# The Modes of the Poisson Distribution of Order 3 and 4 

Yeil Kwon 1,*(©) and Andreas N. Philippou 2 (D)<br>1 Department of Mathematics, University of Central Arkansas, Conway, AR 72035, USA<br>2 Department of Mathematics, University of Patras, 26500 Patras, Greece; anphilip@math.upatras.gr<br>* Correspondence: ykwon1@uca.edu

Citation: Kwon, Y.; Philippou, A.N. The Modes of the Poisson Distribution of Order 3 and 4. Entropy 2023, 25, 699. https:// doi.org/10.3390/e25040699

Academic Editor: Philip Broadbridge
Received: 21 March 2023
Revised: 13 April 2023
Accepted: 19 April 2023
Published: 21 April 2023


Copyright: © 2023 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (https:// creativecommons.org/licenses/by/ $4.0 /$ ).


#### Abstract

In this article, new properties of the Poisson distribution of order $k$ with parameter $\lambda$ are found. Based on them, the modes of the Poisson distributions of order $k=3$ and 4 are derived for $\lambda$ in $(0,1)$. They are $0,3,5$, and $0,4,7,8$, respectively, for $\lambda$ in specified subintervals of $(0,1)$. In addition, using Mathematica, computational results for the modes of the Poisson distributions of order $k=2,3$, and 4 are presented for $\lambda$ in specified subintervals of $(0,2)$.


Keywords: poisson distribution of order $k$; discrete distribution; mode; most probable value; order 4

## 1. Introduction

Following the papers of Philippou and Muwafi [1], Philippou et al. [2], Philippou [3-5], Aki et al. [6] and Aki [7], there has been an upsurge in the study of distributions of order $k$ (distributions of runs) due to their theoretical importance and great applicability in reliability, start-up demonstration tests, sampling inspection, etc. See, e.g., Ling [8], Mohanty [9], Chang [10], Johnson et al. [11], Shmueli and Kohen [12], Balakrishnan and Koutras [13], Fu and Lou [14], Eryilmaz [15], Rakitzis and Antzoulakos [16], Dafnis et al. [17], Sengar et al. [18], Kwon [19], and references therein. However, the modes of these distributions are not yet known, except for the modes of the geometric distribution of order $k$ and partial results for the mode(s) of the Poisson distribution of order $k$ and the negative binomial distribution of the same order derived by Luo [20], Georghiou et al. [21], Philippou [22], Shao and Fu [23], and Georghiou et al. [24].

The Poisson distribution of order $k\left(k \geq 1\right.$, integer) with parameter $\lambda(>0)$ say $P_{k}(\lambda)$, has probability mass function (pmf)

$$
\begin{equation*}
f_{k}(x ; \lambda)=e^{-k \lambda} \sum \frac{\lambda^{x_{1}+x_{2}+\cdots+x_{k}}}{x_{1}!x_{2}!\cdots x_{k}!}, \quad x=0,1,2, \ldots \tag{1}
\end{equation*}
$$

where the summation is taken over all $k$-tuples of non-negative integers $x_{1}, x_{2}, \ldots, x_{k}$ such that $x_{1}+2 x_{2}+\cdots+k x_{k}=x$.

It was derived by Philippou et al. [2] as a limit of the negative binomial distribution of order $k$, and it was named so, since, for $k=1$, it reduces to the Poisson distribution with parameter $\lambda$. It is a special case, for $\lambda_{1}=\lambda_{2}=\cdots=\lambda_{k}=\lambda$, of the multiparameter Poisson distribution of order $k$ (Philippou [25]), also known as $k$ stuttering Poisson distribution (Galliher et al. [26], Patel [27]). The latter author discussed the estimation of the parameters of the triple and quadruple stuttering distributions and noted that the cases $k=2,3$, and 4 are more frequently observed in practice.

Let $m_{k, \lambda}$ denote the mode(s) of $f_{k}(x ; \lambda)$, i.e., the value(s) of $x$ for which $f_{k}(x ; \lambda)$ attains its maximum. It is well known that $m_{1, \lambda}=\lambda$ or $\lambda-1$ if $\lambda \in \mathbb{N}$ and $m_{1, \lambda}=\lfloor\lambda\rfloor$, if $\lambda$ does not belong to $\mathbb{N}$, where $\lfloor\alpha\rfloor$ denotes the greatest integer part of $\alpha$. Philippou [3] derived some properties of $f_{k}(x ; \lambda)$ and posed the problem of finding its mode(s) for $k \geq 2$.

Hirano et al. [28] presented several graphs of $f_{k}(x ; \lambda)$ for $\lambda \in(0,1)$ and $2 \leq k \leq 8$, and Luo [20] derived the following lower bound inequality for $m_{k, \lambda}$,

$$
m_{k, \lambda} \geq k \lambda \sqrt[k]{k!}-\frac{1}{2} k(k+1), \quad k \geq 1, \quad \lambda>0
$$

Georghiou et al. [21] employed the probability generating function of the Poisson distribution of order $k$ to improve the lower bound of Luo [20] and also to give an upper bound for $m_{k, \lambda}$

$$
\begin{equation*}
\frac{1}{2} k(k+1)(\lambda-1)+1-\delta_{k, 1} \leq m_{k, \lambda} \leq\left\lfloor\frac{1}{2} k(k+1) \lambda\right\rfloor=u_{k, \lambda}, \quad k \geq 1, \quad \lambda>0 \tag{2}
\end{equation*}
$$

where $\delta_{k, 1}$ denotes the Kronecker delta. With the bounds of $m_{k, \lambda}$ in (2), they showed that

$$
\begin{equation*}
m_{k, \lambda}=\frac{1}{2} k(k+1) \lambda-\left\lfloor\frac{k}{2}\right\rfloor, \quad 2 \leq k \leq 5, \quad \lambda \in \mathbb{N} . \tag{3}
\end{equation*}
$$

Using the upper bound $u_{k, \lambda}$ of (2) and the definition of $m_{k, \lambda}$, Philippou [22] found that:
(a) For any integer $k \geq 1$ and $0<\lambda<2 / k(k+1)$, the Poisson distribution of order $k$ has a unique mode $m_{k, \lambda}=0$.
(b) The Poisson distribution of order 2 has a unique mode $m_{k, \lambda}=0$ if $0<\lambda<\sqrt{3}-1$; it has two modes $m_{k, \lambda}=0$ and 2 if $0<\lambda=\sqrt{3}-1$, and it has a unique mode $m_{k, \lambda}=2$ if $\sqrt{3}-1<\lambda<1$. (The number $\sqrt{3}-1$ is the positive root (say $r_{2}$ ) of the quadratic equation $\lambda^{2}+2 \lambda-2=0$.)

Remark 1. Since the modes of the Poisson distribution of order $k$ with parameter $\lambda$ are defined as the values of $x \in\{0,1,2, \ldots\}$, which maximize $f_{k}(x ; \lambda)$, they are its most probable values and they may be obtained numerically for any given positive integer $k$ and positive $\lambda$, from

$$
f_{k}\left(m_{k, \lambda} ; \lambda\right)=\max \left\{f_{k}(x ; \lambda) \mid x \in\left\{0,1,2, \ldots, u_{k}(\lambda)\right\}\right\} .
$$

In the present short note, we derive some additional properties of $f_{k}(x ; \lambda)$ and find the modes of the Poisson distribution of order $k=3$ and $k=4$ for $0<\lambda<1$. Furthermore, Section 3 presents computational results for the modes of the Poisson distributions of order $k=2,3$, and 4 for $\lambda \in(0,2)$. Finally, in Section 4, we briefly discuss our results, give the moment estimator of $\lambda(>0)$ for $k \geq 1$, and indicate further research.

## 2. Main and Preliminary Results

The mode(s) of a discrete probability mass function is (are) its most probable value(s). In this section, we derive the modes of the Poisson distributions of order 3 and 4 , respectively, when $0<\lambda<1$ (see Propositions 1 and 2). In order to do so, we first state and prove three lemmas, regarding $h_{k}(x ; \lambda)=e^{k \lambda} f_{k}(x ; \lambda)$, which we use, along with relation (2), to prove the propositions.

Because of (1),

$$
\begin{equation*}
h_{k}(x ; \lambda)=\sum \frac{\lambda^{x_{1}+x_{2}+\cdots+x_{k}}}{x_{1}!x_{2}!\cdots x_{k}!}, \quad x=0,1,2, \ldots ; \lambda>0 \tag{4}
\end{equation*}
$$

where the summation is taken over all $k$-tuples of non-negative integers $x_{1}, x_{2}, \ldots, x_{k}$ such that $x_{1}+2 x_{2}+\cdots+k x_{k}=x$. Note that $h_{k}(0 ; \lambda)=1$ and $h_{k_{1}}(x ; \lambda)=h_{k_{2}}(x ; \lambda)$, for $1 \leq x \leq k_{1} \leq k_{2}$.

Georghiou et al. [21] provided a recursive form of $f_{k}(x ; \lambda)$ as

$$
\begin{equation*}
x f_{k}(x ; \lambda)=\sum_{j=1}^{k} j \lambda f_{k}(x-j ; \lambda), \quad x \geq 1 . \tag{5}
\end{equation*}
$$

It can be restated, in terms of $h_{k}(x ; \lambda)$, as

$$
x h_{k}(x ; \lambda)= \begin{cases}\sum_{j=1}^{x} j \lambda h_{k}(x-j ; \lambda) & 1 \leq x \leq k  \tag{6}\\ \sum_{j=1}^{k} j \lambda h_{k}(x-j ; \lambda) & x>k,\end{cases}
$$

with $h_{k}(0 ; \lambda)=1$.
Lemma 1. For $2 \leq x \leq k$ and a fixed $\lambda>0$,

$$
\lambda \leq h_{k}(x-1 ; \lambda)<h_{k}(x ; \lambda) .
$$

Proof. To avoid the abuse of notation, let $h(x) \equiv h_{k}(x ; \lambda)$. From (4), it is easy to see $h(1)=\lambda$. Using (6), for $2 \leq x \leq k$,

$$
\begin{aligned}
x h(x)-(x-1) h(x-1) & =\lambda \sum_{j=1}^{x} j h(x-j)-\lambda \sum_{j=1}^{x-1} j h(x-1-j) \\
& =\lambda \sum_{j=1}^{x} j h(x-j)-\lambda \sum_{j=2}^{x}(j-1) h(x-j) \\
& =\lambda \sum_{j=1}^{x} j h(x-j)-\lambda \sum_{j=2}^{x} j h(x-j)+\lambda \sum_{j=2}^{x} h(x-j) \\
& =\lambda h(x-1)+\lambda \sum_{j=2}^{x} j h(x-j)-\lambda \sum_{j=2}^{x} j h(x-j)+\lambda \sum_{j=2}^{x} h(x-j) \\
& =\lambda h(x-1)+\lambda \sum_{j=2}^{x} h(x-j) \\
& =\lambda \sum_{j=1}^{x} h(x-j)
\end{aligned}
$$

From (6), since $(x-1) h(x-1)=\sum_{j=1}^{x-1} j \lambda h(x-1-j)$, we have

$$
\begin{aligned}
x[h(x)-h(x-1)] & =x h(x)-(x-1) h(x-1)-h(x-1) \\
& =\lambda\left[\sum_{j=1}^{x} h(x-j)-\frac{1}{x-1} \sum_{j=1}^{x-1} j h(x-1-j)\right] \\
& =\lambda\left[\sum_{j=1}^{x} h(x-j)-\frac{1}{x-1} \sum_{j=2}^{x}(j-1) h(x-j)\right] \\
& =\lambda\left[h(x-1)+\sum_{j=2}^{x} h(x-j)-\frac{1}{x-1} \sum_{j=2}^{x}(j-1) h(x-j)\right] \\
& =\lambda\left[h(x-1)+\sum_{j=2}^{x}\left(1-\frac{j-1}{x-1}\right) h(x-j)\right]>0 .
\end{aligned}
$$

Therefore, $h_{k}(x ; \lambda)>h_{k}(x-1 ; \lambda)$ for $2 \leq x \leq k$ with a fixed value of $\lambda>0$.

Lemma 2. For $k \geq 2$ and $0<\lambda<1$, the equation $h_{k}(k ; \lambda)=h_{k}(0 ; \lambda)$ has exactly one root $\lambda=r_{k}\left(0<r_{k}<1\right)$ such that

$$
\begin{cases}h_{k}(0 ; \lambda)>h_{k}(k ; \lambda), & \text { if } 0<\lambda<r_{k} \\ h_{k}(0 ; \lambda)<h_{k}(k ; \lambda), & \text { if } r_{k}<\lambda<1 .\end{cases}
$$

Proof. First, note that $h_{k}(0 ; 1)=1$ and $\lim _{\lambda \rightarrow 0^{+}} h_{k}(0 ; \lambda)=1$ because the relation (4) implies $h_{k}(0 ; \lambda)=1$ for $\lambda>0$. Since, for $k \geq 1, h_{k}(k ; \lambda)$ is a polynomial function of $\lambda$ with positive coefficient only and without constant term, we have $\lim _{\lambda \rightarrow 0^{+}} h_{k}(k ; \lambda)=0$. Note that $h_{k}(1 ; \lambda)=\lambda$ and by Lemma $1, h_{k}(1 ; \lambda)<h_{k}(2 ; \lambda)<\cdots<h_{k}(k ; \lambda)$ for $\lambda>0$. Thus, $h_{k}(k ; 1)>h_{k}(1 ; 1)=1$ for $k \geq 2$.

Second, let $g_{k}(\lambda)=h_{k}(k ; \lambda)-h_{k}(0 ; \lambda)$. Then, since $\lim _{\lambda \rightarrow 0^{+}} g_{k}(\lambda)=\lim _{\lambda \rightarrow 0^{+}} h_{k}(k ; \lambda)-$ $\lim _{\lambda \rightarrow 0^{+}} h_{k}(0 ; \lambda)=0-1=-1<0$, and $g_{k}(1)=h_{k}(k ; 1)-h_{k}(0 ; 1)>h_{k}(1 ; 1)-h_{k}(0 ; 1)=$ $1-1=0$.

Lastly, Since $h_{k}(x ; \lambda)$ is a polynomial function with positive coefficients only, we have $\frac{\partial}{\partial \lambda} g_{k}(\lambda)>0$ for $\lambda>0$, which implies $g_{k}(\lambda)$ is an increasing function of $\lambda$ for $0<\lambda<1$.

Therefore, $g_{k}(\lambda)$ has a unique root, say $r_{k}$, between 0 and 1 with $h_{k}(0 ; \lambda)>h_{k}(k ; \lambda)$, if $0<\lambda<r_{k}$ and $h_{k}(0 ; \lambda)<h_{k}(k ; \lambda)$, if $r_{k}<\lambda<1$.

Lemma 3. For $k \geq 2$ and $0<\lambda \leq r_{k}<1$,

$$
h_{k}(k ; \lambda)>h_{k}(k+1 ; \lambda)
$$

where $r_{k}$ is defined in Lemma 2.
Proof. For notational simplicity, let $h(x) \equiv h_{k}(x ; \lambda)$. From (6), we have

$$
k h(k)=\sum_{j=1}^{k} j \lambda h_{k}(k-j) \quad \text { and } \quad(k+1) h_{k}(k+1)=\sum_{j=1}^{k} j \lambda h_{k}(k+1-j) .
$$

Thus, with $h(0)=1$,

$$
\begin{align*}
k[h(k)-h(k+1)] & =\sum_{j=1}^{k} j \lambda h(k-j)-\sum_{j=1}^{k} j \lambda h(k+1-j)+h(k+1) \\
& =\sum_{j=1}^{k} j \lambda h(k-j)-\sum_{j=0}^{k-1}(j+1) \lambda h(k-j)+h(k+1) \\
& =\sum_{j=1}^{k} j \lambda h(k-j)-\sum_{j=0}^{k-1} j \lambda h(k-j)-\sum_{j=0}^{k-1} \lambda h(k-j)+h(k+1) \\
& =k \lambda h(0)-\sum_{j=0}^{k-1} \lambda h(k-j)+h(k+1) \\
& =\lambda\left(k-\sum_{j=0}^{k-1} h(k-j)\right)+h(k+1) . \tag{7}
\end{align*}
$$

Since each $h_{k}(1 ; \lambda), h_{k}(2 ; \lambda), \ldots, h_{k}(k ; \lambda)$ are increasing functions of lambda, they all have the maximum value at $\lambda=r_{k}$ on $\lambda \in\left(0, r_{k}\right]$. Moreover, by Lemma 1 and the definition of
$r_{k}$, we have $h_{k}\left(1 ; r_{k}\right)<h_{k}\left(2 ; r_{k}\right)<\cdots<h_{k}\left(k ; r_{k}\right)=1$. Therefore, for $0<\lambda \leq r_{k}$, the first term of (7) can be written as

$$
\begin{aligned}
k-\sum_{j=0}^{k-1} h_{k}(k-j ; \lambda) & \geq k-\sum_{j=0}^{k-1} h_{k}\left(k-j ; r_{k}\right) \\
& =k-\left[h_{k}\left(k ; r_{k}\right)+h_{k}\left(k-1 ; r_{k}\right)+\cdots+h_{k}\left(1 ; r_{k}\right)\right] \\
& >k-k\left[\left(h_{k}\left(k ; r_{k}\right)\right]=0\right.
\end{aligned}
$$

and it implies $h_{k}(k ; \lambda)>h_{k}(k+1 ; \lambda)$.
Proposition 1. Let $m_{3, \lambda}$ denote the mode(s) of the Poisson distribution of order 3 with parameter $\lambda(>0)$. Let $r_{3}(=0.6016791318 \ldots)$ and $r_{3,5}(=0.9962030611 \ldots)$ be the positive roots of the equations $\lambda^{3}+6 \lambda^{2}+6 \lambda-6=0$, and $\lambda^{4}+20 \lambda^{3}+100 \lambda^{2}-120=0$, respectively. Then

$$
m_{3, \lambda}= \begin{cases}0, & \text { if } 0<\lambda<r_{3} \\ 0 \text { and } 3, & \text { if } \lambda=r_{3} \\ 3, & \text { if } r_{3}<\lambda<r_{3,5} \\ 3 \text { and } 5, & \text { if } \lambda=r_{3,5} \\ 5, & \text { if } r_{3,5}<\lambda<1\end{cases}
$$

Proof. The Equation (2) implies

$$
0 \leq m_{3, \lambda} \leq\left\lfloor\frac{3}{2}(3+1) \lambda\right\rfloor=\lfloor 6 \lambda\rfloor \leq 5, \text { for } 0<\lambda<1
$$

Hence, it is enough we compare the magnitudes of $f_{3}(x ; \lambda)$, or the magnitudes of $h_{3}(x ; \lambda)$ for $0 \leq x \leq 5$. By Lemma 1, we have $h_{3}(1 ; \lambda)<h_{3}(2 ; \lambda)<h_{3}(3 ; \lambda)$ for $0<\lambda<1$, and, by Lemma 2 and 3, it is given $1=h_{3}(0 ; \lambda)>h_{3}(3 ; \lambda)>h_{3}(4 ; \lambda)$ for $0<\lambda<r_{3}$. In addition, the maximum of $h_{3}(5 ; \lambda)$ on $\lambda \in\left(0, r_{3}\right]$ is $h_{3}\left(5 ; r_{k}\right)=r_{3}^{2}+r_{3}^{3}+\frac{1}{6} r_{3}^{4}+\frac{1}{120} r_{3}^{5}<1$. Hence, $m_{3, \lambda}=0$ for $0<\lambda<r_{3}$.

Furthermore, $h_{3}(3 ; \lambda)>h_{3}(4 ; \lambda)$ for $0<\lambda<1$. To see this, let $d_{k}(x ; \lambda)=h_{k}(k ; \lambda)-$ $h_{k}(x ; \lambda)$. Then

$$
\begin{aligned}
d_{3}(4 ; \lambda) & =h_{3}(3 ; \lambda)-h_{3}(4 ; \lambda) \\
& =\left(\lambda+\lambda^{2}+\frac{1}{6} \lambda^{3}\right)-\left(\frac{3}{2} \lambda^{2}+\frac{1}{2} \lambda^{3}+\frac{1}{24} \lambda^{4}\right) \\
& =\lambda-\frac{1}{2} \lambda^{2}-\frac{1}{3} \lambda^{3}-\frac{1}{24} \lambda^{4},
\end{aligned}
$$

with $\frac{\partial^{2}}{\partial \lambda^{2}} d_{3}(4 ; \lambda)=-1-2 \lambda-\lambda^{2} / 2<0$, which implies that $d_{3}(4 ; \lambda)$ is concave down for $\lambda>0$. Since $\lim _{\lambda \rightarrow 0^{+}} d_{3}(4 ; \lambda)=0^{+}$, and $d_{3}(4 ; 1)=1 / 8>0$, we have $h_{3}(3 ; \lambda)>h_{3}(4 ; \lambda)$ for $0<\lambda<1$. Hence, it suffices that we compare the magnitude of $h_{3}(0 ; \lambda), h_{3}(3 ; \lambda)$ and $h_{3}(5 ; \lambda)$ to find the modes for $r_{3}<\lambda<1$.

$$
\begin{aligned}
d_{3}(5 ; \lambda) & =h_{3}(3 ; \lambda)-h_{3}(5 ; \lambda) \\
& =\left(\lambda+\lambda^{2}+\frac{1}{6} \lambda^{3}\right)-\left(\lambda^{2}+\lambda^{3}+\frac{1}{6} \lambda^{4}+\frac{1}{120} \lambda^{5}\right) \\
& =\lambda-\frac{5}{6} \lambda^{3}-\frac{1}{6} \lambda^{4}-\frac{1}{120} \lambda^{5} .
\end{aligned}
$$

For $\lambda>0, d_{3}(5 ; \lambda)$ is concave down because $\frac{\partial^{2}}{\partial \lambda^{2}} d_{3}(5 ; \lambda)=-5 \lambda-2 \lambda^{2}-\lambda^{3} / 6<0$. Since $\lim _{\lambda \rightarrow 0^{+}} d_{3}(5 ; \lambda)=0^{+}$and $d_{3}(5 ; 1)=-1 / 120$, the equation $d_{3}(5 ; \lambda)=0$, equivalently
$\lambda^{4}+20 \lambda^{3}+100 \lambda^{2}-120=0$, has one positive root, say $r_{3,5}$. Thus, $h_{3}(3 ; \lambda)>h_{3}(5 ; \lambda)$ for $0<\lambda<r_{3,5}$ and $h_{3}(3 ; \lambda)<h_{3}(5 ; \lambda)$ for $r_{3,5}<\lambda<1$. Since $0<r_{3}=0.601679 \ldots<r_{3,5}=$ 0.996203... $<1$, we have

$$
\begin{cases}f_{3}(0 ; \lambda)>f_{3}(3 ; \lambda)>f_{3}(5 ; \lambda), & \text { if } 0<\lambda<r_{3} \\ f_{3}(3 ; \lambda)>f_{3}(0 ; \lambda) \& f_{3}(3 ; \lambda)>f_{3}(5 ; \lambda), & \text { if } r_{3}<\lambda<r_{3,5} \\ f_{3}(5 ; \lambda)>f_{3}(3 ; \lambda)>f_{3}(0 ; \lambda), & \text { if } r_{3,5}<\lambda<1\end{cases}
$$

Proposition 2. Let $m_{4, \lambda}$ denote the mode(s) of the Poisson distribution of order 4 with parameter $\lambda(>0)$, and let $r_{4}(=0.5203510176 \ldots), r_{4,7}(=0.7947408725 \ldots)$, and $r_{7,8}(=0.8944652714 \ldots)$, respectively, be the positive roots of the equations $\lambda^{4}+12 \lambda^{3}+36 \lambda^{2}+24 \lambda-24=0, \lambda^{6}+42 \lambda^{5}+$ $630 \lambda^{4}+3990 \lambda^{3}+7560 \lambda^{2}-2520 \lambda-5040=0$, and $\lambda^{6}+48 \lambda^{5}+840 \lambda^{4}+6720 \lambda^{3}+18,480 \lambda^{2}-$ $20,160=0$. Then

$$
m_{4, \lambda}= \begin{cases}0, & \text { if } 0<\lambda<r_{4}, \\ 0 \text { and } 4, & \text { if } \lambda=r_{4}, \\ 4, & \text { if } r_{4}<\lambda<r_{4,7}, \\ 4 \text { and } 7, & \text { if } \lambda=r_{4,7}, \\ 7, & \text { if } r_{4,7}<\lambda<r_{7,8}, \\ 7 \text { and } 8, & \text { if } \lambda=r_{7,8} \\ 8, & \text { if } r_{7,8}<\lambda<1 .\end{cases}
$$

Proof. The Equation (2) implies

$$
0 \leq m_{4, \lambda} \leq\left\lfloor\frac{4}{2}(4+1) \lambda\right\rfloor=\lfloor 10 \lambda\rfloor \leq 9, \text { for } 0<\lambda<1
$$

Hence, it is enough we compare the magnitudes of $f_{4}(x ; \lambda)$, or the magnitudes of $h_{4}(x ; \lambda)$ for $0 \leq x \leq 9$. By Lemma 1, we have $h_{4}(1 ; \lambda)<h_{4}(2 ; \lambda)<h_{4}(3 ; \lambda)<h_{4}(4 ; \lambda)$. Thus, we will compare the magnitudes of $h_{4}(x ; \lambda)$ for $4 \leq x \leq 9$, and $h_{4}(0 ; \lambda)$. They are given by

$$
\begin{aligned}
& h_{4}(4 ; \lambda)=\lambda+\frac{3}{2} \lambda^{2}+\frac{1}{2} \lambda^{3}+\frac{1}{24} \lambda^{4} \\
& h_{4}(5 ; \lambda)=2 \lambda^{2}+\lambda^{3}+\frac{1}{6} \lambda^{4}+\frac{1}{120} \lambda^{5} \\
& h_{4}(6 ; \lambda)=\frac{3}{2} \lambda^{2}+\frac{5}{3} \lambda^{3}+\frac{5}{12} \lambda^{4}+\frac{1}{24} \lambda^{5}+\frac{1}{720} \lambda^{6} \\
& h_{4}(7 ; \lambda)=\lambda^{2}+2 \lambda^{3}+\frac{5}{6} \lambda^{4}+\frac{1}{8} \lambda^{5}+\frac{1}{120} \lambda^{6}+\frac{1}{5040} \lambda^{7} \\
& h_{4}(8 ; \lambda)=\frac{1}{2} \lambda^{2}+2 \lambda^{3}+\frac{31}{24} \lambda^{4}+\frac{7}{24} \lambda^{5}+\frac{7}{240} \lambda^{6}+\frac{1}{720} \lambda^{7}+\frac{1}{40,320} \lambda^{8} \\
& h_{4}(9 ; \lambda)=\frac{5}{3} \lambda^{3}+\frac{5}{3} \lambda^{4}+\frac{13}{24} \lambda^{5}+\frac{7}{90} \lambda^{6}+\frac{1}{180} \lambda^{7}+\frac{1}{5040} \lambda^{8}+\frac{1}{362,880} \lambda^{9}
\end{aligned}
$$

Note that $h_{4}(x ; \lambda)$, for $4 \leq x \leq 9$, are strictly increasing functions of $\lambda$ and $h_{4}(x ; 0)=1$. Table 1 displays the function values of $h_{4}(x ; \lambda)$ for $4 \leq x \leq 9$ with $\lambda=0.5,0.6,0.7,0.8,0.9$, 1.0, and 1.1. From the function values of the Table 1 , we can see $h_{4}(4 ; \lambda)=1$ has a root $r_{4} \in(0.5,0.6), h_{4}(4 ; \lambda)=h_{4}(7 ; \lambda)$ has a root $r_{4,7} \in(0.7,0.8)$, and $h_{4}(7 ; \lambda)=h_{4}(8 ; \lambda)$ has a $\operatorname{root} r_{7,8} \in(0.8,0.9)$.

Table 1. The function values of $h_{4}(x ; \lambda)$ for $x=0$ and $4 \leq x \leq 10$ with $\lambda=0.5,0.6,0.7,0.8,0.9$, 1.0, and 1.1. The bolded value in each column stands for the maximum of $h_{4}(x ; \lambda)$, which implies $m_{4, \lambda}=x$ with a value of $\lambda$ given in the corresponding column. Note that $0.5<r_{4}<0.6<0.7<$ $r_{4,7}<0.8<r_{7,8}<0.9$. The function values are calculated based on substantially tight grid for $\lambda$. The table displays only the meaningful $\lambda$ values.

|  |  |  | $\boldsymbol{\lambda}$ |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\boldsymbol{x}$ | $\mathbf{0 . 5}$ | $\mathbf{0 . 6}$ | $\mathbf{0 . 7}$ | $\mathbf{0 . 8}$ | $\mathbf{0 . 9}$ | $\mathbf{1 . 0}$ | $\mathbf{1 . 1}$ |
| 0 | $\mathbf{1 . 0 0 0 0}$ | 1.0000 | 1.0000 | 1.0000 | 1.0000 | 1.0000 | 1.0000 |
| 4 | 0.9401 | $\mathbf{1 . 2 5 3 4}$ | $\mathbf{1 . 6 1 6 5}$ | 2.0331 | 2.5068 | 3.0417 | 3.6415 |
| 5 | 0.6357 | 0.9582 | 1.3644 | 1.8630 | 2.4633 | 3.1750 | 4.0084 |
| 6 | 0.6107 | 0.9573 | 1.4139 | 1.9980 | 2.7287 | 3.6264 | 4.7129 |
| 7 | 0.5561 | 0.9101 | 1.3981 | $\mathbf{2 . 0 4 8 5}$ | 2.8931 | 3.9669 | 5.3085 |
| 8 | 0.4653 | 0.8035 | 1.2937 | 1.9766 | 2.8989 | $\mathbf{4 . 1 1 3 9}$ | 5.6823 |
| 9 | 0.3307 | 0.6219 | 1.0725 | 1.7351 | 2.6724 | 3.9585 | 5.6799 |
| 10 | 0.2686 | 0.5273 | 0.9457 | 1.5861 | 2.5258 | 3.8593 | $\mathbf{5 . 7 0 0 4}$ |

Therefore,

$$
\begin{cases}f_{4}(0 ; \lambda)>f_{4}(x ; \lambda) \text { where } 1 \leq x \leq 9, & \text { if } 0<\lambda<r_{4} \\ f_{4}(4 ; \lambda)>f_{4}(x ; \lambda) \text { where } 0 \leq x \leq 3, \text { or } 5 \leq x \leq 9 & \text { if } r_{4}<\lambda<r_{4,7} \\ f_{4}(7 ; \lambda)>f_{4}(x ; \lambda) \text { where } 0 \leq x \leq 6, \text { or } 8 \leq x \leq 9 & \text { if } r_{4,7}<\lambda<r_{7,8} \\ f_{4}(8 ; \lambda)>f_{4}(x ; \lambda) \text { where } 0 \leq x \leq 7, \text { or } x=9 & \text { if } r_{7,8}<\lambda<1\end{cases}
$$

## 3. More Computational Resutls

This section provides more computational results using the computer algebra system Mathematica. Table 2 shows the modes of Poisson distribution of order $k=2,3$, and 4 for $0<\lambda<2$. For $\lambda>1$, we observe that the mode values frequently change as $\lambda$ increases. However, for every value of $k$, the first two modes are 0 and $k$ for some subintervals of $\lambda$ between zero and one. We also note that for $k=2,3$, and $4, m_{k, 1}=2,5$, and 8 , (the modes of the Poisson distribution of order $k$ with $\lambda=1$ ) as it should, in accordance with (3).

Table 2. The modes of Poisson distribution of order $k=2,3$, and 4 for $0<\lambda<2$. The lower and upper bounds of $\lambda$ are the approximated values.

| $\begin{aligned} & k=2 \\ & \text { Mode } \end{aligned}$ | Interval of $\boldsymbol{\lambda}$ | $\begin{aligned} & k=3 \\ & \text { Mode } \end{aligned}$ | Interval of $\boldsymbol{\lambda}$ | $\begin{aligned} & k=4 \\ & \text { Mode } \end{aligned}$ | Interval of $\boldsymbol{\lambda}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | (0.0000, 0.7321) | 0 | (0.0000, 0.6017) | 0 | (0.0000, 0.5204) |
| 2 | (0.7321, 1.3412) | 3 | (0.6017, 0.9962) | 4 | (0.5204, 0.7947) |
| 4 | (1.3412, 1.8851) | 5 | (0.9962, 1.0612) | 7 | (0.7947, 0.8945) |
| 5 | (1.8851, 2.0000 ) | 6 | (1.0612, 1.3881) | 8 | (0.8945, 1.0950) |
|  |  | 7 | (1.3881, 1.4293) | 10 | (1.0950, 1.2056) |
|  |  | 8 | (1.4293, 1.6286) | 11 | (1.2056, 1.3244) |
|  |  | 9 | (1.6286, 1.8197) | 12 | (1.3244, 1.4332) |
|  |  | 10 | (1.8197, 1.9590) | 13 | (1.4332, 1.5124) |
|  |  | 11 | (1.9590, 2.0000) | 14 | (1.5124, 1.6183) |
|  |  |  |  | 15 | (1.6183, 1.7215) |
|  |  |  |  | 16 | (1.7215, 1.8210) |
|  |  |  |  | 17 | (1.8210, 1.9180) |
|  |  |  |  | 18 | (1.9180, 2.0000) |

## 4. Moment Estimation of the Parameter $\lambda$ of $P_{k}(\lambda)$, Discussion, and Further Research

Despite the upsurge of the study of distributions of order $k$ or runs since the early 1980s, their modes, due to the difficulty of obtaining them, are not known, except for
the mode of the geometric distribution of order $k$ and partial results for the modes of the negative binomial and Poisson distributions of order $k$. Their probability generating functions, however, and moments are well known.

The mean and variance of $P_{k}(\lambda)$, for example, are (a) $k(k+1) \lambda / 2$ and (b) $k(k+1)$ $(2 k+1) \lambda / 6$ (see, e.g., Philippou [3,25]). By means of them and the method of moments estimation, we now give the moment estimator $\hat{\lambda}$ of $\lambda$ of $P_{k}(\lambda)$. Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample of size $n$ from $P_{k}(\lambda)$, and set $\bar{X}=\left(X_{1}+X_{2}+\cdots+X_{n}\right) / n$. Then, the moment estimator of $\lambda$ is $\hat{\lambda}=2 \bar{X} /[k(k+1)]$. It is unbiased, and has variance $\operatorname{Var}(\hat{\lambda})=2(2 k+1) \lambda /[3 k(k+1) n]$. In fact, by the method of moments and (a), $\bar{X}=k(k+1) \hat{\lambda} / 2$, which implies $\hat{\lambda}=2 \bar{X} /[k(k+1)]$. It follows by (a) and (b), respectively, that $\hat{\lambda}$ is unbiased for $\lambda$, since $E(\hat{\lambda})=2 E(\bar{X}) /[k(k+1)]=\lambda$, and has variance $\operatorname{Var}(\hat{\lambda})=4 \operatorname{Var}(\bar{X}) /\left[k^{2}(k+1)^{2}\right]=\left[4 / k^{2}(k+1)^{2}\right] \cdot[k(k+1)(2 k+1) \lambda /(6 n)]=2(2 k+$ 1) $\lambda /[3 k(k+1) n]$, which was to be shown.

In the present article, in addition to the above paragraph regarding $P_{k}(\lambda)$, we derived a few new properties of the Poisson distribution of order $k$, and using them, along with a result of Georghiou et al. [21], we found the modes or most probable values of the Poisson distributions of order 3 and 4 for $\lambda$ in the interval $(0,1)$. In addition, using Mathematica and a personal computer, we found the modes of the Poisson distributions of order 2, 3, and 4 for $\lambda \in(0,2)$. We observe that for $k=2,3$, and 4 , the first two modes are 0 for $0<\lambda<r_{k}$, and $k$ for $r_{k}<\lambda<r_{k}+l_{k}$, where $l_{k}$ stands for the length of the interval on which $k$ is the mode of the Poisson distribution of order $k$. Further research may include several interesting problems: Is it generally true that $m_{k, \lambda}=0$ for $k \geq 2$ and $0<\lambda<r_{k}$, and $m_{k, \lambda}=k$ for $k \geq 2$ and $r_{k}<\lambda<r_{k}+l_{k}$ ? Does $r_{k}$ decrease as $k$ increases? If it does, how fast is $r_{k}$ decreasing? What positive integers cannot be modes of the Poisson distribution of order $k$ ?

Author Contributions: Conceptualization, Y.K. and A.N.P.; methodology, Y.K. and A.N.P.; software, Y.K.; validation, Y.K. and A.N.P.; formal analysis, Y.K. and A.N.P.; resources, A.N.P.; writing-original draft, Y.K. and A.N.P.; writing-review and editing, Y.K. and A.N.P. All authors have read and agreed to the published version of the manuscript.

Funding: This research received no external funding.
Institutional Review Board Statement: Not applicable.
Data Availability Statement: Data sharing not applicable.
Acknowledgments: The authors would like to thank the Editorial Board and the referees for their constructive suggestions and comments, which helped us improve the presentation of the manuscript. The authors also thank Nikolaos Ioakimidis for his dedicated support of the computational results.

Conflicts of Interest: The authors declare no conflict of interest.

## References

1. Philippou, A.N.; Muwafi, A.A. Waiting for the $k$-th consecutive success and the Fibonacci sequence of order $k$. Fibonacci Q. 1982, 20, 28-32.
2. Philippou, A.N.; Georghiou, C.; Philippou, G.N. A generalized geometric distribution and some of its properties. Stat. Probab. Lett. 1983, 1, 171-175. [CrossRef]
3. Philippou, A.N. Poisson and compound Poisson distributions of order $k$ and some of their properties. Zap. Nauchnykh Semin. LOMI 1983, 130, 175-180.
4. Philippou, A.N. The negative binomial distribution of order $k$ and some of its properties. Biom. J. 1984, 26, 789-794. [CrossRef]
5. Philippou, A.N. Distributions and Fibonacci polynomials of order $k$, longest runs, and reliability of consecutive- $k$-out-of- $n$ :F systems. In Fibonacci Numbers and Their Applications; Philippou, A.N., Ed.; Reidel: Dordrecht, The Netherland, 1986.
6. Aki, S.; Kuboki, H.; Hirano, K. On discrete distributions of order k. Ann. Inst. Stat Math. 1984, 36, 431-440. [CrossRef]
7. Aki, S. Discrete distributions of order $k$ on a binary sequence. Ann. Inst. Statist. Math. 1985, 37, 205-224. [CrossRef]
8. Ling, K.D. A new class of negative binomial distributions of order k. Stat. Probab. Lett. 1989, 7, 371-376. [CrossRef]
9. Mohanty, S.G. Success runs of length $k$ in Markov dependent trials. ANnals Inst. Stat. Math. 1994, 46, 777-796. [CrossRef]
10. Chang, G.J.; Cui, L.; Hwang, F.K. Reliabilities of Consecutive-k Systems; Kluwer Academic Publishers: Dordrecht, The Netherland, 2000.
11. Johnson, N.L.; Kotz, S.; Balakrishnan, N. Discrete Multivariate Distributions; Wiley: New York, NY, USA, 1997.
12. Shmueli, G.; Cohen, A. Run-related probability functions applied to sampling inspection. Technometrics 2000, 42, 188-202. [CrossRef]
13. Balakrishnan, N.; Koutras, M.V. Runs and Scans with Applications. Wiley Series in Probability and Statistics; Wiley: Chichester, UK, 2002.
14. Fu, J.C.; Lou, W.Y.W. Distribution Theory of Runs and Patterns and Its Applications: A Finite Markov Chain Imbedding Approach; World Scientific: Singapore, 2003.
15. Eryilmaz, S. Geometric distribution of order $k$ with a reward. Stat. Probab. Lett. 2014, 92, 53-58. [CrossRef]
16. Rakitzis, A.C.; Antzoulakos, D. Start-up demonstration tests with three-level classification. Stat. Pap. 2015, 56, 1-21. [CrossRef]
17. Dafnis, S.D.; Makri, F.S.; Philippou, A.N. The reliability of a generalized consecutive system. Appl. Math. Comput. 2019, 359, 186-193. [CrossRef]
18. Sengar, A.S.; Maheshwari, A.; Upadhye, N.S. Time-changed Poisson processes of order k. Stoch. Anal. Appl. 2020, 38, 1-25. [CrossRef]
19. Kwon, Y. A comparison of the method of moments estimator and maximum likelihood estimator for the success probability in the Fibonacci-type probability distribution. Stat. Transit. 2022, 23, 27-47. [CrossRef]
20. Luo, X.H. Poisson distribution of order $k$ and its properties, Kexue Tongbao. Foreign Lang. Ed. 1987, 32, 873-874.
21. Georghiou, C.; Philippou, A.N.; Saghafi, A. On the Poisson distribution of order k. Fibonacci Q. 2013, 51, 44-48.
22. Philippou, A.N. A note on the Poisson distribution of order $k$. Fibonacci Q. 2014, 52, 203-205.
23. Shao, J.; Fu, S. On the modes of the negative binomial distribution of order k. J. Appl. Stat. 2016, 43, 2131-2149. [CrossRef]
24. Georghiou, C.; Philippou, A.N.; Psillakis, Z.M. On the modes of the negative binomial distribution of order $k$, type I. Commun. Stat. Simul. Comput. 2021, 50, 1217-1230. [CrossRef]
25. Philippou, A.N. On multiparameter distributions of order k. Ann. Inst. Stat. Math. 1988, 40, 467-475. [CrossRef]
26. Galliher, H.P.; Morse, P.M.; Simond, M. Dynamics of Two Classes of Continuous-Review Inventory Systems. Oper. Res. 1959, 7, 362-383. [CrossRef]
27. Patel, Y.C. Estimation of the Parameters of the Triple and Quadruple Stuttering-Poisson Distributions. Technometrics 1976, 18, 67-73. [CrossRef]
28. Hirano, K.; Kuboki, H.; Aki, S.; Kuribayashi, A. Figures of probability density functions in statistics II: Discrete univariate case. Comput. Sci. Monogr. 1984, 20, 53-102.

Disclaimer/Publisher's Note: The statements, opinions and data contained in all publications are solely those of the individual author(s) and contributor(s) and not of MDPI and/or the editor(s). MDPI and/or the editor(s) disclaim responsibility for any injury to people or property resulting from any ideas, methods, instructions or products referred to in the content.

