

1. Skewness tests
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5. Mixture/latent class SKEWT illustration from Mplus
6. Power analysis exhibit from <https://clincalc.com/stats/samplesize.aspx>
7. Test of differences in attrition from https://www.medcalc.org/calc/comparison_of_proportions.php

1. Skewness tests

Skewness/Kurtosis tests for Normality

| Variable | Obs | Pr(Skewness) | Pr(Kurtosis) | adj chi2(2) | Prob>chi2 |
|----------|-----|--------------|--------------|-------------|-----------|
| totmeth1 | 282 | 0.0000 | 0.0000 | . | 0.0000 |
| totmeth2 | 214 | 0.0000 | 0.0033 | 40.26 | 0.0000 |
| totsith1 | 277 | 0.0016 | 0.0000 | 26.40 | 0.0000 |
| totsith2 | 204 | 0.0087 | 0.0000 | 21.22 | 0.0000 |
| jobperf1 | 277 | 0.0001 | 0.0690 | 15.68 | 0.0004 |
| jobperf2 | 211 | 0.0000 | 0.0241 | 24.66 | 0.0000 |

2. Tests of attrition for the 3 baseline outcomes

2.i. ttest totmeth1 , by(dropped)

Two-sample t test with equal variances

| Group | Obs | Mean | Std. Err. | Std. Dev. | [95% Conf. Interval] | |
|----------|-----|----------|-----------|-----------|----------------------|----------|
| 0 | 214 | 22.0761 | 1.547672 | 22.64049 | 19.02538 | 25.12681 |
| 1 | 68 | 20.15502 | 2.02663 | 16.71202 | 16.10985 | 24.2002 |
| combined | 282 | 21.61286 | 1.271369 | 21.34991 | 19.11024 | 24.11548 |
| diff | | 1.921074 | 2.97516 | | -3.935447 | 7.777594 |

diff = mean(0) - mean(1) t = 0.6457
 Ho: diff = 0 degrees of freedom = 280

Pr(T < t) = 0.7405 Pr(|T| > |t|) = 0.5190 Pr(T > t) = 0.2595
 Ha: diff < 0 Ha: diff != 0 Ha: diff > 0

2.ii. ttest totsith1 , by(dropped)

Two-sample t test with equal variances

| Group | Obs | Mean | Std. Err. | Std. Dev. | [95% Conf. Interval] | |
|----------|-----|----------|-----------|-----------|----------------------|----------|
| 0 | 210 | 9.442857 | .1653386 | 2.395984 | 9.116912 | 9.768802 |
| 1 | 67 | 9.328358 | .3704289 | 3.032091 | 8.588773 | 10.06794 |
| combined | 277 | 9.415162 | .1537501 | 2.558912 | 9.112491 | 9.717834 |
| diff | | .1144989 | .3596304 | | -.5934794 | .8224773 |

diff = mean(0) - mean(1) t = 0.3184
 Ho: diff = 0 degrees of freedom = 275

Ha: diff < 0 Ha: diff != 0 Ha: diff > 0
 Pr(T < t) = 0.6248 Pr(|T| > |t|) = 0.7504 Pr(T > t) = 0.3752

2.iii. ttest jobperf1 , by(dropped)

Two-sample t test with equal variances

| Group | Obs | Mean | Std. Err. | Std. Dev. | [95% Conf. Interval] | |
|-------|-----|------|-----------|-----------|----------------------|--|
|-------|-----|------|-----------|-----------|----------------------|--|

```

-----+-----
      0 |      211      7.251185      .0923241      1.341086      7.069184      7.433186
      1 |       66      6.681818      .2084006      1.693055      6.265613      7.098023
-----+-----
diff |      .5693667      .201977      .1717491      .9669842
-----+-----
diff = mean(0) - mean(1)                                t = 2.8190
Ho: diff = 0                                           degrees of freedom = 275

```

```

Ha: diff < 0
Pr(T < t) = 0.9974
Ha: diff != 0
Pr(|T| > |t|) = 0.0052
Ha: diff > 0
Pr(T > t) = 0.0026

```

```

-----+-----
Variable |      Obs      Mean      Std. Dev.      Min      Max
-----+-----
PostMPAMET |      143      297.4825      453.3436      0      2520

```

3. Illustrative example of how FIML works

3.i. A simple summary: parameter estimates for 1 outcome pre and post:

ec: SAMPLE SIZE FOR POST IS 142 ONLY

```

. sum totmeth1 totmeth2 if group ==1
Variable |      Obs      Mean      Std. Dev.      Min      Max
-----+-----
totmeth1 |      196      22.8708      21.90028      0      119
totmeth2 |      142      27.75332      26.07248      0      119

```

EC: SAMPLE SIZE FOR PAIRED T-TEST DROPS TO POST SAMPLE SIZE 142

```

. ttest totmeth1=totmeth2 if group ==1
ec: SAMPLE SIZE FOR BOTH IS 142 ONLY, SO THE BASELINE MEAN NOW SHIFTED 24.12038 >22.8708
Paired t test

```

```

-----+-----
Variable |      Obs      Mean      Std. Err.      Std. Dev.      [95% Conf. Interval]
-----+-----
totmeth1 |      142      24.12038      1.975248      23.5378      20.21545      28.02531
totmeth2 |      142      27.75332      2.187954      26.07248      23.42789      32.07876
diff |      142      -3.632946      2.231634      26.59299      -8.044734      .7788419

```

```

mean(diff) = mean(totmeth1 - totmeth2)                                t = -1.6279
Ho: mean(diff) = 0                                           degrees of freedom = 141
Ha: mean(diff) < 0
Pr(T < t) = 0.0529
Ha: mean(diff) != 0
Pr(|T| > |t|) = 0.1058
Ha: mean(diff) > 0
Pr(T > t) = 0.9471

```

3.ii. A SEM non-FIML similar descriptive model

. sem totmeth1 totmeth2 if group ==1

(54 observations with missing values excluded)

EC: SEM DROPS THE MISSING POST CASES AND USES ONLY 142 FOR THE ENTIRE MODEL: HENCE THE MEAN ESTIMATE FOR BASELINE FOR WHICH ALL CASES HAD VALUES NOW WILL BE ONLY BASED ON THE 142, NOT 196: SO THE MEAN ESTIMATES DIFFER, SEE BLUE HIGHLIGHTS: 24.12038 > 22.8708

Exogenous variables

Observed: totmeth1 totmeth2

Fitting target model:

Iteration 0: log likelihood = -1299.0972

Iteration 1: log likelihood = -1299.0972

Structural equation model

Number of obs = 142

Estimation method = ml

Log likelihood = -1299.0972

```

-----+-----
|      Coef.      OIM      Std. Err.      z      P>|z|      [95% Conf. Interval]
-----+-----
mean(totmeth1) |      24.12038      1.968281      12.25      0.000      20.26262      27.97813
mean(totmeth2) |      27.75332      2.180236      12.73      0.000      23.48014      32.02651

```

```

-----+-----
      var(totmeth1)|    550.1263    65.28798                435.9567    694.1949
      var(totmeth2)|    674.987    80.10622                534.9047    851.7546
-----+-----
cov(totmeth1,totmeth2)|    261.4532    55.64512        4.70    0.000        152.3908    370.5156
-----+-----
LR test of model vs. saturated: chi2(0)    =        0.00, Prob > chi2 =        .

```

3.iii. A FIML SEM model

```

. sem totmeth1 totmeth2 if group ==1, method(mlmv)
EC: SEM WITH FIML USES NOW ALL 196 CASES TO ESTIMATE BOTH PRE AND POST MEANS AND VARIANCES, SO
N=196: SO THE BASELINE MEAN ESTIMATES ARE THE SAME, SEE BLUE HIGHLIGHTS ABOVE (THE I. SUMMARY
ESTIMATES ): 22.8708 = 22.8708

```

Exogenous variables

Observed: totmeth1 totmeth2

Fitting saturated model:

```

Iteration 0: log likelihood = -1533.3765
Iteration 1: log likelihood = -1532.3527
Iteration 2: log likelihood = -1532.1521
Iteration 3: log likelihood = -1532.152

```

Fitting baseline model:

```

Iteration 0: log likelihood = -1546.6569
Iteration 1: log likelihood = -1546.5972
Iteration 2: log likelihood = -1546.5971

```

Fitting target model:

```

Iteration 0: log likelihood = -1532.152
Iteration 1: log likelihood = -1532.152

```

Structural equation model

Number of obs = 196

Estimation method = mlmv

Log likelihood = -1532.152

```

-----+-----
              |          OIM
              |          Coef.   Std. Err.      z    P>|z|    [95% Conf. Interval]
-----+-----
mean(totmeth1)|    22.8708    1.56031    14.66    0.000    19.81265    25.92895
mean(totmeth2)|    27.15945    2.10696    12.89    0.000    23.02988    31.28901
-----+-----
var(totmeth1)|    477.1751    48.20197                391.4655    581.6506
var(totmeth2)|    658.5094    76.42463                524.5345    826.7037
-----+-----
cov(totmeth1,totmeth2)|    226.7824    46.15245        4.91    0.000    136.3253    317.2395
-----+-----
LR test of model vs. saturated: chi2(0)    =        0.00, Prob > chi2 =        .

```

3.iv. A SEM (with Stata gsem) SEM model - default is FIML for gsem!!!

```

. gsem (-> totmeth1 totmeth2) if group ==1

```

EC: gsem LISTS THE SAMPLES WITH VALID SCORES IN EACH VARIABLE BUT USES AS DEFAULT FIML AND HENCE THE ENTIRE SAMPLE TO ESTIMATE ALL PARAMETERS, LIKE SEM WITH method(mlmv)

```

Iteration 0: log likelihood = -1546.5971
Iteration 1: log likelihood = -1546.5971

```

Generalized structural equation model

Number of obs = 196

Response : totmeth1
Family : Gaussian
Link : identity

Number of obs = 196

Response : totmeth2
Family : Gaussian
Link : identity
Log likelihood = -1546.5971

Number of obs = 142

| | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
|-----------------|----------|-----------|-------|-------|----------------------|----------|
| totmeth1 | | | | | | |
| _cons | 22.8708 | 1.56031 | 14.66 | 0.000 | 19.81265 | 25.92895 |
| totmeth2 | | | | | | |
| _cons | 27.75332 | 2.180236 | 12.73 | 0.000 | 23.48014 | 32.02651 |
| var(e.totmeth1) | 477.1751 | 48.20197 | | | 391.4655 | 581.6506 |
| var(e.totmeth2) | 674.987 | 80.10622 | | | 534.9047 | 851.7546 |

4. Latent change score Stata model results:

Physical activity

```
. sem (latentch21 ctotmeth1@1 _cons@0 ->ctotmeth2) (ctotmeth1 _cons ->latentch21),
group(group) ginvariant(n
> one) nocapslatent latent(latentch21) var(e.ctotmeth2@0 e.latentch21 ) method(mlmv)
showginvariant
```

Endogenous variables

Observed: ctotmeth2
Latent: latentch21

Exogenous variables

Observed: ctotmeth1

Structural equation model

Grouping variable = group

Estimation method = mlmv

Log likelihood = -2259.5882

Number of obs = 282
Number of groups = 2

- (1) [ctotmeth2]1bn.group#c.latentch21 = 1
- (2) [ctotmeth2]1bn.group#c.ctotmeth1 = 1
- (3) [/]var(e.ctotmeth2)#1bn.group = 0
- (4) [ctotmeth2]1bn.group = 0
- (5) [ctotmeth2]2.group#c.latentch21 = 1
- (6) [ctotmeth2]2.group#c.ctotmeth1 = 1
- (7) [/]var(e.ctotmeth2)#2.group = 0
- (8) [ctotmeth2]2.group = 0

| | Coef. | OIM Std. Err. | z | P> z | [95% Conf. Interval] | |
|-------------------|-----------|------------------|-------|-------|----------------------|-----------|
| Structural | | | | | | |
| ctotmeth2 | | | | | | |
| latentch21 | | | | | | |
| [*] | 1 | (constrained) | | | | |
| ctotmeth1 | | | | | | |
| [*] | 1 | (constrained) | | | | |
| _cons | | | | | | |
| [*] | 0 | (constrained) | | | | |
| latentch21 | | | | | | |
| ctotmeth1 | | | | | | |
| Guangzhou | -.5247397 | .0839641 | -6.25 | 0.000 | -.6893062 | -.3601731 |
| Beijing | -.8910333 | .1797252 | -4.96 | 0.000 | -1.243288 | -.5387785 |
| _cons | | | | | | |
| Guangzhou | 5.79507 | 1.999515 | 2.90 | 0.004 | 1.876092 | 9.714048 |
| Beijing | 7.409048 | 3.64259 | 2.03 | 0.042 | .2697015 | 14.54839 |
| var(e.ctotmeth2) | | | | | | |
| [*] | 0 | (constrained) | | | | |
| var(e.latentch21) | | | | | | |

| | | | | |
|-----------|----------|----------|---------|---------|
| Guangzhou | 550.7286 | 65.35945 | 436.434 | 694.955 |
| Beijing | 946.4345 | 157.7385 | 682.691 | 1312.07 |

Note: [*] identifies parameter estimates constrained to be equal across groups.
 LR test of model vs. saturated: $\chi^2(0) = 0.00$, Prob > $\chi^2 = .$

```
. * intercepts:
. lincom _b[latentch21:1bn.group] - _b[latentch21:2.group]

(1) [latentch21]1bn.group - [latentch21]2.group = 0
```

| | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] |
|-----|-----------|-----------|-------|-------|----------------------|
| (1) | -1.613977 | 4.155301 | -0.39 | 0.698 | -9.758217 6.530263 |

```
. * prop-growth paths:
. lincom _b[latentch21:1bn.group#c.ctotmeth1]- _b[latentch21:2.group#c.ctotmeth1]

(1) [latentch21]1bn.group#c.ctotmeth1 - [latentch21]2.group#c.ctotmeth1 = 0
```

| | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] |
|-----|----------|-----------|------|-------|----------------------|
| (1) | .3662937 | .1983711 | 1.85 | 0.065 | -.0225066 .7550939 |

WITHOUT PROPORTIONAL GROWTH FOR ILLUSTRATION:

| | | OIM | | | | |
|------------|--------------------------|-----------|---------------|-------|-------|----------------------|
| | | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] |
| Structural | | | | | | |
| totmeth2 | | | | | | |
| | totmeth1 | | | | | |
| | [*] | 1 | (constrained) | | | |
| | latentch21 | | | | | |
| | [*] | 1 | (constrained) | | | |
| | _cons | | | | | |
| | [*] | 0 | (constrained) | | | |
| | | | | | | |
| | mean(totmeth1) | | | | | |
| | Guangzhou | 22.8708 | 1.56031 | 14.66 | 0.000 | 19.81265 25.92895 |
| | Beijing | 18.74593 | 2.129581 | 8.80 | 0.000 | 14.57203 22.91983 |
| | mean(latentch21) | | | | | |
| | Guangzhou | 4.288648 | 2.135355 | 2.01 | 0.045 | .103429 8.473867 |
| | Beijing | 8.52647 | 4.09408 | 2.08 | 0.037 | .5022211 16.55072 |
| | | | | | | |
| | var(e.totmeth2) | | | | | |
| | [*] | 0 | (constrained) | | | |
| | var(totmeth1) | | | | | |
| | Guangzhou | 477.1751 | 48.20197 | | | 391.4655 581.6506 |
| | Beijing | 390.02 | 59.47749 | | | 289.2543 525.8888 |
| | var(latentch21) | | | | | |
| | Guangzhou | 682.1197 | 78.84195 | | | 543.8455 855.5505 |
| | Beijing | 1256.091 | 206.6782 | | | 909.8372 1734.117 |
| | | | | | | |
| | cov(totmeth1,latentch21) | | | | | |
| | Guangzhou | -250.3927 | 47.38154 | -5.28 | 0.000 | -343.2588 -157.5266 |
| | Beijing | -347.521 | 87.87576 | -3.95 | 0.000 | -519.7543 -175.2877 |

Note: [*] identifies parameter estimates constrained to be equal across groups.
 LR test of model vs. saturated: $\chi^2(0) = 0.00$, Prob > $\chi^2 = .$

```
. *nlcom _b[/mean(latentch21)#1bn.group] - _b[/mean(latentch21)#2.group]
. lincom _b[/mean(latentch21)#1bn.group] - _b[/mean(latentch21)#2.group]
```

```
( 1)  [/]mean(latentch21)#1bn.group - [/]mean(latentch21)#2.group = 0
```

| | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
|-----|-----------|-----------|-------|-------|----------------------|----------|
| (1) | -4.237821 | 4.617492 | -0.92 | 0.359 | -13.28794 | 4.812296 |

Work performance

```
. sem (latentch21@1 cjobperf1@1 _cons@0 ->cjobperf2) (cjobperf1 _cons -> latentch21),
group(group) ginvariant(non
> e) nocapslatent latent(latentch21) var(e.cjobperf2@0 e.latentch21 ) method(mlmv) showginvariant
(2 all-missing observations excluded)
note: Missing values found in observed exogenous variables. Using the noxconditional behavior.
Specify the
```

forcexconditional option to override this behavior.

Endogenous variables

Observed: cjobperf2
Latent: latentch21

Exogenous variables

Observed: cjobperf1

```
Structural equation model      Number of obs      =      280
Grouping variable = group      Number of groups   =       2
Estimation method = mlmv
Log likelihood      = -872.68286
```

```
( 1)  [cjobperf2]1bn.group#c.latentch21 = 1
( 2)  [cjobperf2]1bn.group#c.cjobperf1 = 1
( 3)  [/]var(e.cjobperf2)#1bn.group = 0
( 4)  [cjobperf2]1bn.group = 0
( 5)  [cjobperf2]2.group#c.latentch21 = 1
( 6)  [cjobperf2]2.group#c.cjobperf1 = 1
( 7)  [/]var(e.cjobperf2)#2.group = 0
( 8)  [cjobperf2]2.group = 0
```

| | Coef. | OIM Std. Err. | z | P> z | [95% Conf. Interval] | |
|-----------------|-----------|------------------|-------|-------|----------------------|-----------|
| Structural | | | | | | |
| cjobperf2 | | | | | | |
| latentch21 | | | | | | |
| [*] | 1 | (constrained) | | | | |
| cjobperf1 | | | | | | |
| [*] | 1 | (constrained) | | | | |
| _cons | | | | | | |
| [*] | 0 | (constrained) | | | | |
| latentch21 | | | | | | |
| cjobperf1 | | | | | | |
| Guangzhou | -.5391502 | .0912282 | -5.91 | 0.000 | -.7179542 | -.3603463 |
| Beijing | -1.241835 | .1647798 | -7.54 | 0.000 | -1.564797 | -.9188723 |
| _cons | | | | | | |
| Guangzhou | -.0327876 | .1178219 | -0.28 | 0.781 | -.2637142 | .1981391 |
| Beijing | .6914081 | .2440482 | 2.83 | 0.005 | .2130825 | 1.169734 |
| mean(cjobperf1) | | | | | | |
| Guangzhou | -.1025018 | .1012937 | -1.01 | 0.312 | -.3010337 | .0960301 |

| | | | | | | |
|-------------------|-----------------|----------|------|-------|----------|----------|
| Beijing | .6282801 | .1531447 | 4.10 | 0.000 | .3281221 | .9284382 |
| ----- | | | | | | |
| var(e.cjobperf2) | | | | | | |
| [*] | 0 (constrained) | | | | | |
| var(e.latentch21) | | | | | | |
| Guangzhou | 1.952079 | .2331145 | | | 1.544714 | 2.466872 |
| Beijing | 2.985658 | .504855 | | | 2.143421 | 4.158843 |
| var(c.jobperf1) | | | | | | |
| Guangzhou | 1.983593 | .2019436 | | | 1.624779 | 2.421647 |
| Beijing | 1.970682 | .3040412 | | | 1.456435 | 2.666501 |

Note: [*] identifies parameter estimates constrained to be equal across groups.

LR test of model vs. saturated: $\chi^2(0) = 0.00$, Prob > $\chi^2 = .$

```
. lincom _b[latentch21:1bn.group] - _b[latentch21:2.group]
( 1) [latentch21]1bn.group - [latentch21]2.group = 0
```

| | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
|-----|-----------|-----------|-------|-------|----------------------|-----------|
| (1) | -.7241957 | .2710009 | -2.67 | 0.008 | -1.255348 | -.1930436 |

* prop-growth paths:

```
. lincom _b[latentch21:1bn.group#c.cjobperf1] - _b[latentch21:2.group#c.cjobperf1]
( 1) [latentch21]1bn.group#c.cjobperf1 - [latentch21]2.group#c.cjobperf1 = 0
```

| | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
|-----|----------|-----------|------|-------|----------------------|---------|
| (1) | .7026845 | .188348 | 3.73 | 0.000 | .3335292 | 1.07184 |

WITHOUT PROPORTIONAL GROWTH FOR ILLUSTRATION:

```
. sem (latentch21 -> jobperf2@1) (jobperf1@1 _cons@0 -> jobperf2), group(group)
ginvariant(none) nocapslat
> ent latent(latentch21) means(latentch21) var(e.jobperf2@0 latentch21) method(mlmv)
showginvariant
(2 all-missing observations excluded)
```

Endogenous variables

Observed: jobperf2

Exogenous variables

Observed: jobperf1

Latent: latentch21

Structural equation model

Grouping variable = group

Estimation method = mlmv

Log likelihood = -872.68286

Number of obs = 280

Number of groups = 2

- (1) [jobperf2]1bn.group#c.jobperf1 = 1
- (2) [jobperf2]1bn.group#c.latentch21 = 1
- (3) [/]var(e.jobperf2)#1bn.group = 0
- (4) [jobperf2]1bn.group = 0
- (5) [jobperf2]2.group#c.jobperf1 = 1
- (6) [jobperf2]2.group#c.latentch21 = 1
- (7) [/]var(e.jobperf2)#2.group = 0
- (8) [jobperf2]2.group = 0

| | | OIM | | | |
|--|-------|-----------|---|------|----------------------|
| | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] |

```

Structural
  jobperf2
    jobperf1
      [*] 1 (constrained)
    latentch21
      [*] 1 (constrained)
    _cons
      [*] 0 (constrained)
-----
mean(jobperf1)
  Guangzhou 6.897498 .1012937 68.09 0.000 6.698966 7.09603
  Beijing 7.62828 .1531447 49.81 0.000 7.328122 7.928438
mean(latentch21)
  Guangzhou .0224763 .1303576 0.17 0.863 -.23302 .2779726
  Beijing -.088812 .281877 -0.32 0.753 -.6412808 .4636568
-----
var(e.jobperf2)
  [*] 0 (constrained)
var(jobperf1)
  Guangzhou 1.983593 .2019436 1.624779 2.421647
  Beijing 1.970682 .3040412 1.456435 2.666501
var(latentch21)
  Guangzhou 2.528675 .3107017 1.987489 3.217224
  Beijing 6.024752 1.059153 4.26872 8.503165
-----
cov(jobperf1,latentch21)
  Guangzhou -1.069455 .2112161 -5.06 0.000 -1.483431 -.6554785
  Beijing -2.447261 .4976786 -4.92 0.000 -3.422693 -1.471829
-----LR test
of model vs. saturated: chi2(0) = 0.00, Prob > chi2 = .
. lincom _b[/mean(latentch21)#1bn.group] - _b[/mean(latentch21)#2.group]

( 1) [/]mean(latentch21)#1bn.group - [/]mean(latentch21)#2.group = 0
-----
| Coef. Std. Err. z P>|z| [95% Conf. Interval]
-----+-----
(1) | .1112883 .3105604 0.36 0.720 -.4973989 .7199756
-----

```

Sitting hours

```

. sem (latentch21 totsith1@1 _cons@0 ->totsith2) (totsith1 _cons ->latentch21), gr
> oup(group) ginvariant(none) nocapslatent latent(latentch21) var(e.totsith2@0 e.la
> tentch21 ) method(mlmv) showginvariant
(3 all-missing observations excluded)

```

note: Missing values found in observed exogenous variables. Using the
noxconditional behavior. Specify the forcexconditional option to override
this behavior.

Endogenous variables

Observed: totsith2
Latent: latentch21

Exogenous variables

Observed: totsith1

Structural equation model
Grouping variable = group
Estimation method = mlmv
Log likelihood = -1102.5334

Number of obs = 279
Number of groups = 2

```

( 1) [totsith2]1bn.group#c.latentch21 = 1
( 2) [totsith2]1bn.group#c.totsith1 = 1

```



```
( 3) [/]var(e.totsith2)#1bn.group = 0
( 4) [totsith2]1bn.group = 0
( 5) [totsith2]2.group#c.latentch21 = 1
( 6) [totsith2]2.group#c.totsith1 = 1
( 7) [/]var(e.totsith2)#2.group = 0
( 8) [totsith2]2.group = 0
```

| | Coef. | OIM Std. Err. | z | P> z | [95% Conf. Interval] | |
|-------------------|-----------|------------------|--------|-------|----------------------|-----------|
| Structural | | | | | | |
| totsith2 | | | | | | |
| latentch21 | | | | | | |
| [*] | 1 | (constrained) | | | | |
| totsith1 | | | | | | |
| [*] | 1 | (constrained) | | | | |
| _cons | | | | | | |
| [*] | 0 | (constrained) | | | | |
| latentch21 | | | | | | |
| totsith1 | | | | | | |
| Guangzhou | -.5851763 | .0826396 | -7.08 | 0.000 | -.7471469 | -.4232056 |
| Beijing | -1.101411 | .1097504 | -10.04 | 0.000 | -1.316518 | -.8863041 |
| _cons | | | | | | |
| Guangzhou | 5.6775 | .8156443 | 6.96 | 0.000 | 4.078867 | 7.276134 |
| Beijing | 10.34172 | 1.037181 | 9.97 | 0.000 | 8.30888 | 12.37456 |
| mean(totsith1) | | | | | | |
| Guangzhou | 9.507772 | .1934689 | 49.14 | 0.000 | 9.12858 | 9.886964 |
| Beijing | 9.20747 | .2403564 | 38.31 | 0.000 | 8.73638 | 9.67856 |
| var(e.totsith2) | | | | | | |
| [*] | 0 | (constrained) | | | | |
| var(e.latentch21) | | | | | | |
| Guangzhou | 5.499059 | .6693242 | | | 4.33195 | 6.980609 |
| Beijing | 3.996969 | .6806905 | | | 2.86266 | 5.580741 |
| var(totsith1) | | | | | | |
| Guangzhou | 7.224033 | .7353872 | | | 5.917383 | 8.819211 |
| Beijing | 4.851632 | .7485811 | | | 3.585518 | 6.564833 |

Note: [*] identifies parameter estimates constrained to be equal across groups.

LR test of model vs. saturated: chi2(0) = 0.00, Prob > chi2 = .

. * intercepts:

. lincom _b[latentch21:1bn.group] - _b[latentch21:2.group]

```
( 1) [latentch21]1bn.group - [latentch21]2.group = 0
```

| | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
|-----|-----------|-----------|-------|-------|----------------------|----------|
| (1) | -4.664218 | 1.319478 | -3.53 | 0.000 | -7.250346 | -2.07809 |

. * prop-growth paths:

. lincom _b[latentch21:1bn.group#c.totsith1] - _b[latentch21:2.group#c.totsith1]

```
( 1) [latentch21]1bn.group#c.totsith1 - [latentch21]2.group#c.totsith1 = 0
```

| | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
|-----|----------|-----------|------|-------|----------------------|----------|
| (1) | .5162347 | .1373844 | 3.76 | 0.000 | .2469663 | .7855031 |

WITHOUT PROPORTIONAL GROWTH FOR ILLUSTRATION:

```
. sem (latentch21 totsith1@1 _cons@0 ->totsith2),          group(group) ginvariant(no
> ne) nocapslatent latent(latentch21) means(latentch21 ) var(e.totsith2@0 latentch21
> ) method(mlmv) showginvariant
(3 all-missing observations excluded)
```

Endogenous variables

Observed: totsith2

Exogenous variables

Observed: totsith1

Latent: latentch21

| | | | |
|---------------------------|------------------|------------|-----|
| Structural equation model | Number of obs | = | 279 |
| Grouping variable = group | Number of groups | = | 2 |
| Estimation method = mlmv | | | |
| Log likelihood | = | -1102.5334 | |

- (1) [totsith2]1bn.group#c.totsith1 = 1
- (2) [totsith2]1bn.group#c.latentch21 = 1
- (3) [/]var(e.totsith2)#1bn.group = 0
- (4) [totsith2]1bn.group = 0
- (5) [totsith2]2.group#c.totsith1 = 1
- (6) [totsith2]2.group#c.latentch21 = 1
- (7) [/]var(e.totsith2)#2.group = 0
- (8) [totsith2]2.group = 0

| | Coef. | OIM Std. Err. | z | P> z | [95% Conf. Interval] | |
|------------------|-----------|------------------|-------|-------|----------------------|-----------|
| Structural | | | | | | |
| totsith2 | | | | | | |
| totsith1 | | | | | | |
| [*] | 1 | (constrained) | | | | |
| latentch21 | | | | | | |
| [*] | 1 | (constrained) | | | | |
| _cons | | | | | | |
| [*] | 0 | (constrained) | | | | |
| mean(totsith1) | | | | | | |
| Guangzhou | 9.507772 | .1934689 | 49.14 | 0.000 | 9.12858 | 9.886964 |
| Beijing | 9.20747 | .2403564 | 38.31 | 0.000 | 8.73638 | 9.67856 |
| mean(latentch21) | | | | | | |
| Guangzhou | .1137779 | .2314559 | 0.49 | 0.623 | -.3398674 | .5674232 |
| Beijing | .2005102 | .3573284 | 0.56 | 0.575 | -.4998406 | .900861 |
| var(e.totsith2) | | | | | | |
| [*] | 0 | (constrained) | | | | |
| var(totsith1) | | | | | | |
| Guangzhou | 7.224033 | .7353872 | | | 5.917383 | 8.819211 |
| Beijing | 4.851632 | .7485811 | | | 3.585518 | 6.564833 |
| var(latentch21) | | | | | | |
| Guangzhou | 7.972794 | .999788 | | | 6.23548 | 10.19415 |
| Beijing | 9.882513 | 1.629717 | | | 7.153136 | 13.65332 |
| cov(totsith1, | | | | | | |
| latentch21) | | | | | | |
| Guangzhou | -4.227333 | .7359236 | -5.74 | 0.000 | -5.669716 | -2.784949 |
| Beijing | -5.34364 | .9811248 | -5.45 | 0.000 | -7.266609 | -3.420671 |

Note: [*] identifies parameter estimates constrained to be equal across groups.

LR test of model vs. saturated: chi2(0) = 0.00, Prob > chi2 = .

```
. *nlcom _b[/mean(latentch21)#1bn.group] - _b[/mean(latentch21)#2.group]
```

```
. lincom _b[/mean(latentch21)#1bn.group] - _b[/mean(latentch21)#2.group]
```

```
( 1)  [/]mean(latentch21)#1bn.group - [/]mean(latentch21)#2.group = 0
```

| ----- | | | | | | |
|-------|-----------|-----------|-------|-------|----------------------|----------|
| | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
| ----- | | | | | | |
| (1) | -.0867323 | .425741 | -0.20 | 0.839 | -.9211694 | .7477048 |
| ----- | | | | | | |

5. Mixture/latent class SKEWT illustration from Mplus

```
---3-----T LCS model-----Sitting-----
```

```
NS.C > NS.I, NS.Diff
```

```
SYNTAX for SKEWT Mplus model:
```

```
Usevariables are !
```

```

y2 y1 ;
IDVARIABLE = id;
CLASSES = c1t2 (2) ;
      KNOWNCLASS = c1t2 (group = 1 group = 2);
Define:
y1 = ctotsith1; !centered on PRE C=BJ mean
y2 = ctotsith2; !centered on PRE C=BJ mean
ANALYSIS: TYPE = MIXTURE;
      COVERAGE = 0;
      DISTRIBUTION = SKEWT;! SKEW SKEWT normal tdistribution
      STARTS = 32 8; ! typically not necessary
      PROCESSORS = 8;
```

```
Model:
```

```
%OVERALL%
```

```

LCS by y2@1;
y2@0;
[y2@0];
[LCS] ;
LCS on y1;
y2 on y1@1;
!y1 on ;! covariates possible
!LCS on tot1 ;! proportional growth path; also covariates possible
```

```
%c1t2#1%
```

```

LCS by y2@1;
y2@0;
[y2@0];
[LCS] (CLCSInt);
LCS on y1;
y2 on y1@1;
!y1 on ;! covariates possible
!LCS on tot1 ;! proportional growth path; also covariates possible
```

```
%c1t2#2%
```

```

LCS by y2@1;
y2@0;
[y2@0];
[LCS] (TLCSInt);
LCS on y1;
y2 on y1@1;
!y1 on ;! covariates possible
!LCS on tot1 ;! proportional growth path; also covariates possible
```

```
Model constraint:
```

```
New (TE);
```

```
TE=TLCSInt-CLCSInt; !T-C difference tested for significance
```

```
THE MODEL ESTIMATION TERMINATED NORMALLY
```

```
MODEL FIT INFORMATION
```

```
Number of Free Parameters
```

```
10
```

```
Loglikelihood
```

```
H0 Value
```

```
-564.172
```

H0 Scaling Correction Factor 1.1269
for MLR
Information Criteria
Akaike (AIC) 1148.343
Bayesian (BIC) 1181.426
Sample-Size Adjusted BIC 1149.744
(n* = (n + 2) / 24)

FINAL CLASS COUNTS AND PROPORTIONS FOR THE LATENT CLASSES
BASED ON THE ESTIMATED MODEL

Latent
Classes
!EC: these are actually observed classes/known
1 135.00000 0.66832
2 67.00000 0.33168

MODEL RESULTS

| | Estimate | S.E. | Est./S.E. | Two-Tailed P-Value |
|-----------------------------------|----------|-------|-----------|-----------------------|
| Latent Class 1 (1) = INTERVENTION | | | | |
| LCS BY | | | | |
| Y2 | 1.000 | 0.000 | 999.000 | 999.000 |
| LCS ON | | | | |
| Y1 | -0.517 | 0.169 | -3.057 | 0.002 |
| Y2 ON | | | | |
| Y1 | 1.000 | 0.000 | 999.000 | 999.000 |
| Intercepts | | | | |
| Y2 | 0.000 | 0.000 | 999.000 | 999.000 |
| LCS | 0.014 | 0.586 | 0.024 | 0.981 |
| Residual Variances | | | | |
| Y2 | 0.000 | 0.000 | 999.000 | 999.000 |
| LCS | 2.186 | 0.575 | 3.799 | 0.000 |
| Latent Class 2 (2) = CONTROLS | | | | |
| LCS BY | | | | |
| Y2 | 1.000 | 0.000 | 999.000 | 999.000 |
| LCS ON | | | | |
| Y1 | -1.097 | 0.064 | -17.162 | 0.000 |
| Y2 ON | | | | |
| Y1 | 1.000 | 0.000 | 999.000 | 999.000 |
| Intercepts | | | | |
| Y2 | 0.000 | 0.000 | 999.000 | 999.000 |
| LCS | -0.304 | 0.680 | -0.448 | 0.654 |
| Residual Variances | | | | |
| Y2 | 0.000 | 0.000 | 999.000 | 999.000 |
| LCS | 2.186 | 0.575 | 3.799 | 0.000 |

Categorical Latent Variables

Means
C1T2#1 0.701 0.149 4.688 0.000

!EC: MPLUS TESTS FOR SKEW THE DV/EFFECT VARIABLES ONLY, AND THE LCAS MODEL SEEMS TO SHOW THE LCS SCORE FOR PA IS NOT SKEWED

Skew and Df Parameters

| | | | | |
|---------------------------|--------|-------|--------|-------|
| Latent Class 1 (1) | | | | |
| LCS | 0.353 | 0.662 | 0.534 | 0.594 |
| DF | 2.868 | 0.840 | 3.412 | 0.001 |
| Latent Class 2 (2) | | | | |
| LCS | 0.863 | 0.807 | 1.069 | 0.285 |
| DF | 5.013 | 2.731 | 1.836 | 0.066 |
| New/Additional Parameters | | | | |
| TE | -0.318 | 0.839 | -0.380 | 0.704 |

IF ONE TESTS ONLY tdistribution BY CHANGING ONLY:

DISTRIBUTION = tdistribution;

MODEL RESULTS

Two-Tailed

| | Estimate | S.E. | Est./S.E. | P-Value |
|-----------------------------------|----------|-------|-----------|---------|
| Latent Class 1 (1) = INTERVENTION | | | | |
| LCS BY | | | | |
| Y2 | 1.000 | 0.000 | 999.000 | 999.000 |
| LCS ON | | | | |
| Y1 | -0.530 | 0.168 | -3.165 | 0.002 |
| Y2 ON | | | | |
| Y1 | 1.000 | 0.000 | 999.000 | 999.000 |
| Intercepts | | | | |
| Y2 | 0.000 | 0.000 | 999.000 | 999.000 |
| LCS | 0.337 | 0.206 | 1.634 | 0.102 |
| Residual Variances | | | | |
| Y2 | 0.000 | 0.000 | 999.000 | 999.000 |
| LCS | 2.451 | 0.494 | 4.966 | 0.000 |
| Latent Class 2 (2) = CONTROLS | | | | |
| LCS BY | | | | |
| Y2 | 1.000 | 0.000 | 999.000 | 999.000 |
| LCS ON | | | | |
| Y1 | -1.085 | 0.068 | -16.056 | 0.000 |
| Y2 ON | | | | |
| Y1 | 1.000 | 0.000 | 999.000 | 999.000 |
| Intercepts | | | | |
| Y2 | 0.000 | 0.000 | 999.000 | 999.000 |
| LCS | 0.441 | 0.241 | 1.832 | 0.067 |
| Residual Variances | | | | |
| Y2 | 0.000 | 0.000 | 999.000 | 999.000 |
| LCS | 2.451 | 0.494 | 4.966 | 0.000 |
| Categorical Latent Variables | | | | |
| Means | | | | |
| C1T2#1 | 0.701 | 0.149 | 4.688 | 0.000 |
| Skew and Df Parameters | | | | |
| Latent Class 1 (1) | | | | |
| DF | 3.108 | 0.812 | 3.825 | 0.000 |
| Latent Class 2 (2) | | | | |
| DF | 5.332 | 3.099 | 1.721 | 0.085 |
| New/Additional Parameters | | | | |
| TE | 0.104 | 0.305 | 0.341 | 0.733 |



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6. Power analysis exhibit from <https://clincalc.com/stats/samplesize.aspx>

Appendix Move-It: A cluster-randomised digital worksite exercise intervention

≡ Menu

Sample Size Calculator

Determines the minimum number of subjects for adequate study power

 [ClinCalc.com \(/\)](https://clincalc.com/) » [Statistics \(/Statistics\)](/Statistics/) » Sample Size Calculator

Study Group Design



vs.

Two independent
study groups

vs.

One study group
vs. population

Two study groups will each receive different treatments.

Primary Endpoint



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Top



Dichotomous
(yes/no)

Sample Size Calculator



Continuous
(means)

The primary endpoint is **binomial** - only two possible outcomes.
Eg, mortality (dead/not dead), pregnant (pregnant/not)

Statistical Parameters

Anticipated Means

Group 1 [?](#)

5 ± 15

Group 2 [?](#)

12

Mean ▼

Enrollment ratio [?](#)

1

Type I/II Error Rate

Alpha [?](#)

0.05

Power [?](#)

80%

Reset

Calculate

⬆️ Back to Top
⬆️ Top

RESULTS

Continuous Endpoint, Two Independent Sample Study

| Sample Size | |
|--------------|------------|
| Group 1 | 72 |
| Group 2 | 72 |
| Total | 144 |

| Study Parameters | |
|------------------|------|
| Mean, group 1 | 5 |
| Mean, group 2 | 12 |
| Alpha | 0.05 |
| Beta | 0.2 |
| Power | 0.8 |

 **View Power Calculations**

About This Calculator

This calculator uses a number of different equations to determine the minimum number of subjects that need to be enrolled in a study in order to have sufficient statistical power to detect a treatment effect.¹

Before a study is conducted, investigators need to determine how many subjects should be included. By enrolling too few subjects, a study may not have enough statistical power to detect a difference (type II error). Enrolling too many patients can be unnecessarily costly or time-consuming.

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Generally speaking, statistical power is determined by the following variables:

- **Baseline Incidence:** If an outcome occurs infrequently, many more patients are needed in order to detect a difference.
- **Population Variance:** The higher the variance (standard deviation), the more patients are needed to demonstrate a difference.
- **Treatment Effect Size:** If the difference between two treatments is small, more patients will be required to detect a difference.
- **Alpha:** The probability of a type-I error -- finding a difference when a difference does not exist. Most medical literature uses an alpha cut-off of 5% (0.05) -- indicating a 5% chance that a significant difference is actually due to chance and is not a true difference.
- **Beta:** The probability of a type-II error -- not detecting a difference when one actually exists. Beta is directly related to study power ($\text{Power} = 1 - \beta$). Most medical literature uses a beta cut-off of 20% (0.2) -- indicating a 20% chance that a significant difference is missed.

Post-Hoc Power Analysis

To calculate the post-hoc statistical power of an existing trial, please visit the [post-hoc power analysis calculator \(Power.aspx\)](#).

References and Additional Reading

1. Rosner B. *Fundamentals of Biostatistics*. 7th ed. Boston, MA: Brooks/Cole; 2011.

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Related Calculators

- [Post-hoc Power Calculator \(Power.aspx\)](#)

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Comparison of proportions calculator

7. Test of differences in attrition from https://www.medcalc.org/calc/comparison_of_proportions.php
 Appendix Move-It: A cluster-randomised digital worksite exercise intervention 1

Sample 1

Proportion (%): 27

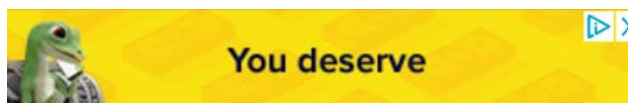
Sample size: 196

Sample 2

Proportion (%): 15.1

Sample size: 86

Test



Results

| | |
|--------------------|---------------------|
| Difference | 11.9 % |
| 95% CI | 1.1934% to 20.8701% |
| Chi-squared | 4.710 |
| DF | 1 |
| Significance level | P = 0.0300 |

Computational notes

MedCalc uses the "N-1" Chi-squared test as recommended by Campbell (2007) and Richardson (2011).

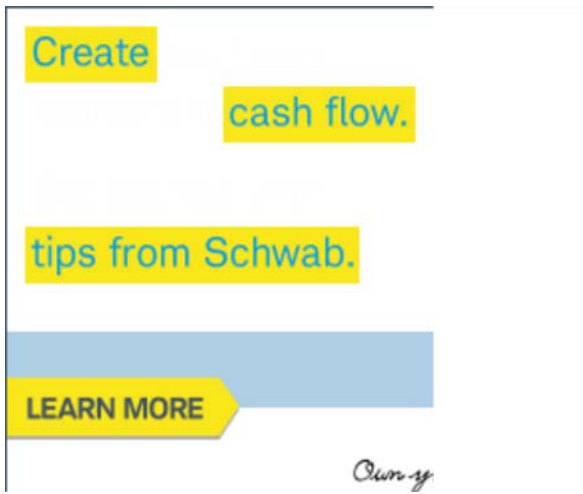
The confidence interval is calculated according to the recommended method given by Altman et al. (2000).

Literature

- Altman DG, Machin D, Bryant TN, Gardner MJ (Eds) (2000) Statistics with confidence (2nd edn) (London: BMJ Books, 2000) 49)

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- Campbell I (2007) Chi-squared and Fisher-Irwin tests of two-by-two tables with small sample recommendations. *Statistics in Medicine* 26:3661-3675. [PubMed](#)
- Richardson JTE (2011) The analysis of 2 x 2 contingency tables - Yet again. *Statistics in Medicine* 30:890. [PubMed](#)



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