

Article

Factor Analysis of Genetic Parameters for Body Conformation Traits in Dual-Purpose Simmental Cattle

Lei Xu ¹, Hanpeng Luo ², Xiaoxue Zhang ¹, Haibo Lu ³, Menghua Zhang ¹, Jianjun Ge ⁴, Tao Zhang ¹, Mengjie Yan ¹, Xueting Tan ¹, Xixia Huang ^{1,*} and Yachun Wang ^{2,*} 

¹ College of Animal Science, Xinjiang Agricultural University, Urumqi 830052, China

² College of Animal Science and Technology, China Agricultural University, Beijing 100193, China

³ Beijing SUNLON Biological Seed Industry Innovation Technology Limited Company, Beijing 101206, China

⁴ Xinjiang Hutubi Farm, Changji 831200, China

* Correspondence: au-huangxixia@163.com (X.H.); wangyachun@cau.edu.cn (Y.W.);

Tel.: +86-1399-999-6861 (X.H.); +86-1580-159-5851 (Y.W.)

Simple Summary: Body conformation traits are closely related to economically important characteristics and should be considered in cattle breeding programs. A variety of body conformation traits recorded by classifiers can complicate the analysis process. Factor analysis can reduce the number of variables by combining two or more variables into a single factor, which has biological significance. The results of this study could be used by breeders to define conformation indexes and implement genetic assessments for conformation traits in dual-purpose breeds.

Abstract: In this study, we estimated the genetic parameters for 6 composite traits and 27 body conformation traits of 1016 dual-purpose Simmental cattle reared in northwestern China from 2010 to 2019 using a linear animal mixed model. To integrate these traits, a variety of methods were used as follows: (1) genetic parameters estimates for composite and individual body conformation traits based on the pedigree relationship matrix (A) and combined genomic-pedigree relationship matrix (H); (2) factor analysis to explore the relationships among body conformation traits; and (3) genetic parameters of factor scores estimated using A and H, and the correlations of EBVs of the factor scores and EBVs of the composite traits. Heritability estimates of the composite traits using A and H were low to medium (0.07–0.47). The 24 common latent factors explained 96.13% of the total variance. Among factors with eigenvalues ≥ 1 , F1 was mainly related to body frame, muscularity, and rump; F2 was related to feet and legs; F3, F4, F5, and F6 were related to teat placement, teat size, udder size, and udder conformation; and F7 was related to body frame. Single-trait analysis of factor scores yielded heritability estimates that were low to moderate (0.008–0.43 based on A and 0.04–0.43 based on H). Spearman and Pearson correlations, derived from the best linear unbiased prediction analysis of composite traits and factor scores, showed a similar pattern. Thus, incorporating factor analysis into the morphological evaluation to simplify the assessment of body conformation traits may improve the genetics of dual-purpose Simmental cattle.

Keywords: Simmental cattle; genetic parameter; body conformation traits; factor analysis



Citation: Xu, L.; Luo, H.; Zhang, X.; Lu, H.; Zhang, M.; Ge, J.; Zhang, T.; Yan, M.; Tan, X.; Huang, X.; et al.

Factor Analysis of Genetic Parameters for Body Conformation Traits in Dual-Purpose Simmental Cattle. *Animals* **2022**, *12*, 2433.

<https://doi.org/10.3390/ani12182433>

Academic Editor: Martino Cassandro

Received: 17 July 2022

Accepted: 12 September 2022

Published: 15 September 2022

Publisher's Note: MDPI stays neutral with regard to jurisdictional claims in published maps and institutional affiliations.



Copyright: © 2022 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (<https://creativecommons.org/licenses/by/4.0/>).

1. Introduction

Dual-purpose Simmental cattle, a popular breed, exhibit high milk and meat production, good fertility and profitability [1–4]. Dual-purpose Simmental cattle were introduced to Northwest China in the 1950s. Core breeding farms have resulted in improvements in more than 400,000 cattle in China over several years [5]. These core breeding farms record milk production, milk composition, body measurement, and body weight every month. The breeding goals for dual-purpose Simmental cattle in Northwest China are milk production, milk quality, and growth traits. The average 305-day milk yield reached

5469 kg, the average fat content was 4.13%, the protein content was 3.33% [3], and the average daily weight gain from birth to 24 months of age was $0.70 \text{ kg}\cdot\text{d}^{-1}$ [6].

Body conformation traits are closely related to economically important traits, such as milk production [7], reproduction [8], health [9], profitability [10], and lifespan [11]. Therefore, studying the genetics of body conformation traits is as important as other production traits from an economic perspective. Understanding the genetic parameters of body conformation traits is crucial for implementing breeding programmes. Multiple regression has been used to analyze the relationship between body conformation traits, and the results have shown that the traits are correlated genetically and phenotypically [12]. For instance, Simčič et al. [12] found high genetic and phenotypic correlations between body frame traits in first parity Rendena cows. In some other cattle breeds, VanRaden et al. [13] and Mazza et al. [14], found strong genetic correlations between rear udder height and rear udder width, with values ranging from 0.85 to 0.95. Using a large number of traits with common information in multiple regression can lead to biased estimates of their relationship with productive traits [15]. Factor analysis is a useful multivariate technique for analyzing correlated traits, and it can remove redundant information introduced by incorporating multiple variables [16,17]. Mazza et al. [18] and Olasege et al. [19] suggested the use of latent factors in genetic evaluation, avoiding the analysis of highly correlated traits, thereby improving precision and reducing computational burdens for large datasets.

To date, a linear identification of body conformation traits of dual-purpose Simmental cattle in Northwest China has not been conducted, and breeding programmes do not include linear-type traits. This study aims to estimate genetic parameters for body conformation traits in dual-purpose Simmental cattle using factor analysis. The results could be used to develop a national genetic evaluation framework for the improvement of body conformation traits for dual-purpose Simmental cattle in China.

2. Materials and Methods

Approval from the local Institutional Animal Care and Use Committee was not required for this study because the data were obtained by field measurements, and no animal experiments were conducted.

2.1. Phenotypic and Pedigree Data

In 2020, our team applied for a dual-purpose cattle type classification project. After deliberation, we identified 6 composite traits and 27 body conformation traits. There were 17 measured traits and 10 scored traits. The measured traits included stature (ST), body depth (BD), chest width (CW), withers width (WW), hind leg half circumference (HLHC), rear leg height (RLH), rump length (RL), rump width (RW), rump angle (RA), rear udder height (RUH), rear udder width (RUW), median suspensory (MS), udder depth (UD), fore udder length (FUL), front teat length (FTL), front teat diameter (FTD), and heel depth (HD). The scored traits included ribs and bone (RB), rear legs side view (RLSV), bone quality (BQ), foot angle (FA), rear legs rear view (RLRV), fore udder attachment (FUA), rear udder length (RUL), udder balance (UB), fore teat placement (FTP), and rear teat placement (RTP). The individual body conformation traits are specific to certain body regions of the animals, including the body frame, muscularity, rump, feet and legs, and mammary system. The six composite traits summarized body frame, muscularity, rump, feet and legs, mammary system, and final score.

Conformation traits of 1200 dual-purpose Simmental cattle born from 2010 to 2019 were measured at Xinjiang Hutubi Farm, Kekedala Chuangjin Farm, and Xinjiang Haozi Animal Husbandry Farm in Northwest China. After quality control was applied using the threshold of the mean \pm three times the standard deviation, 1016 conformation records remained and were used for analysis. The pedigree file used for the analysis included data for 1988 animals, with each animal being traced back three generations. In the full datasets, one sire had a maximum of 195 offspring with records, whereas 16 sires had only one offspring. More than 62 dams had 2 or more offspring.

2.2. Genotype Data

The Illumina 100K Bovine BeadChip was used to genotype 516 Simmental cows. For analysis, common SNPs were obtained from 100K bead chips as target files. Quality control of the SNP genotyping was carried out with PLINK 1.07 software (Boston, MA, USA) [20]. All genotyped animals had a call rate greater than 0.90. SNPs were removed if the call rate was less than 0.90 and the minor allele frequency (MAF) was less than 0.01. After quality control, data for 88,913 SNPs from 516 animals remained for analysis.

2.3. Genetic Connectedness

To estimate genetic parameters, the genetic connectedness of the animals in the dataset must be determined. If the breeding process affects the genetic connectedness of cattle at different farms over time, this change would be reflected in a change in the average relatedness in birth-year cohorts and among the populations at different farms. The indirect method of Sargolzaei [21], implemented in the software package CFC, was used to compute the coefficient of relationship among animals.

2.4. Variance Component Estimates for Body Conformation Traits

The single-trait animal model was used to estimate genetic and residual variance for the 6 composite traits and 27 individual body conformation traits with the average information-restricted maximum likelihood (AI-REML) method. The AIREMLF90 procedure of BLUPF90 1.0.1 software (Athens, GA, USA) was used [22]. The animal linear mixed model for the single trait analysis was as follows:

$$Y = X\beta + Za + e \quad (1)$$

where Y indicates the vector of 6 composite traits and 27 individual body conformation traits; β is the vector of the fixed effects, including herd-year of evaluation (5 different levels), days in milk (ten classes: from 10 to 30 days after calving, from 31 to 270 days after calving in 30-day intervals, and >270 days after calving), age at first calving (seven classes: <23 months, from 23 to 34 months in 2-month intervals, and >34 months), and parity (four classes: 1, 2, 3, and ≥ 4); a is the vector of random animal additive genetic effects; e is the vector of random residual effects; and X , and Z are the incidence matrices assigning observations to fixed and random animal effects.

The genetic effect was modeled using two kinds of genetic variance-covariance matrices: the pedigree relationship matrix (A) [23], and the combined genomic-pedigree relationship matrix (H) [24].

The heritability and standard error were estimated according to Austin Putz et al. [25] using the following formula:

$$h^2 = \frac{\sigma_a^2}{\sigma_a^2 + \sigma_e^2} \quad (2)$$

$$SE(h^2) = \left(\frac{h^2(1-h^2)}{a} - \frac{h^2(h^2)}{a} \right) \begin{pmatrix} var(a) & cov(a,e) \\ cov(e,a) & var(e) \end{pmatrix} \begin{pmatrix} \frac{h^2(1-h^2)}{a} \\ -\frac{h^2(h^2)}{a} \end{pmatrix} \quad (3)$$

2.5. Factor Analysis

Factor analysis was performed using the FACTOR procedure in SAS 8.0 software (Cary, NC, USA). In this analysis, a set of n observation variables (y_1, \dots, y_n) is synthesized into a new set of p ($p < n$) latent variables (X_1, \dots, X_p), which are referred to as common latent factors. As described by Kaiser (1960), varimax rotation was used to maintain the orthogonality of the extracted factors. Only components with eigenvalues ≥ 1 were retained for the analyses (i.e., the Kaiser criterion; Russel [26], Mazza et al. [18], Olasege et al. [19]). By observing the individual body conformation trait loadings, the analysis was interpreted from a biological point of view. Based on the standardized scoring coefficients, we calculated the sample scores for each animal. According to Russel [27], the classic factor analysis

equation specifies that a measure being factored can be represented by the following equation accounting for n factors:

$$X_m = W_{m1}F_1 + W_{m2}F_2 + \dots + W_{mn}F_n + W_{mn}U_n + e_m \quad (4)$$

where X_m is the m -th measure, F_n is the n -th common factor that underlies the m -th measure being analyzed, and U_n is the n -th factor that is unique to each m -th measure. Furthermore, W_{mn} represents the n -th factor loading coefficients or loadings of each m -th measure on the respective factors, whereas e_m reflects the random measurement errors in each m -th measure. Using this equation, we can divide the variance in the measure being factored into three parts. The first part of the variance of the measure reflects the impact of the common factors, the second part reflects the influence of the unique factor associated with the measure, and the third is the variance of the random error [26].

2.6. Estimation of Genetic Parameters Using Factor Analysis

In the single-trait animal model presented above, genetic parameters were estimated by fitting the factor scores as y . The estimated breeding values (EBVs) of the factor scores were then subjected to rank correlation analysis with the EBVs of the composite traits using SPSS.

3. Results

3.1. Phenotype

The descriptive statistics for the 6 composite traits and 27 individual body conformation traits, including the mean, standard deviation, minimum, maximum, and coefficient of variation are summarized in Table 1. In the composite traits, the coefficient of variation ranged from 2.52% (final score) to 7.73% (rump). For the individual body conformation traits, the coefficient of variation ranged from 3.20% (stature) to 77.16% (udder depth).

Table 1. Description of body conformation traits in dual-purpose Simmental Cattle.

Traits	Number	Minimum	Maximum	Average	SD	CV (%)
Composite trait						
Final score (points)	1016	74.6	87.1	82.2	2.1	2.5
Body frame (points)	1016	68.5	95.0	85.1	4.6	5.4
Muscularity (points)	1016	70.5	91.5	80.7	3.1	3.9
Rump (points)	1016	58.0	95.0	80.4	6.2	7.7
feet and legs (points)	1016	73.0	94.3	86.7	3.6	4.2
Mammary system (points)	1016	64.2	88.1	78.8	4.2	5.4
Individual body conformation trait						
Body frame						
Stature (cm)	1016	126	154	140.7	4.5	3.2
Body depth (cm)	1016	61	90	78.4	6.0	7.6
Chest width (cm)	1016	18	39	27.6	4.3	15.6
Muscularity						
Withers width (cm)	1016	10	28	18.4	2.9	15.6
Hind leg half circumference (cm)	1016	33	52	42.5	3.0	6.9
Rear leg height (cm)	1016	62	88	76.6	4.3	5.6
Rib and bone (points)	1016	1	9	6.8	1.2	18.0
Rump						
Rump length (cm)	1016	43	59	51.8	3.1	5.9
Rump width (cm)	1016	17	29	22.7	1.9	8.4
Rump angle (cm)	1016	−5	17	6.3	3.3	52.7
Feet and legs						
Heel depth (cm)	1016	1	8	4.1	0.9	21.7
Foot angle (points)	1016	2	9	5.4	1.1	20.0
Rear legs side view (points)	1016	2	9	5.2	1.2	22.4
Bone quality (points)	1016	4	8	6.1	0.6	9.8

Table 1. *Cont.*

Traits	Number	Minimum	Maximum	Average	SD	CV (%)
Rear legs rear view (points)	1016	2	9	5.2	1.2	22.4
Mammary system						
Rear Udder height (cm)	1016	14	38	29.0	4.2	14.4
Rear Udder width (cm)	1016	7	24	13.0	2.9	22.4
Median suspensory (cm)	1016	0	8	3.5	1.5	42.9
Udder depth (cm)	1016	−19	22	6.7	5.2	77.2
Fore udder length (cm)	1016	10	29	17.7	3.4	19.2
Front teat length (cm)	1016	2	10	4.8	1.3	27.2
Front teat diameter (cm)	1016	1	4	2.5	0.5	21.3
Fore udder attachment (points)	1016	1	8	4.2	1.3	32.0
Rear udder length (points)	1016	1	9	5.0	1.5	29.7
Udder balance (points)	1016	1	9	4.7	0.9	19.0
Fore teat placement (points)	1016	1	8	4.0	1.1	26.4
Rear teat placement (points)	1016	1	9	5.2	1.0	19.5

3.2. Genetic Connectedness

All three farms use artificial insemination for breeding. Figure 1 shows the trend of the average coefficient of relationship by year of birth of Simmental cattle. The relationship coefficient among Simmental cattle born on the three farms from 2010 to 2019 ranged from 0.0179 to 0.0613, with irregular variation. Of the annual changes in the coefficient of relationship, the largest change was from 2017 to 2018.

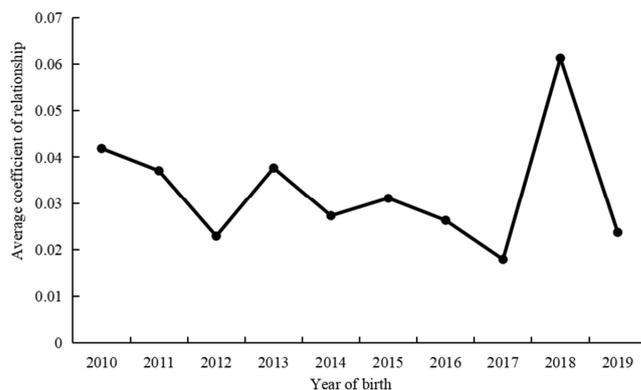


Figure 1. Evolution of average coefficient of relationship between three farms subpopulations according to the year of birth of Simmental cattle.

3.3. Heritability of Conformation Traits

The heritability estimates for the 6 composite traits and 27 individual body conformation traits are presented in Table 2. Using the pedigree relationship matrix (A), the estimates for the composite traits ranged from 0.07 (muscularity composite) to 0.43 (body frame composite); using the combined genomic-pedigree matrix (H), the estimates for the composite traits ranged from 0.10 (muscularity composite) to 0.47 (body frame composite). The estimation for the final score was 0.18 from A and 0.14 from H. In general, the highest estimates of heritability were obtained in body frame traits, whereas the estimates for muscularity traits and feet and legs traits were low. For the individual body conformation, the heritability estimates from A ranged from 0.05 (HLHC and RAB) to 0.56 (ST), and from H, they ranged from 0.03 (rear udder length; RUL) to 0.65 (ST). The standard errors of heritability estimates were all ≤ 0.10 , except for those for ST and RL. In addition, estimates of other body conformation traits from A and H were very similar. There was little improvement in the accuracy of the estimated heritability, i.e., the standard errors of the two models were similar. However, the average heritability estimated by H was higher than that of A, except for individual muscularity traits.

Table 2. Variance components and heritability of body conformation traits in dual-purpose Simmental Cattle obtained using the pedigree relationship matrix (A) and combined genomic-pedigree relationship matrix (H).

Traits ¹	A Matrix				H Matrix			
	σ_a^2	σ_e^2	σ_p^2	$h^2 \pm SE$	σ_a^2	σ_e^2	σ_p^2	$h^2 \pm SE$
	Composite trait							
Final score	0.72	3.33	4.05	0.18 ± 0.08	0.55	3.50	4.05	0.14 ± 0.07
Body frame	8.47	11.27	19.74	0.43 ± 0.13	9.36	10.48	19.84	0.47 ± 0.12
Muscularity	0.62	8.86	9.48	0.07 ± 0.05	0.99	8.53	9.52	0.10 ± 0.06
Rump	7.57	23.94	31.51	0.24 ± 0.10	6.65	24.74	31.39	0.21 ± 0.09
Feet and leg	1.29	11.37	12.66	0.10 ± 0.06	1.27	11.41	12.68	0.10 ± 0.06
Mammary system	2.74	10.63	13.37	0.20 ± 0.09	3.11	10.33	13.44	0.23 ± 0.09
Average	3.57	11.57	15.14	0.20 ± 0.09	3.66	11.50	15.15	0.21 ± 0.08
	Individual body conformation							
Body frame								
ST	11.38	8.93	20.31	0.56 ± 0.12	13.48	7.17	20.66	0.65 ± 0.11
BD	5.27	25.33	30.60	0.17 ± 0.08	6.00	24.70	30.70	0.20 ± 0.09
CW	2.29	15.72	18.01	0.13 ± 0.08	2.04	15.96	18.00	0.11 ± 0.07
Average	6.31	16.66	22.97	0.29 ± 0.09	7.17	15.94	23.12	0.32 ± 0.09
	Muscularity							
WW	0.77	7.51	8.28	0.09 ± 0.07	0.53	7.73	8.26	0.06 ± 0.06
HLHC	0.53	7.55	8.08	0.07 ± 0.05	0.64	7.45	8.09	0.08 ± 0.05
RLH	0.86	14.95	15.81	0.05 ± 0.05	0.94	14.88	15.82	0.06 ± 0.05
RAB	0.07	1.29	1.36	0.05 ± 0.05	0.52	1.31	1.83	0.04 ± 0.04
Average	0.56	7.83	8.38	0.07 ± 0.06	0.66	7.84	8.50	0.06 ± 0.05
	Rump							
RL	2.24	5.57	7.81	0.29 ± 0.11	2.67	5.23	7.90	0.34 ± 0.12
RW	0.68	2.40	3.08	0.22 ± 0.09	0.62	2.46	3.08	0.20 ± 0.08
RA	1.39	7.85	9.24	0.15 ± 0.07	1.94	7.35	9.29	0.21 ± 0.08
Average	1.44	5.27	6.71	0.22 ± 0.09	1.74	5.01	6.76	0.25 ± 0.09
	Feet and legs							
HD	0.03	0.59	0.62	0.05 ± 0.05	0.05	0.58	0.63	0.07 ± 0.05
FA	0.11	0.83	0.94	0.11 ± 0.06	0.15	0.78	0.93	0.16 ± 0.07
RLSV	0.11	1.18	1.29	0.09 ± 0.06	0.09	1.20	1.29	0.07 ± 0.05
BQ	0.02	0.32	0.34	0.07 ± 0.05	0.04	0.30	0.34	0.12 ± 0.06
RLRV	0.19	1.34	1.53	0.12 ± 0.07	0.20	1.33	1.53	0.13 ± 0.07
Average	0.09	0.85	0.94	0.09 ± 0.06	0.11	0.84	0.94	0.11 ± 0.06
	Mammary system							
RUH	3.08	14.72	17.80	0.17 ± 0.07	3.00	14.82	17.82	0.17 ± 0.07
RUW	0.57	7.06	7.63	0.07 ± 0.05	0.83	6.81	7.64	0.11 ± 0.06
MS	0.11	1.73	1.84	0.06 ± 0.05	0.18	1.67	1.85	0.10 ± 0.06
UD	4.01	14.15	18.16	0.22 ± 0.09	4.71	13.55	18.26	0.26 ± 0.09
FUL	1.02	7.91	8.93	0.11 ± 0.06	1.08	7.87	8.95	0.12 ± 0.07
FTL	0.21	1.57	1.78	0.12 ± 0.06	0.20	1.58	1.78	0.11 ± 0.06
FTD	0.05	0.23	0.28	0.18 ± 0.07	0.05	0.23	0.28	0.18 ± 0.07
FUA	0.31	1.34	1.65	0.19 ± 0.08	0.44	1.23	1.67	0.27 ± 0.09
RUL	0.14	1.75	1.89	0.07 ± 0.06	0.05	1.83	1.88	0.03 ± 0.05
UB	0.22	0.61	0.83	0.26 ± 0.10	0.28	0.54	0.82	0.34 ± 0.09
FTP	0.19	0.78	0.97	0.20 ± 0.08	0.20	0.78	0.98	0.20 ± 0.08
RTP	0.16	0.76	0.92	0.17 ± 0.08	0.20	0.72	0.92	0.22 ± 0.09
Average	0.84	4.38	5.22	0.15 ± 0.07	0.94	4.30	5.24	0.18 ± 0.07

¹ ST: stature; BD: body depth; CW: chest width; WW: withers width; HLHC: hind leg half circumference; RLH: Rear leg height; RAB: Rib and bone; RL: rump length; RW: rump width; RA: rump angle; HD: heel depth; FA: foot angle; RLSV: rear legs side view; BQ: Bone quality; RLRV: rear legs rear view; RUH: rear udder height; RUW: rear udder width; MS: median suspensory; UD: udder depth; FUL: fore udder length; FTL: fore teat length; FTD: fore teat diameter; FUA: fore udder attachment; RUL: rear udder length; UB: udder balance; FTP: fore teat placement; RTP: rear teat placement; σ_a^2 : additive genetic variance; σ_e^2 : residual variance; σ_p^2 : phenotype variance; h^2 : heritability; SE: standard error.

3.4. Factor Analysis

The eigenvalues and the proportion of total and cumulative variance explained by each factor are listed in Table 3. The 24 factors after varimax rotation explained 96.13% of the total variation among the 27 individual body conformation traits. The first factor (F1) accounted for the largest proportion (13.51%) of the total variability. The first 9 factors with eigenvalues ≥ 1 were retained for further analysis. The varimax rotated factor patterns coefficients and communalities are reported in Table 4. Only loading coefficients $\geq |0.40|$ [27] were reported for each body conformation trait. F1 was heavily loaded for ST (0.65), BD (0.46), HLHC (0.60), RL (0.72), and RW (0.57). F1 accounted for traits belonging to the body frame, muscularity, and rump. Factor 2 (F2), with 8.13% variability, had higher loadings on RLH (0.46), HD (0.76), and FA (0.77), whereas the third factor (F3) accounted for 7.32% of the proportional variability and loaded heavily on FTP (0.79) and RTP (0.73). The fourth factor (F4) explained 6.01% of the variability and loaded heavily on FTL (0.80) and FTD (0.81), whereas the fifth factor (F5) explained 5.18% of the variability and had higher loadings on FUL (0.70). F6 accounted for 4.93% of the proportional variability and higher loadings for UB (0.65), whereas F7 accounted for 4.99% of the proportional variability and had higher loadings for MS (0.66). F8 accounted for 4.18% of the proportional variability and loaded heavily on WW (−0.44) and RLRV (0.79), whereas F9 accounted for 3.80% of the proportional variability and only loaded on RA (0.83). Because the remaining subsequent factors with eigenvalues < 1 explained a small amount of variance, they were not considered for further analysis. The range of communality of variables for the 27 body conformation traits was between 0.39 and 0.72.

Table 3. Eigenvalues and proportion of total and cumulative variance explained by factor analysis of the body conformation traits in dual-purpose Simmental cattle.

Factor	Eigenvalue	Proportional Variance (%)	Cumulative Variance (%)
F1	3.65	13.51	13.51
F2	2.20	8.13	21.65
F3	1.98	7.32	28.96
F4	1.62	6.01	34.97
F5	1.40	5.18	40.15
F6	1.33	4.93	45.08
F7	1.21	4.49	49.57
F8	1.13	4.18	53.75
F9	1.03	3.80	57.55
F10	0.96	3.56	61.12
F11	0.93	3.46	64.58
F12	0.89	3.30	67.88
F13	0.86	3.18	71.05
F14	0.81	3.01	74.06
F15	0.74	2.74	76.81
F16	0.73	2.73	79.54
F17	0.71	2.65	82.19
F18	0.67	2.47	84.66
F19	0.62	2.29	86.94
F20	0.59	2.17	89.11
F21	0.56	2.09	91.20
F22	0.49	1.81	93.00
F23	0.43	1.60	94.60
F24	0.41	1.52	96.13

Table 4. Latent factors, loading of individual body conformation traits (loading coefficients $\geq |0.40|$), and communality after varimax rotation of the 27 body conformation traits in dual-purpose Simmental cattle.

Trait ¹	Varimax Latent Factors									Communality
	F1	F2	F3	F4	F5	F6	F7	F8	F9	
ST	0.65									0.68
BD	0.46						−0.51			0.68
CW							−0.45			0.39
WW								−0.44		0.54
HLHC	0.60									0.43
RLH		0.46								0.54
RAB					−0.42					0.44
RL	0.72									0.62
RW	0.57									0.49
RA									0.83	0.72
HD		0.76								0.64
FA		0.77								0.67
RLSV										0.44
BQ										0.46
RLRV								0.79		0.68
RAH						−0.53				0.52
RUW					0.58					0.60
MS							0.66			0.51
UD						0.50				0.63
FUL					0.70					0.60
FTL				0.80						0.68
FTD				0.81						0.72
FUA						0.53				0.51
RUL					0.50					0.55
UB						0.65				0.48
FTP			0.79							0.63
RTP			0.73							0.68
Variance explained (%)	2.32	1.98	1.91	1.85	1.69	1.68	1.52	1.33	1.26	

¹ ST: stature; BD: body depth; CW: chest width; WW: withers width; HLHC: hind leg half circumference; RLH: Rear leg height; RAB: Rib and bone; RL: rump length; RW: rump width; RA: rump angle; HD: heel depth; FA: foot angle; RLSV: rear legs side view; BQ: Bone quality; RLRV: rear legs rear view; RUH: rear udder height; RUW: rear udder width; MS: median suspensory; UD: udder depth; FUL: fore udder length; FTL: fore teat length; FTD: fore teat diameter; FUA: fore udder attachment; RUL: rear udder length; UB: udder balance; FTP: fore teat placement; RTP: rear teat placement.

3.5. Heritability of Factor Scores

Variance components for nine different factor scores using the different relationship matrices are shown in Table 5. Heritability estimates for the pedigree relationship matrix had a mean value of 0.18 with a standard error of 0.08, whereas for the combined genomic-pedigree matrix, the mean value of heritability was 0.20 with a standard error of 0.08 for all considered factor scores. In particular, the lowest heritability estimates were for F2 (feet and legs factor score and muscularity factor score) and F4 based on both methods. However, for both matrices, the highest values of heritability observed were for F1, a factor score accounting for the body frame and rump individual body conformation traits. Factor 3 and Factor 5 (i.e., the mammary system factor score) exhibited medium heritability values based on both matrices. In general, there was no significant difference in the heritability estimates of factor scores obtained using the pedigree relationship matrix or the combined genomic-pedigree matrix.

Table 5. Estimated variance components, heritability and standard errors for 9 factor scores obtained using the pedigree relationship matrix (A) and combined genomic-pedigree relationship matrix (H).

Factor Score	A Matrix				H Matrix			
	σ_a^2	σ_e^2	σ_p^2	$h^2 \pm SE$	σ_a^2	σ_e^2	σ_p^2	$h^2 \pm SE$
F1	0.39	0.52	0.91	0.43 ± 0.13	0.39	0.52	0.91	0.43 ± 0.12
F2	0.04	0.64	0.68	0.06 ± 0.05	0.07	0.62	0.69	0.10 ± 0.06
F3	0.16	0.61	0.77	0.21 ± 0.08	0.16	0.62	0.78	0.21 ± 0.09
F4	0.008	0.92	0.93	0.008 ± 0.04	0.04	0.90	0.94	0.04 ± 0.05
F5	0.20	0.71	0.91	0.22 ± 0.10	0.24	0.68	0.92	0.26 ± 0.11
F6	0.09	0.70	0.79	0.11 ± 0.07	0.11	0.68	0.79	0.14 ± 0.08
F7	0.06	0.60	0.66	0.09 ± 0.05	0.07	0.59	0.66	0.10 ± 0.06
F8	0.08	0.58	0.66	0.12 ± 0.06	0.10	0.57	0.67	0.15 ± 0.07
F9	0.31	0.58	0.89	0.35 ± 0.12	0.34	0.55	0.89	0.38 ± 0.12

3.6. Correlations between EBV of Composite Traits and EBV of Factor Scores

Results of the Spearman and Pearson correlation analyses (only absolute values ≥ 0.20 reported) between EBV of composite traits and EBV of factor scores are reported in Tables 6 and 7. The results of the Spearman and Pearson correlation analyses were very similar. Correlation coefficients exhibited patterns very similar to the loading coefficients of the individual body conformation traits for F1, F2, and F4. EBVs obtained for F1 were highly positively correlated with the EBVs of the body frame, muscularity, and rump traits. In addition, Spearman and Pearson correlations between EBVs of F2 and EBVs of muscularity and feet and legs-related traits were positive, consistent with results previously reported for the loading coefficients between individual traits and the second latent factor. This pattern was also observed for F4; Spearman and Pearson correlations between EBVs of F4 and EBVs of mammary system-related traits were positive. However, similar results were not observed for F3, F5, F6, F7, F8, and F9.

Table 6. Spearman correlation coefficients (only values $\geq |0.20|$) between EBVs estimated for body conformation traits and EBVs obtained from 9 factor scores (from F1 to F9) in dual-purpose Simmental cattle.

Composite Traits	F1	F2	F3	F4	F5	F6	F7	F8	F9
Body frame	0.67			−0.31			0.28		
Muscularity	0.55	0.24		0.23	0.20	−0.24	0.45		
Rump	0.58		0.24		0.23		0.44		
feet and legs		0.20							
Mammary system				0.30				0.31	0.44

Table 7. Pearson correlation coefficients (only values $\geq |0.20|$) between EBVs estimated for body conformation traits and EBVs obtained from 9 factor scores (from F1 to F9) in dual-purpose Simmental cattle.

Composite Traits	F1	F2	F3	F4	F5	F6	F7	F8	F9
Body frame	0.69			−0.34			0.36		
Muscularity	0.59	0.31		0.23	0.22	−0.25	0.58		
Rump	0.64		0.33		0.28		0.51		
feet and legs		0.33							
Mammary system				0.39				0.32	0.50

4. Discussion

In this study, the heritability of 6 composite traits and 27 individual body conformations ranged from 0.03 to 0.65. The 24 common latent factors explained 96.13% of the total variation in 27 individual body conformation traits. Heritability estimates for the factor scores ranged from 0.008 to 0.43. The Spearman and Pearson correlation results revealed

that the correlation coefficients between the EBVs of the factor scores and the EBVs of the composite traits exhibited a very similar pattern to that of the loading coefficients of the individual body conformation traits for the F1, F2, and F4.

4.1. Phenotype

Linear scoring of dual-purpose Simmental cattle has not been conducted in Northwest China. We quantified the body conformation traits that could be reliably measured to develop a linear scoring criterion. However, some individual body conformation traits were difficult to measure, such as rib and bone, rear leg rear view, and fore udder attachment. We scored ten hard-to-measure individual body conformation traits on a 9-point linear scale. The means of all scored traits ranged from 4 to 6. Similar findings were reported by Strapáková et al. [28] and Zavadilová et al. [29]. In addition, similar findings have been reported for Chinese Holstein cattle [19], US Brown Swiss dairy cattle [9], and Rendena and Aosta Red Pied dual-purpose breeds [18].

The mean values for body frame (85.14) and feet and legs scores (86.72) were slightly higher than those for Slovenian Simmental dairy cows (81.35 and 81.06, respectively). In contrast, scores for muscularity (80.72) and the mammary system (78.77) were similar for the two species [28]. For body measurement traits, the mean values of stature (140.65 cm) and body depth (78.41 cm) were slightly lower than those of Slovakia Simmental dairy cows (144.31 cm and 82.92 cm, respectively) [28] and slightly higher than those of Czech Fleckvieh cows (137.40 cm and 77.40 cm, respectively) [29]. In addition, the rump length (51.83 cm) of dual-purpose Simmental cattle in Northwest China was lower than that of Slovakia Simmental dairy cows (53.31 cm) [29] and Czech Fleckvieh cows (52.80 cm) [28]. In general, the body size of dual-purpose Simmental cattle in northwest China needs to be improved.

4.2. Heritability

In this study, except for muscularity traits, the composite traits were consistent with those of dual-purpose Rendena cattle [14], dual-purpose autochthonous Valdostana cattle [30], German and French dairy cattle [31], and Czech Fleckvieh cattle [32]. Previous studies [32,33] reported muscularity to be a medium to high heritability trait; this differs from the conclusion based on the results of our study and may be due to different definitions of muscularity. The heritability of individual body frame traits was high (0.11 to 0.65), followed by the rump traits (0.15 to 0.34) and mammary system traits (0.03 to 0.34), and the heritability was low for feet and legs traits (0.07 to 0.16) and muscularity traits (0.04 to 0.09). As reported by Kern et al. [34], Gibson et al. [9], and Spehar et al. [35], RUW and MS are moderately heritable traits (0.12–0.17) in Brazilian Holstein cattle, American Brown Swiss cattle, and Slovenian Brown Swiss cattle, while the results of this study indicated low heritability of these traits (0.04–0.06). Roveglia et al. [36] reported a heritability of 0.07 for RUW in their study of Italian Jersey cattle, which was similar to the results of our study (0.07). The differences in magnitude observed across these studies may be due to the scales used for measurement and scoring, the number of animals, breeds, statistical models, data editing procedures, and consistency among evaluators [37].

In addition, we investigated the heritability for composite traits and individual body conformation traits using the H matrix in dual-purpose Simmental cattle. Comparing the estimates of heritability and their standard errors for each trait using A and H, there was no significant difference for any conformation trait, except for ST. Based on the standard errors, there were no significant differences in any traits. It is possible that the construction of H was primarily determined by the information found in A since there were too few individuals with genotypes. Therefore, they had little influence on the genetic parameter estimates. In addition, the small number of individuals with phenotypic and pedigree records, and the specific environmental effects on farms have influenced the results of the genetic parameter estimation in this study. Some researchers have shown that genomic information can improve the accuracy of genetic parameter estimation for breeding target

traits. For example, Veerkamp et al. [38] and Wei et al. [39] demonstrated that rescaling H according to the eigenvalues of A slightly changed the genetic variances. However, there are several reasons why estimated genetic variances differ between the models using pedigree and genomic relationships. First, A and H use different scales for the diagonal elements, especially when considering the Mendelian sampling component in H. A second reason is related to the genetic structure of cattle populations. The third reason is the accuracy of pedigree records. Naserkheil et al. [40] and Song et al. [41] demonstrated that single-step GBLUP provides a more accurate prediction than traditional BLUP for all the studied traits. There are so few individuals with available genotypes in the current study that combining data from genotyped and nongenotyped animals are not worthwhile. In future studies, additional individuals with available genotypes will be included in the genomic analysis of conformation traits in dual-purpose Simmental cattle.

4.3. Factor Analysis

In multivariate statistical analysis, factor analysis is one of the classical tools [18,19,42,43]. Several studies have shown that a small number of factors can be used to accurately describe the cow's conformation without reducing accuracy [18,19,44,45]. An essential aspect of the present study is the algebraic sign and magnitude of the loading coefficients and the percentage of the total variance explained by each factor. A trait with a high loading coefficient contributes more to the factor than one with a low loading coefficient [44]. Once the loading coefficients are determined, with a varimax rotation, it is possible to posit a biological interpretation of the factors [46]. Kaiser introduced the varimax rotation criterion; it maximizes the sum of variances between variables and factors squared [47]. Generally, factor analysis can be understood as a data-reduction technique that removes duplicate information from a collection of correlated variables [43]. The most frequently used factor analysis procedure in the literature has been the matrix transformation step, followed by the extraction of all factors using the principal-factor method with eigenvalues ≥ 1.0 and then rotation of these factors by varimax [18,19,45,48]. Finally, the factors can be explained by identifying the traits with the largest values. Since the procedure is available in many computer statistical packages (e.g., SAS and SPSS), it is relatively easy to use.

Phenotypic factor analysis in this study extracted 24 principal components, which accounted for 96.13% of the total variance among the 27 body conformation traits. Chu and Shi [43] found that eigenvalues > 1 explained 49.1% of the total variance in type traits of Holstein cows in the Beijing area. A similar value was found for the first six latent factors in a study of Aosta Red Pied cattle [18]. The approximate percentage of the total variance explained was determined in a factor analysis of the Rendena breed conducted by Mantovani et al. [48]. The results of the analysis indicated that high values of F1, representing body frame, rump, and muscularity traits, were associated with greater height and buttocks size. High correlations were also noted in previous studies [49,50]. As reported by Manafizar et al. [51], ST, CW, and BD exhibited a strong genetic correlation with residual feed intake, and the work of Dadati et al. [52] indicated that the rump score had the highest genetic correlation with easier calving. Therefore, selection based on F1 may improve the milk production, meat production and reproductive performance of dual-purpose cattle. F2 is characterized by high and positive loading coefficients for heel depth and foot angle and is usually associated with lameness. Thus, the selection of dual-purpose cattle based on high scores for feet and legs, steeper foot angle, straighter legs, and fine bone structure might improve locomotion and lower the risk of claw disorder [53,54]. F3, F4, F5, and F6 were udder trait-related factors, indicating the size and quality of the mammary system, respectively. Mazza et al. [18] reported that F3 and F4 reflected the mammary system in dual-purpose autochthonous breeds. High values of F3 were associated with teat placement, high values of F4 were associated with thick and long teats, high values of F5 were related to large udders, and high values of F6 were associated with shallow, strong, and balanced udders. Researchers have extensively studied the genetic correlation between conformation traits and SCS [9,55–57]. The strongest correlations were obtained

for FUA, FTP, and UD. F3, F4, F5, and F6 included traits that are usually associated with SCS, and a selection index based on higher udders with tighter attachments and closer teats would be favorable for reducing SCS [58]. Dube et al. [59] found that narrow teat placement and low, shallow udders were strongly correlated with low SCC in the South African Holstein population. Evaluating latent factors instead of original traits is an interesting approach. However, in reality, there is almost no routine application of conducting a genetic evaluation on the factors in any country.

4.4. Correlations between EBV of Composite Traits and EBV of Factor Scores

Consistent with the findings of Mantovani et al. [48], the Spearman correlation (r_s) analyses between composite trait EBVs and factor score EBVs showed very similar patterns to the loading coefficients of individual traits in latent factors. For instance, the EBVs obtained for F1 exhibited a high correlation with the EBVs of body frame, muscularity, and rump traits ($0.55 < r_s < 0.67$ and $0.59 < r < 0.69$). Additionally, correlation between EBVs of F2 and EBVs of feet and legs also showed positive correlations ($r_s = 0.20$ and $r = 0.33$), and EBVs for mammary system and udder conformation factors (i.e., F4) also showed positive correlations ($r_s = 0.30$ and $r = 0.39$). The Spearman and Pearson correlation results for F7 and the composite trait EBVs were similar to those of F1, exhibiting high correlations with body frame, muscularity, and rump. Because of the generally high Spearman and Pearson correlations between the EBVs of factor scores and the respective EBVs of body conformation traits associated with those factors, factor scores could be used to guide animal breeding. However, it is crucial to select factors prudently since the random error could attenuate any further analysis based on the newly extracted variable in the factor score [26]. Heritability estimates of the nine factor scores showed that in both matrices, the most heritable factor was related to body frame, muscularity, and rump traits (F1), whereas the least heritable factor was related to feet and legs traits (F2).

The amount of phenotypic and genotypic data collected was small due to the late start of the linear assessment of body conformation and genomic evaluation of dual-purpose Simmental cattle in Northwest China. In the future, we will establish a protocol for type classification in dual-purpose Simmental cattle and conduct annual body conformation linear identification and incorporate it into the breeding programme.

5. Conclusions

In this study, the body conformation traits showed a range of heritability from low to high, with stature yielding the highest estimates. The number of animals with recorded body conformation traits is rather low, especially the number of genotyped animals, which led to little difference in the precision of estimating heritability using the pedigree relationship matrix and the combined genomic-pedigree matrix. The factor scores exhibited low to medium heritability, and the generally high Spearman and Pearson correlations between EBVs of Factor 1, Factor 2, and Factor 4 and the corresponding EBVs for composite traits suggested their utility in selection programmes. These analyses suggest that a few factors can describe a variety of body conformation traits without reducing the accuracy of genetic assessments.

Author Contributions: Conceptualization, X.H. and Y.W.; methodology, H.L. (Hanpeng Luo); formal analysis, L.X. and X.Z.; investigation, L.X., X.Z., M.Z., T.Z., M.Y. and X.T.; resources, J.G.; data curation, L.X.; writing—original draft preparation, L.X., H.L. (Haibo Lu) and X.Z.; writing—review and editing, L.X., H.L. (Haibo Lu), Y.W. and X.H.; supervision, X.H. All authors have read and agreed to the published version of the manuscript.

Funding: This work was funded by the National Key Research and Development Program (no. 2021YFD1200903). The study was also supported by China Agriculture Research System of MOF and MARA (no. CARS-36).

Institutional Review Board Statement: Not applicable.

Informed Consent Statement: Not applicable.

Data Availability Statement: The data that support the findings of this study are available from the corresponding author, L.X., upon reasonable request.

Acknowledgments: We greatly thank the staffs of Xinjiang Hutubi Farm, Kekedala Chuangjin Farm, and Xinjiang Haozi Animal Husbandry Farm for providing us with working conditions.

Conflicts of Interest: We certify that there are no conflicts of interest with any financial organizations regarding the material discussed in the manuscript.

References

1. Perišić, P.; Skalicki, Z.; Petrovic, M.; Bogdanović, V.; Ružić-Muslić, D. Simmental cattle breed in different production systems. *Biotechnol. Anim. Husbandry* **2009**, *25*, 315–326. [[CrossRef](#)]
2. Amaya, A.; Garrick, D.; Martínez, R.; Cerón-Muñoz, M. Economic values for index improvement of dual-purpose Simmental cattle. *Livest. Sci.* **2020**, *240*, 104224. [[CrossRef](#)]
3. Wei, C.; Luo, H.; Wang, Y.; Huang, X.; Zhang, M.; Zhang, X.; Wang, D.; Ge, J.; Xu, L.; Jiang, H.; et al. Analyses of the genetic relationships between lactose, somatic cell score, and growth traits in Simmental cattle. *Animal* **2021**, *15*, 100027. [[CrossRef](#)] [[PubMed](#)]
4. Edel, C.; Schwarzenbacher, H.; Hamann, H.; Neuner, S.; Emmerling, R.; Götz, K. The German-Austrian genomic evaluation system for Fleckvieh (Simmental) cattle. *Interbull Bull.* **2011**, *44*, 152–156.
5. Chen, W. *Estimation of Breeding Values in Simmental Cattle Using Animal BLUP Model*; Xinjiang Agricultural University: Urumqi, China, 2017.
6. Zhu, B.; Li, J.; Wang, C.; Xu, L.; Chen, Y.; Gao, X.; Zhang, L.; Gao, H.; Li, J. Genetic parameter and genetic gain estimation for growth and development traits in Chinese Simmental beef cattle. *Acta Vet. Zootech. Sin.* **2020**, *51*, 1833–1844.
7. Stefani, G.; El Faro, L.; Santana, M.; Tonhati, H. Association of longevity with type traits, milk yield and udder health in Holstein cows. *Livest. Sci.* **2018**, *218*, 1–7. [[CrossRef](#)]
8. Zink, V.; Štípková, M.; Lassen, J. Genetic parameters for female fertility, locomotion, body condition score, and linear type traits in Czech Holstein cattle. *J. Dairy Sci.* **2011**, *94*, 5176–5182. [[CrossRef](#)]
9. Gibson, K.D.; Dechow, C.D. Genetic parameters for yield, fitness, and type traits in US Brown Swiss dairy cattle. *J. Dairy Sci.* **2017**, *101*, 1251–1257. [[CrossRef](#)]
10. Pérez-Cabal, M.; Alenda, R. Genetic relationships between lifetime profit and type traits in Spanish Holstein cows. *J. Dairy Sci.* **2003**, *85*, 3480–3491. [[CrossRef](#)]
11. Török, E.; Komlosi, I.; Szőnyi, V.; Béla, B.; Mészáros, G.; Posta, J. Combinations of linear type traits affecting the longevity in Hungarian Holstein-Friesian cows. *Animals* **2021**, *11*, 3065. [[CrossRef](#)]
12. Simčič, M.; Luštrek, B.; Štepec, M.; Logar, B.; Potočnik, K. Estimation of genetic parameters of type traits in first parity cows of the Autochthonous Cika cattle in Slovenia. *Front. Genet.* **2021**, *12*, 724058. [[CrossRef](#)] [[PubMed](#)]
13. Vanraden, P.M.; Jensen, E.L.; Lawlor, T.; Funk, D.A. Prediction of transmitting abilities for Holstein type traits. *J. Dairy Sci.* **1990**, *73*, 191–197. [[CrossRef](#)]
14. Mazza, S.; Guzzo, N.; Sartori, C.; Berry, D.P.; Mantovani, R. Genetic parameters for linear type traits in the Rendena dual-purpose breed. *J. Anim. Breed. Genet.* **2013**, *131*, 27–35. [[CrossRef](#)] [[PubMed](#)]
15. Macciotta, N.; Cecchinato, A.; Mele, M.; Bittante, G. Use of multivariate factor analysis to define new indicator variables for milk composition and coagulation properties in Brown Swiss cows. *J. Dairy Sci.* **2012**, *95*, 7346–7354. [[CrossRef](#)] [[PubMed](#)]
16. Corrales Alvarez, J.; Cerón-muñoz, M.; Cañas-Álvarez, J.; Herrera, R.; Calvo Cardona, S. Relationship between type traits and milk production in Holstein cows from Antioquia, Colombia. *Rev. Mvz. Córdoba* **2011**, *16*, 2507–2513.
17. Vukasinovic, N.; Moll, J.; Künzi, N. Factor analysis for evaluating relationships between herd life and type traits in Swiss Brown cattle. *Lives. Prod. Sci.* **1997**, *49*, 227–234. [[CrossRef](#)]
18. Mazza, S.; Guzzo, N.; Sartori, C.; Mantovani, R. Factor analysis for genetic evaluation of linear type traits in dual-purpose autochthonous breeds. *Animal* **2015**, *10*, 372–380. [[CrossRef](#)]
19. Olasege, B.; Zhang, S.; Zhao, Q.; Liu, D.; Sun, H.; Wang, Q.; Ma, P.P.; Pan, Y. Genetic parameter estimates for body conformation traits using composite index, principal component, and factor analysis. *J. Dairy Sci.* **2019**, *102*, 5219–5229. [[CrossRef](#)]
20. Purcell, S.; Neale, B.; Todd-Brown, K.; Thomas, L.; Ferreira, M.; Bender, D.; Maller, J.; Sklar, P.; Bakker, P.; Daly, M.; et al. Plink: A tool set for whole-genome association and population-based linkage analyses. *Am. J. Hum. Genet.* **2007**, *81*, 559–575. [[CrossRef](#)]
21. Sargolzaei, M.; Iwaisaki, H.; Colleau, J.J. CFC: A tool for monitoring genetic diversity. In Proceedings of the 8th World Congress on Genetics Applied to Livestock Production, Belo Horizonte, Brazil, 13–18 August 2006; pp. 27–28.
22. Misztal, I.; Tsuruta, S.; Lourenco, D.; Masuda, Y.; Aguilar, I.; Legarra, A.; Vitezica, Z. *Manual for BLUPF90 Family of Programs*; University of Georgia: Athens, Greece, 2016.
23. Henderson, C. Best linear unbiased estimation and prediction under a selection model. *Biometrics* **1975**, *31*, 423–447. [[CrossRef](#)]
24. Lourenco, D.; Legarra, A.; Tsuruta, S.; Masuda, Y.; Aguilar, I.; Misztal, I. Single-step genomic evaluations from theory to practice: Using SNP chips and sequence data in BLUPF90. *Genes* **2020**, *11*, 790. [[CrossRef](#)]
25. Putz, A.; Tiezzi, F.; Maltecca, C.; Gray, K.; Kanuer, M. Variance component estimates for alternative litter size traits in swine. *J. Anim. Sci.* **2015**, *93*, 5153–5163. [[CrossRef](#)] [[PubMed](#)]

26. Russell, D. In search of underlying dimensions: The use (and abuse) of factor analysis in personality and social psychology bulletin. *Pers. Soc. Psychol. B* **2002**, *28*, 1629–1646. [[CrossRef](#)]
27. Hair, J.F.J.; Black, W.C.; Babin, B.J.; Anderson, R.E. Exploratory Factor Analysis. In *Multivariate Data Analysis: Pearson New International Edition*, 7th ed.; Prentice Hall: Upper Saddle River, NJ, USA, 2014; p. 118.
28. Strapáková, E.; Strapák, P.; Candrak, J.; Pavlík, I.; Dočkalová, K. Fleckscore system of exterior evaluation as a more accurate indirect predictor of longevity in Slovak Simmental dairy cows. *Czech. J. Anim. Sci.* **2021**, *66*, 487–494. [[CrossRef](#)]
29. Zavadilová, L.; Němcová, E.; Štípková, M.; Bouška, J. Relationships between longevity and conformation traits in Czech Fleckvieh cows. *Czech. J. Anim. Sci.* **2009**, *54*, 387–394. [[CrossRef](#)]
30. Mazza, S.; Sartori, C.; Mantovani, R. Genetic parameters of type traits in two strains of dual purpose autochthonous Valdostana cattle. *Livest. Sci.* **2015**, *178*, 35–42. [[CrossRef](#)]
31. Tarres, J.; Liu, Z.; Reinhardt, F.; Reents, R.; Ducrocq, V. Binational evaluation of type traits from Germany and France with a single-trait MACE animal model. *Animal* **2009**, *3*, 925–932. [[CrossRef](#)]
32. Novotný, L.; Frelich, J.; Beran, J.; Zavadilová, L. Genetic relationship between type traits, number of lactations initiated, and lifetime milk performance in Czech Fleckvieh cattle. *Czech. J. Anim. Sci.* **2017**, *62*, 2017–2501. [[CrossRef](#)]
33. Pfeiffer, C.; Fuerst-Waltl, B.; Ducrocq, V.; Fuerst, C. Approximate multivariate genetic evaluation of functional longevity and type traits in Austrian Fleckvieh cattle. In Proceedings of the 10th World Congress on Genetics Applied to Livestock Production, Vancouver, BC, Canada, 17–22 August 2014; p. 386.
34. Kern, E.; Cobuci, J.; Costa, C.; Mcmanus, C.; Braccini, N.J. Genetic association between longevity and linear type traits of Holstein cows. *Sci. Agric.* **2015**, *72*, 203–209. [[CrossRef](#)]
35. Špehar, M.; Štepec, M.; Potočnik, K. Variance components estimation for type traits in Slovenian brown Swiss cattle. *Acta Agric. Slov.* **2012**, *100*, 107–115.
36. Roveglia, C.; Niero, G.; Bobbo, T.; Penasa, M.; Finocchiaro, R.; Visentin, G.; Lopez-Villalobos, N.; Cassandro, M. Genetic parameters for linear type traits including locomotion in Italian Jersey cattle breed. *Livest. Sci.* **2019**, *229*, 131–136. [[CrossRef](#)]
37. Campos, R.; Cobuci, J.; Costa, C.; Braccini, N.J. Genetic parameters for type traits in Holstein cows in Brazil. *Rev. Bras. Zootecn.* **2012**, *41*, 2150–2161. [[CrossRef](#)]
38. Veerkamp, R.; Mulder, H.; Thompson, R.; Calus, M. Genomic and pedigree-based genetic parameters for scarcely recorded traits when some animals are genotyped. *J. Dairy Sci.* **2011**, *94*, 4189–4197. [[CrossRef](#)] [[PubMed](#)]
39. Wei, C.; Luo, H.; Zhao, B.; Tian, K.; Huang, X.; Wang, Y.; Fu, X.; Tian, Y.; Di, J.; Xu, X.; et al. The effect of integrating genomic information into genetic evaluations of Chinese Merino sheep. *Animals* **2020**, *10*, 569. [[CrossRef](#)] [[PubMed](#)]
40. Naserkheil, M.; Lee, D.; Mehrban, H. Improving the accuracy of genomic evaluation for linear body measurement traits using single-step genomic best linear unbiased prediction in Hanwoo beef cattle. *BMC Genet.* **2020**, *21*, 144. [[CrossRef](#)] [[PubMed](#)]
41. Song, H.; Zhang, J.; Zhang, Q.; Ding, X. Using different single-step strategies to improve the efficiency of genomic prediction on body measurement traits in Pig. *Front. Genet.* **2019**, *9*, 730. [[CrossRef](#)]
42. Larsson, R.; Rydén, J. Applications of discrete factor analysis. *Commun. Stat-Simul. C* **2021**, *50*, 1–11. [[CrossRef](#)]
43. Chu, M.; Shi, S. Phenotypic factor analysis for linear type traits in Beijing Holstein cows. *Asian Australas. J. Anim. Sci.* **2002**, *15*, 1527–1530. [[CrossRef](#)]
44. Sieber, M.; Freeman, A.E.; Hinz, P.N. Factor analysis for evaluating relationships between first lactation type scores and production data of Holstein dairy cows. *J. Dairy Sci.* **1987**, *70*, 1018–1026. [[CrossRef](#)]
45. Ali, A.K.; Koots, K.R.; Burnside, E.B. Factor analysis of genetic evaluations for type traits of Canadian Holstein sires and cows. *Asian Australas. J. Anim. Sci.* **1998**, *11*, 463–469. [[CrossRef](#)]
46. Brown, J.E.; Brown, C.J.; Butts, W.T. Evaluating relationships among immature measures of size, shape and performance of beef bulls. I. Principal Components as Measures of Size and Shape in Young Hereford and Angus Bulls. *J. Anim. Sci.* **1973**, *36*, 1010–1020. [[CrossRef](#)]
47. Kaiser, H. The varimax criterion for analytic rotation in factor analysis. *Psychometrika* **1958**, *23*, 187–200. [[CrossRef](#)]
48. Mantovani, R.I.C.; Contiero, B. Factor analysis for genetic evaluation of linear type traits in dual purpose breeds. *Ital. J. Anim. Sci.* **2010**, *4*, 31–33. [[CrossRef](#)]
49. Samoré, A.; Rizzi, R.; Rossoni, A.; Bagnato, A. Genetic parameters for functional longevity, type traits, SCS, milk flow and production in the Italian Brown Swiss. *Ital. J. Anim. Sci.* **2010**, *9*, 28. [[CrossRef](#)]
50. Budianto, D.; Widi, T.S.; Panjono, P.; Budisatria, I.G.S.I.; Hartatik, T. Estimation of body weight using linear body measurements in two crossbred beef cattle. In Proceedings of the 9th International Seminar on Tropical Animal Production (ISTAP 2021), Yogyakarta, Indonesia, 21–22 September 2021; pp. 332–337.
51. Manafiazar, G.; Goonewardene, L.; Miglior, F.; Crews, D.; Basarab, J.; Okine, E.; Wang, A. Genetic and phenotypic correlations among feed efficiency, production and selected conformation traits in dairy cows. *Animal* **2015**, *10*, 1–9. [[CrossRef](#)]
52. Dadati, E.; Kennedy, B.; Burnside, E. Relationships between conformation and reproduction in Holstein cows: Type and calving performance. *J. Dairy Sci.* **1985**, *68*, 2639–2645. [[CrossRef](#)]
53. Van, D.R.; Boettcher, P.; Schaeffer, L. Genetics of locomotion. *Livest. Prod. Sci.* **2004**, *90*, 247–253.
54. Pérez-Cabal, M.; Charfeddine, N. Short communication: Association of foot and leg conformation and body weight with claw disorders in Spanish Holstein cows. *J. Dairy Sci.* **2016**, *99*, 9104–9108. [[CrossRef](#)]

55. Rogers, G.W.; Banos, G.; Nielsen, U.; Philipsson, J. Genetic correlations among somatic cell scores, productive life, and type traits from the United States and udder health measures from Denmark and Sweden. *J. Dairy Sci.* **1998**, *81*, 1445–1453. [[CrossRef](#)]
56. Chrystal, M.; Seykora, A.; Hansen, L.B. Heritabilities of teat end shape and teat diameter and their relationships with somatic cell score. *J. Dairy Sci.* **1999**, *82*, 2017–2022. [[CrossRef](#)]
57. Němcová, E.; Štípková, M.; Zavadilová, L.; Bouška, J.; Vacek, M. The relationship between somatic cell count, milk production and six linearly scored type traits in Holstein cows. *Czech. J. Anim. Sci.* **2007**, *52*, 437–446.
58. DeGroot, B.; Keown, J.; Van Vleck, L.; Marotz, E. Genetic parameters and responses of linear type, yield traits, and somatic cell scores to divergent selection for predicted transmitting ability for type in Holsteins1. *J. Dairy Sci.* **2002**, *85*, 1578–1585. [[CrossRef](#)]
59. Dube, B.; Dzama, K.; Banga, C. Genetic analysis of somatic cell score and udder type traits in South African Holstein cows. *S. Afr. J. Anim. Sci.* **2008**, *38*, 1–11. [[CrossRef](#)]