

Article

Identification of Dual-Rate Sampled Hammerstein Systems with a Piecewise-Linear Nonlinearity Using the Key Variable Separation Technique

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Abstract: The identification difficulties for a dual-rate Hammerstein system lie in two aspects. First, the identification model of the system contains the products of the parameters of the nonlinear block and the linear block, and a standard least squares method cannot be directly applied to the model; second, the traditional single-rate discrete-time Hammerstein model cannot be used as the identification model for the dual-rate sampled system. In order to solve these problems, by combining the polynomial transformation technique with the key variable separation technique, this paper converts the Hammerstein system into a dual-rate linear regression model about all parameters (linear-in-parameter model) and proposes a recursive least squares algorithm to estimate the parameters of the dual-rate system. The simulation results verify the effectiveness of the proposed algorithm.

Keywords: Hammerstein system; dual-rate; key variable separation technique; polynomial transformation; least squares

1. Introduction

A traditional discrete-time system is called a single-rate system, in which the input refreshing period equals the output sampling period [1,2]; In some complex nonlinear systems, the sampling rates of the output and the input are different due to the limitation of the measurement technology and method. The system, which has two different input-output operating frequencies, is called a dual-rate system [3,4]. Take the dual-rate warp yarn dyeing process with indigo dye for dyeing blue denim products as an example [5]; the measurement of leuco-indigo concentration in the process needs at least 15 min by the automatic redox titration method, while one can adjust/control the dosage of leuco-indigo more quickly (at any time) by feeding indigo solution with different concentrations into the indigo dyeing bath.

In the area of dual-rate/multirate sampled system identification, Chen proposed three gradient parameter estimation methods for dual-rate sampled systems [4]; Ding *et al.* explored a hierarchical least squares method for dual-rate sampled systems [6]; Liu *et al.* studied a hierarchical identification method for general dual-rate sampled systems [7]. By using T-S (Takagi and Sugeno) fuzzy models, Huang *et al.* proposed a filtering method for multirate nonlinear sampled-data systems [8].

The piecewise-linear characteristic is often encountered in control systems, either alone or in cascade with linear dynamic systems to describe processes operating with different gains in different input intervals [9–11], such as in the nonlinear servomechanism [10], in the heating and cooling processes [11], *etc.* It is also well known that piecewise-linear functions can be used as a general tool to approximate nonlinear functions [10]. In previous applications, Hammerstein systems with a two-segment piecewise-linear nonlinearity have been studied in [12], and an extension to a discontinuous two-segment piecewise-linearity with preloads and dead-zones nonlinearity was discussed [12]. Vörös extended the key variable separation identification method for a Hammerstein system with a two-segment piecewise-linear nonlinearity [13] to a Hammerstein system with a multi-segment piecewise-linear characteristic [14] and with a time-varying backlash [15].

This paper deal with dual-rate Hammerstein systems with a two-segment piecewise-linear nonlinearity in cascade with a linear dynamic system. In the literature of single-rate nonlinear Hammerstein system identification, Deng and Ding studied a Newton iterative identification method for an input nonlinear finite impulse response system with moving average noise using the key variable separation technique [16]. Li explored a maximum likelihood estimation algorithm for Hammerstein CARARMA systems based on the Newton iteration [17]. Wang and Tang developed an auxiliary model-based recursive least squares algorithm for Hammerstein linear-in-parameters output error moving average systems [18]. Salimifard *et al.* presented iterative algorithms to identify nonlinear MIMO (Multi-Input Multi-Output) Hammerstein systems with moving average noises [19].

Due to the structure of a nonlinear block plus a linear block, the identification model of a Hammerstein system naturally contains the products of the parameters of the nonlinear block and the linear block. Moreover, the traditional single-rate discrete-time model is not suitable for the dual-rate sampled-data of the Hammerstein system. These bring difficulties to directly using a standard least squares method. The intent of this paper is to study identification methods of Hammerstein nonlinear systems with dual-rate sampling period in input-output signals. The contribution of this paper is that, by combining

the polynomial transformation technique [20] with the key variable separation technique [12], a dual-rate linear-in-parameter identification model for the dual-rate sampled Hammerstein system is derived, which is suitable for the dual-rate sampled-data, and it is easy to use the standard least squares method to identify the system parameters.

The organization of this paper is as follows. Section 2 describes the problem formulation of a dual-rate Hammerstein system. Section 3 transforms the Hammerstein system into a dual-rate linear-in-parameter identification model using the polynomial transformation technique and the key variable separation technique. Section 4 derives the key variable separation-based recursive least squares estimation algorithms to estimate the parameters of the dual-rate/single-rate models. Section 5 analyzes the convergence performance of the presented least squares algorithm. Section 6 provides an experiment to verify the effectiveness of the proposed algorithm. Finally, some concluding remarks are summarized in Section 7.

2. The Description of the Dual-Rate Hammerstein System

Let us introduce some notation. The symbol I stands for an identity matrix of appropriate size; the superscript T denotes the matrix/vector transpose. $\hat{x}(t)$ stands for the estimate of x at time t.

A Hammerstein system with two sampling periods in input-output signals is depicted in Figure 1, where P_c is a continuous-time Hammerstein process. Under the dual-rate framework, the discrete control signal $u(kT_1)$ is generated by a computer; the input u(t) is produced by a zero-order hold H_{T_1} with period $T_1 = T$ and takes the piecewise constant values u(kT) within the updating intervals; the output signal y(t) is sampled by a sampler S_{T_2} with period $T_2 := qT$, yielding a discrete-time signal y(kqT). That means that the multiple of the output sampling period $T_2 := qT$ to the input sampling period $T_1 = T$ is $T_2 = T_1$ and $T_2 = T_2$ is written as $T_2 = T_1$ is written as $T_2 = T_2$ is written as T_2

$$u(k)$$
 H_{T_1} $u(t)$ P_c $y(t)$ S_{T_2} $y(kq)$

[u(kT) written as u(k); y(kqT) written as y(kq)

Figure 1. The dual-rate Hammerstein system.

Assume that the corresponding discrete-time singe-rate Hammerstein system consists of a two-segment piecewise-linear nonlinearity with slopes m_1 and m_2 and a linear model with polynomials A(z) and B(z) as depicted in Figure 2,

$$x(k) = \begin{cases} m_1 u(k), \ u(k) \ge 0, \\ m_2 u(k), \ u(k) < 0, \end{cases}$$
 (1)

$$A(z)y(k) = B(z)x(k), (2)$$

where x(k) is the internal variable; A(z) and B(z) are defined by:

$$A(z) := 1 + a_1 z^{-1} + a_2 z^{-2} + \dots + a_n z^{-n},$$

$$B(z) := b_0 + b_1 z^{-1} + b_2 z^{-2} + \dots + b_n z^{-n}.$$

$$\begin{array}{c|c}
u(k) & \hline m_1 & \\
\hline m_2 & \hline x(k) & \hline B(z) & y(k) \\
\hline M(kT) & \hline M(k), y(kT) & \hline M(k) \\
\end{array}$$

Figure 2. The discrete-time Hammerstein system.

Introduce a switching function h(k) as:

$$h(k) = h[u(k)] = 0.5\{1 + \operatorname{sgn}[u(k)]\},\$$

where:

$$sgn(u) = \begin{cases} 1, & u > 0, \\ 0, & u = 0, \\ -1, & u < 0. \end{cases}$$

Then, Equation (1) can be rewritten as:

$$x(k) = m_2 u(k) + (m_1 - m_2)h(k)u(k).$$
(3)

From Equation (2), we have:

$$y(k) = b_0 x(k) + b_1 x(k-1) + \dots + b_n x(k-n) - a_1 y(k-1) - a_2 y(k-2) - \dots - a_n y(k-n).$$
 (4)

Inserting Equation (3) into x(k-i) in Equation (4) gives:

$$y(k) = \sum_{i=0}^{n} b_i [m_2 u(k-i) + (m_1 - m_2) h(k-i) u(k-i)] - \sum_{i=1}^{n} a_i y(k-i).$$

Define:

$$\mathbf{U}(k) := \begin{bmatrix} u(k) & h(k)u(k) \\ u(k-1) & h(k-1)u(k-1) \\ \vdots \\ u(k-n) & h(k-n)u(k-n) \end{bmatrix} \in \mathbb{R}^{(n+1)\times 2},
\mathbf{\phi}(k) := [-y(k-1), -y(k-2), \cdots, -y(k-n)]^{\mathsf{T}} \in \mathbb{R}^{n},
\mathbf{a} := [a_{1}, a_{2}, \cdots, a_{n}]^{\mathsf{T}} \in \mathbb{R}^{n},
\bar{\mathbf{b}} := [b_{0}, b_{1}, b_{2}, \cdots, b_{n}]^{\mathsf{T}} \in \mathbb{R}^{n+1},
\mathbf{M} := [m_{2}, m_{1} - m_{2}]^{\mathsf{T}} \in \mathbb{R}^{2}.$$

Then, we have:

$$y(k) = \bar{\boldsymbol{b}}^{\mathrm{T}} \boldsymbol{U}(k) \boldsymbol{M} + \boldsymbol{\phi}^{\mathrm{T}}(k) \boldsymbol{a},$$

Replacing k with kq gives:

$$y(kq) = \bar{\boldsymbol{b}}^{\mathrm{T}} \boldsymbol{U}(kq) \boldsymbol{M} + \boldsymbol{\phi}^{\mathrm{T}}(kq) \boldsymbol{a}. \tag{5}$$

The identification problems caused by dual-rate Hammerstein systems with piecewise-linear nonlinearity exist in two aspects.

- For the dual-rate Hammerstein system in this paper, all of the input data $\{u(k): k=0,1,2,\cdots\}$ are measurable; only a part of the output data $\{y(kq): k=0,1,2,\cdots\}$ can be measured, where $q\geq 2$ is a positive integer. Intersample outputs (missing outputs) $\{y(kq-i): k=0,1,2,\cdots,0 < i < q\}$ are unknown. Because there exist unavailable outputs y(kq-i) in the information vector $\phi(kq)$, with kq-i not equaling the integer multiples of q, the single-rate model in Equation (5) cannot be used as the identification model for the dual-rate system.
- Moreover, the single-rate model in Equation (5) contains the products of the parameters of the nonlinear block and the linear block, *i.e.*, the output of the model is a bilinear function about the parameter vectors \bar{b} and M, and a standard least squares method cannot be directly applied to the model.

The intent of this paper is to convert the Hammerstein system in Equations (1) and (2) into a dual-rate linear regression model about all parameters (linear-in-parameter model), by combining the polynomial transformation technique [20] with the key variable separation technique [12]. The obtained dual-rate linear regression model is suitable for the dual-rate sampled data, and a standard identification technique can be directly applied to the model to estimate the parameters of the model.

3. The Dual-Rate Identification Model of the Hammerstein System

In this section, we transform the Hammerstein system in Equations (1) and (2) into a dual-rate linear-in-parameter identification model, which is suitable for the dual-rate sampled data, by using the polynomial transformation technique [20] and the key variable separation technique [12].

Referring to [20], assume that the roots of A(z) are z_i , $i = 1, 2, \dots, n$. Then, we have:

$$A(z) = (1 - z_1 z^{-1})(1 - z_2 z^{-1}) \cdots (1 - z_n z^{-1}).$$

Define:

$$\epsilon(z) := (1 + z_1 z^{-1} + z_1^2 z^{-2} + \dots + z_1^{q-1} z^{-q+1}) (1 + z_2 z^{-1} + z_2^2 z^{-2} + \dots + z_2^{q-1} z^{-q+1}) \dots
(1 + z_n z^{-1} + z_n^2 z^{-2} + \dots + z_n^{q-1} z^{-q+1})
= 1 + \epsilon_1 z^{-1} + \epsilon_2 z^{-2} + \dots + \epsilon_{n(q-1)} z^{-n(q-1)}.$$

Multiplying both sides of Equation (2) by $\epsilon(z)$ gives:

$$\alpha(z)y(k) = \beta(z)x(k).$$

Using the identity:

$$1 - z_i^q z^{-q} \equiv (1 - z_i z^{-1})(1 + z_i z^{-1} + z_i^2 z^{-2} + \dots + z_i^{q-1} z^{-(q-1)}),$$

one obtains:

$$\alpha(z) := \epsilon(z)A(z)$$

$$= (1 + z_{1}z^{-1} + z_{1}^{2}z^{-2} + \dots + z_{1}^{q-1}z^{-(q-1)})(1 - z_{1}z^{-1})$$

$$(1 + z_{2}z^{-1} + z_{2}^{2}z^{-2} + \dots + z_{2}^{q-1}z^{-(q-1)})(1 - z_{2}z^{-1})$$

$$\dots$$

$$(1 + z_{n}z^{-1} + z_{n}^{2}z^{-2} + \dots + z_{n}^{q-1}z^{-(q-1)})(1 - z_{n}z^{-1})$$

$$= (1 - z_{1}^{q}z^{-q})(1 - z_{2}^{q}z^{-q}) \dots (1 - z_{n}^{q}z^{-q})$$

$$=: 1 + \alpha_{1}z^{-q} + \alpha_{2}z^{-2q} + \dots + \alpha_{n}z^{-nq},$$

$$\beta(z) := \epsilon(z)B(z)$$

$$= (1 + \epsilon_{1}z^{-1} + \epsilon_{2}z^{-2} + \dots + \epsilon_{n(q-1)}z^{-n(q-1)})(b_{0} + b_{1}z^{-1} + b_{2}z^{-2} + \dots + b_{n}z^{-n})$$

$$=: \beta_{0} + \beta_{1}z^{-1} + \beta_{2}z^{-2} + \dots + \beta_{nq}z^{-nq}.$$

$$(6)$$

Consider a disturbance in a practical system; introducing a zero-mean random white noise v(k), we have:

$$\alpha(z)y(k) = \beta(z)x(k) + v(k). \tag{8}$$

Substituting the polynomials $\alpha(z)$ and $\beta(z)$ into Equation (8) gives:

$$y(k) = [\beta_0 + \beta_1 z^{-1} + \beta_2 z^{-2} + \dots + \beta_{nq} z^{-nq}] x(k) - [\alpha_1 z^{-q} + \alpha_2 z^{-2q} + \dots + \alpha_n z^{-nq}] y(k) + v(k),$$

$$= \beta_0 x(k) + \sum_{i=1}^{nq} \beta_i x(k-i) - \sum_{i=1}^{n} \alpha_i y(k-iq) + v(k).$$
(10)

In order to get a unique solution, we assume that $\beta_0 = b_0 = 1$ [21]. The first term x(k) (its coefficient is one) on the right-hand side is a separated key variable, and x(k-i) $(i=1,2,\cdots,n_{\beta})$ are not separated variables. Substituting (3) into the separated key variable x(k) in Equation (10) gives:

$$y(k) = m_2 u(k) + (m_1 - m_2)h(k)u(k) + \sum_{i=1}^{n_q} \beta_i x(k-i) - \sum_{i=1}^n \alpha_i y(k-iq) + v(k).$$

Define the information vector $\psi(k)$ and the parameter vector Θ as:

$$\psi(k) := [u(k), h(k)u(k), x(k-1), x(k-2), \cdots, x(k-nq), -y(k-q), -y(k-2q), \cdots, \\ -y(t-nq)]^{\mathsf{T}} \in \mathbb{R}^{nq+n+2},$$

$$\Theta := [m_2, m_1 - m_2, \beta_1, \beta_2, \cdots, \beta_{nq}, \alpha_1, \alpha_2, \cdots, \alpha_n]^{\mathsf{T}} \in \mathbb{R}^{nq+n+2}.$$

Then, we have:

$$y(k) = \boldsymbol{\psi}^{\mathrm{T}}(k)\boldsymbol{\Theta} + v(k).$$

Replacing k with kq gives the following dual-rate linear-in-parameter identification model,

$$y(kq) = \boldsymbol{\psi}^{\mathsf{T}}(kq)\boldsymbol{\Theta} + v(kq), \tag{11}$$

with:

$$\psi(kq) := [u(kq), h(kq)u(kq), x(kq-1), x(kq-2), \cdots, x(kq-nq), -y(kq-q), -y(kq-2q), \cdots, -y(kq-nq)]^{\mathsf{T}}.$$
(12)

The output of the obtained model is linear about all parameters. Note that y(kq - iq) is available by sampling the output signal in every qT interval, and y(kq) in Equation (11) is a linear function about Θ . Thus, it is easy to estimate the parameter vector Θ by using the standard least squares method.

4. The Dual-Rate/Single-Rate Parameter Estimation

In this section, the parameter estimates (in the following $\hat{\Theta}(kq)$) of the dual-rate model are computed by a recursive least squares algorithm; then by comparing the coefficients of z^{-i} on both sides of a coefficient polynomial equation containing the parameters of the dual-rate/single-rate models, the parameter estimates (in the following $\hat{\theta}(kq)$) of the single-rate model are derived.

Suppose that the data length $N \gg nq + n + 2$. Define a quadratic cost function,

$$J(\boldsymbol{\Theta}) := \sum_{k=1}^N [y(kq) - \boldsymbol{\psi}^{\mathsf{T}}(kq)\boldsymbol{\Theta}]^2.$$

The information vector $\psi(kq)$ in $J(\Theta)$ contains the unknown internal variables $\{x(kq-i), i=1,2,\cdots,nq\}$. Here, the solution is to replace x(kq-i) with its estimate $\hat{x}(kq-i)$; then the estimate of $\psi(kq)$ can be written as:

$$\hat{\psi}_{(kq)} := [u(kq), h(kq)u(kq), \hat{x}(kq-1), \hat{x}(kq-2), \cdots, \hat{x}(kq-nq), -y(kq-q), -y(kq-2q), \cdots, -y(kq-nq)]^{\mathsf{T}}.$$
(13)

the estimate $\hat{x}(kq-i)$ can be computed by replacing m_1 and m_2 in Equation (3) with their estimates $\hat{m}_1(kq)$ and $\hat{m}_2(kq)$,

$$\hat{x}(kq - i) = \hat{m}_2(kq)u(kq - i) + [\hat{m}_1(kq) - \hat{m}_2(kq)]h(kq - i)u(kq - i).$$

Minimizing the cost functions $J(\Theta)$, replacing x(kq-i) in $\psi(kq)$ with its estimate $\hat{x}(kq-i)$ and by using the least squares method, we obtain the following key variable separation-based recursive least squares algorithm for estimating Θ of the dual-rate Hammerstein system,

$$\hat{\boldsymbol{\Theta}}(kq) = \hat{\boldsymbol{\Theta}}(kq - q) + \boldsymbol{L}(kq)[y(kq) - \hat{\boldsymbol{\psi}}^{\mathsf{T}}(kq)\hat{\boldsymbol{\Theta}}(kq - q)], \tag{14}$$

$$\boldsymbol{L}(kq) = \frac{\boldsymbol{P}(kq-q)\hat{\boldsymbol{\psi}}(kq)}{1+\hat{\boldsymbol{\psi}}^{\mathrm{T}}(kq)\boldsymbol{P}(kq-q)\hat{\boldsymbol{\psi}}(kq)},$$
(15)

$$\mathbf{P}(kq) = [\mathbf{I} - \mathbf{L}(kq)\hat{\boldsymbol{\psi}}^{\mathsf{T}}(kq)]\mathbf{P}(kq - q), \ \mathbf{P}(0) = p_0\mathbf{I},$$
(16)

(17)

$$\hat{\psi}(kq) = [u(kq), h(kq)u(kq), \hat{x}(kq-1), \hat{x}(kq-2), \cdots, \hat{x}(kq-nq), \\
-y(kq-q), -y(kq-2q), \cdots, -y(kq-nq)]^{\mathsf{T}},$$

$$\hat{x}(kq+i) = \hat{m}_2(kq)u(kq+i) + [\hat{m}_1(kq) - \hat{m}_2(kq)]h(kq+i)u(kq+i), \tag{18}$$

$$\hat{\Theta}(kq) = [\hat{m}_2(kq), \hat{m}_1(kq) - \hat{m}_2(kq), \hat{\beta}_1(kq), \hat{\beta}_2(kq), \cdots, \hat{\beta}_{nq}(kq),$$

$$\hat{\alpha}_1(kq), \alpha_2(kq), \cdots, \alpha_n(kq)]^{\mathrm{T}}. \tag{19}$$

Use the obtained $\Theta(kq)$ to construct the polynomials,

$$\hat{\alpha}(kq,z) = 1 + \hat{\alpha}_1(kq)z^{-q} + \hat{\alpha}_2(kq)z^{-2q} + \dots + \hat{\alpha}_n(kq)z^{-nq},$$

$$\hat{\beta}(kq,z) = 1 + \hat{\beta}_1(kq)z^{-1} + \hat{\beta}_2(kq)z^{-2} + \dots + \hat{\beta}_{nq}(kq)z^{-nq}.$$

From Equations (6) and (7), we can get:

$$\frac{\alpha(z)}{\beta(z)} = \frac{A(z)}{B(z)}.$$

Let the estimates of A(z) and B(z) at time kq be:

$$\hat{A}(kq,z) = 1 + \hat{a}_1(kq)z^{-1} + \hat{a}_2(kq)z^{-2} + \dots + \hat{a}_n(kq)z^{-n},$$

$$\hat{B}(kq,z) = 1 + \hat{b}_1(kq)z^{-1} + \hat{b}_2(kq)z^{-2} + \dots + \hat{b}_n(kq)z^{-n}.$$

According to the model equivalence principle, let:

$$\frac{\hat{\alpha}(kq,z)}{\hat{\beta}(kq,z)} = \frac{\hat{A}(kq,z)}{\hat{B}(kq,z)},$$

then we have

$$\hat{\alpha}(kq,z)\hat{B}(kq,z) = \hat{\beta}(kq,z)\hat{A}(kq,z).$$

By comparing the coefficients of z^{-i} on both sides of this equation, we can establish a series of linear equations about $\hat{a}_i(kq)$ and $\hat{b}_i(kq)$ and build a matrix equation as:

$$T(kq)\hat{\boldsymbol{\theta}}(kq) = \boldsymbol{\rho}(kq),$$
 (20)

where:

$$\hat{\boldsymbol{\theta}}(kq) = [\hat{a}_{1}(kq), \hat{a}_{2}(kq), \cdots, \hat{a}_{n}(kq), \hat{b}_{1}(kq), \hat{b}_{2}(kq), \cdots, \hat{b}_{n}(kq)]^{\mathsf{T}} \in \mathbb{R}^{2n},$$

$$\boldsymbol{\rho}(kq) = [\hat{\beta}_{1}(kq) - \lambda_{1}(kq), \hat{\beta}_{2}(kq) - \lambda_{2}(kq), \cdots, \hat{\beta}_{nq}(kq) - \lambda_{nq}(kq), 0, \cdots, 0]^{\mathsf{T}} \in \mathbb{R}^{nq+n}, (21)$$

$$\boldsymbol{T}(kq) = [-\boldsymbol{T}_{\beta}(kq) \, \boldsymbol{T}_{\alpha}(kq)] \in \mathbb{R}^{(nq+n)\times 2n}, \qquad (22)$$

$$\begin{bmatrix}
1 & 0 & \cdots & 0 \\
\lambda_{1}(kq) & 1 & \vdots \\
\lambda_{2}(kq) & \lambda_{1}(kq) & \ddots & 0 \\
\vdots & \lambda_{2}(kq) & \ddots & 1 \\
\vdots & \ddots & \lambda_{1}(kq) \\
\lambda_{nq}(kq) & \lambda_{2}(kq)
\end{bmatrix}$$

$$\boldsymbol{\tau}_{\alpha}(kq) = \begin{bmatrix}
\vdots & \ddots & \lambda_{1}(kq) \\
0 & \lambda_{nq}(kq) & \vdots \\
\vdots & \ddots & \vdots \\
0 & \cdots & \lambda_{nq}(kq)
\end{bmatrix}$$

$$\boldsymbol{\tau}_{\alpha}(kq) = \begin{bmatrix}
\vdots & \ddots & \ddots & \vdots \\
0 & \cdots & \lambda_{nq}(kq)
\end{bmatrix}$$

$$\lambda_{i}(kq) = \begin{cases} \hat{\alpha}_{j}(kq), & i = jq, j = 1, 2, \cdots, n, \\ 0, & others, \end{cases}$$

$$\mathbf{T}_{\beta}(kq) = \begin{cases} 1 & 0 & \cdots & 0 \\ \hat{\beta}_{1}(kq) & 1 & 0 \\ \hat{\beta}_{2}(kq) & \hat{\beta}_{1}(kq) & \vdots \\ \vdots & \hat{\beta}_{2}(kq) & \ddots & 1 \\ \vdots & \ddots & \hat{\beta}_{1}(kq) \\ \hat{\beta}_{nq}(kq) & & \hat{\beta}_{2}(kq) \\ 0 & \hat{\beta}_{nq}(kq) & \vdots \\ \vdots & & \ddots & \vdots \\ 0 & \cdots & 0 & \hat{\beta}_{nq}(kq) \end{cases}$$

$$(24)$$

Then, the least squares solution for Equation (20) is given by:

$$\hat{\boldsymbol{\theta}}(kq) = [\boldsymbol{T}^{\mathsf{T}}(kq)\boldsymbol{T}(kq)]^{-1}\boldsymbol{T}^{\mathsf{T}}(kq)\boldsymbol{\rho}(kq). \tag{26}$$

The computation process for estimates $\hat{\Theta}(kq)$ and $\hat{\theta}(kq)$ of the key variable separation-based recursive least squares algorithm for the dual-rate Hammerstein system is summarized as follows:

Step 1. To initialize, let k = 1 and $\hat{\Theta}(0) = \mathbf{1}_{nq+n+2}/p_0$, and $P(0) = p_0 I$, $\hat{x}(kq) = 0$, u(k) = 0, y(k) = 0, for $k \le 0$, $p_0 = 10^6$.

Step 2. Collect the input-output data u(k) and y(kq), and form $\hat{\psi}(kq)$ by Equation (17).

Step 3. Compute L(kq) by Equation (15) and P(kq) by Equation (16).

Step 4. Update the dual-rate parameter estimate $\hat{\Theta}(kq)$ by Equation (14).

Step 5. Compute $\hat{x}(kq+i)$ by Equation (18).

Step 6. Form $\rho(kq)$, $T_{\alpha}(kq)$, $T_{\beta}(kq)$ and T(kq) by Equation (21), Equation (23), Equation (25) and Equation (22).

Step 7. Update the single-rate parameter estimate $\hat{\theta}(kq)$ by Equation (26).

Step 8. Increase k by one and go to Step 2.

The procedure of computing the estimates $\hat{\Theta}(kq)$ and $\hat{\theta}(kq)$ of the key variable separation-based recursive least squares algorithm is shown in Figure 3.

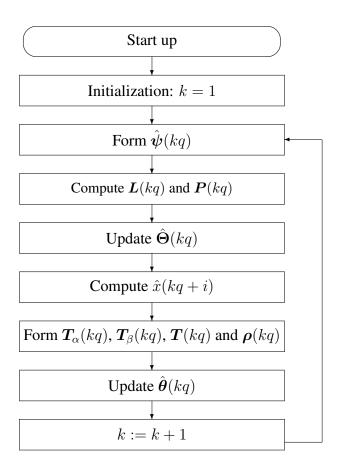


Figure 3. The flowchart for computing the estimates $\hat{\Theta}(kq)$ and $\hat{\theta}(kq)$.

5. The Convergence Analysis

The convergence analysis of the presented key variable separation-based recursive least squares algorithm for a dual-rate Hammerstein system is simply explained as follows.

Assume that the σ algebra sequence $\mathcal{F}_{kq} = \sigma(v(kq), v(kq-1), v(kq-2), \cdots)$ generated by v(kq), and $\{v(kq), \mathcal{F}_{kq}\}$ is a martingale difference sequence on a probability space $\{\Omega, \mathcal{F}, P\}$ [22]. The sequence $\{v(kq)\}$ satisfies:

(A1)
$$E[v(kq)|\mathcal{F}_{kq-q}] = 0, \text{ a.s.}$$

(A2)
$$E[v^2(kq)|\mathcal{F}_{kq-q}] = \sigma^2(kq) \leqslant \bar{\sigma}^2 < \infty$$
, a.s.

(A3)
$$\limsup_{t \to \infty} \frac{1}{t} \sum_{i=q}^{kq} v^2(i) \leqslant \bar{\sigma}^2 < \infty, \text{ a.s.}$$

Theorem 1: For the key variable separation-based recursive least squares algorithm in this paper, assume that (A1)–(A3) hold and that the estimated information vector $\hat{\psi}(kq)$ and matrix T(kq) are

persistently exciting, *i.e.*, there exist constants $0 < C_1 \le C_2 < \infty$ and an integer $N > n_0$ $(n_0 = nq + n + 2)$, such that for $k > n_0$, the following strong persistent excitation conditions hold:

(A4)
$$C_1 \mathbf{I} \leqslant \frac{1}{N} \sum_{k=0}^{N-1} \hat{\boldsymbol{\psi}}(kq) \hat{\boldsymbol{\psi}}^{\mathrm{T}}(kq) \leqslant C_2 \mathbf{I}$$
, a.s.

(A5)
$$C_1 \mathbf{I} \leqslant \frac{1}{N} \sum_{k=0}^{N-1} \mathbf{T}^{\mathsf{T}}(kq) \mathbf{T}(kq) \leqslant C_2 \mathbf{I}$$
, a.s.

Then, the parameter estimation vectors $\hat{\Theta}(kq)$ and $\hat{\theta}(kq)$ consistently converge to the true parameter vectors Θ and θ .

In a similar way to the method in [23], Theorem 1 can be proven, and its proof is omitted here.

6. Experiment

Consider the following Hammerstein system:

$$A(z)y(k) = B(z)x(k) + v(k),$$

$$x(k) = \begin{cases} m_1u(k) & u(k) \ge 0 \\ m_2u(k) & u(k) < 0 \end{cases},$$

$$A(z) = 1 + a_1z^{-1} + a_2z^{-2}$$

$$= 1 + 0.20z^{-1} - 0.35z^{-2},$$

$$B(z) = 1 + b_1z^{-1} + b_2z^{-2}$$

$$= 1 + 0.80z^{-1} + 0.60z^{-2},$$

$$m_1 = 1.50, m_2 = -1.00.$$

Suppose q = 2; then, we have:

$$\varepsilon(z) = 1 - 0.20z^{-1} - 0.35z^{-2}.$$

The corresponding dual-rate model can be expressed as:

$$\alpha(z)y(k) = \beta(z)x(k),$$

$$\alpha(z) = 1 + \alpha_1 z^{-2} + \alpha_2 z^{-4}$$

$$= 1 - 0.74z^{-2} + 0.1225z^{-4},$$

$$\beta(z) = 1 + \beta_1 z^{-1} + \beta_2 z^{-2} + \beta_3 z^{-3} + \beta_4 z^{-4}$$

$$= 1 + 0.60z^{-1} + 0.09z^{-2} - 0.40z^{-3} - 0.21z^{-4},$$

$$\boldsymbol{a} = [0.20, -0.35]^{\mathrm{T}}, \, \boldsymbol{b} = [0.80, 0.60]^{\mathrm{T}},$$

$$\boldsymbol{\theta} = \begin{bmatrix} \boldsymbol{a} \\ \boldsymbol{b} \end{bmatrix},$$

$$\boldsymbol{\alpha} = [-0.74, 0.1225]^{\mathrm{T}}, \, \boldsymbol{\beta} = [0.60, 0.09, -0.40, -0.21]^{\mathrm{T}},$$

$$\boldsymbol{\Theta} = [m_2, m_1 - m_2, \boldsymbol{\beta}^{\mathrm{T}}, \boldsymbol{\alpha}^{\mathrm{T}}]^{\mathrm{T}}.$$

In the simulation experiment, the input $\{u(t)\}$ is taken as an uncorrelated persistently-excited normal distribution signal sequence with zero mean and unit variance, and $\{v(t)\}$ is taken as a white noise sequence with zero mean and variance $\sigma^2=0.50^2$ and $\sigma^2=1.00^2$. Apply the proposed algorithm to estimate the parameters of this system. The parameter estimation errors of the dual-rate model are:

$$\delta := \sqrt{\frac{\|\hat{m}_1(kq) - m_1\|^2 + \|\hat{m}_2(kq) - m_2\|^2 + \|\hat{\boldsymbol{\beta}}(kq) - \boldsymbol{\beta}\|^2 + \|\hat{\boldsymbol{\alpha}}(kq) - \boldsymbol{\alpha}\|^2}{m_1^2 + m_2^2 + \|\boldsymbol{\beta}\|^2 + \|\boldsymbol{\alpha}\|^2}} \times 100\%,$$

the parameter estimation errors of the single-rate model are:

$$\delta := \sqrt{\frac{\|\hat{m}_1(kq) - m_1\|^2 + \|\hat{m}_2(kq) - m_2\|^2 + \|\hat{\boldsymbol{a}}(kq) - \boldsymbol{a}\|^2 + \|\hat{\boldsymbol{b}}(kq) - \boldsymbol{b}\|^2}{m_1^2 + m_2^2 + \|\boldsymbol{a}\|^2 + \|\boldsymbol{b}\|^2}} \times 100\%.$$

The parameter estimates and their errors are shown in Tables 1 and 2. Obviously, the parameter estimation errors become (generally) smaller and smaller with the data length k increasing. This shows that the proposed algorithm is effective.

σ^2		0.50^{2}				1.00^{2}		
t	100	1000	2000	3000	100	1000	2000	3000
m_1 = 1.5000	1.64510	1.49306	1.48069	1.48674	1.64332	1.47369	1.45489	1.46773
m_2 =-1.0000	-1.10137	-1.00671	-0.98836	-0.99494	-1.07085	-0.98956	-0.96557	-0.98115
β_1 = 0.6000	0.44924	0.56619	0.58761	0.60112	0.46040	0.54639	0.58138	0.60706
$\beta_2 = 0.0900$	-0.02088	0.04436	0.07612	0.08695	-0.07241	0.04529	0.09448	0.10813
$\beta_3 = -0.4000$	-0.10421	-0.33348	-0.35710	-0.36657	-0.13676	-0.33817	-0.35584	-0.36336
$\beta_4 = -0.2100$	-0.16787	-0.18590	-0.20263	-0.20855	-0.14399	-0.16044	-0.19223	-0.20651
$\alpha_1 = -0.7400$	-0.70650	-0.75683	-0.74221	-0.73669	-0.77498	-0.75709	-0.72961	-0.72063
α_2 = 0.1225	0.15543	0.14357	0.12683	0.12660	0.18953	0.13963	0.11305	0.11506
δ (%)=0.0000	18.97213	4.54356	2.51473	1.76071	18.54456	5.34749	3.71454	2.84976

Table 1. The dual-rate parameter estimates and errors.

Table 2. The single-rate parameter estimates and errors.

σ^2		0.50^{2}				1.00^{2}		
t	100	1000	2000	3000	100	1000	2000	3000
m_1 = 1.5000	1.64332	1.47369	1.45489	1.46773	1.64510	1.49306	1.48069	1.48674
m_2 =-1.0000	-1.07085	-0.98956	-0.96557	-0.98115	-1.10137	-1.00671	-0.98836	-0.99494
a_1 = 0.2000	-0.13036	0.18131	0.18225	0.16469	-0.16282	0.13686	0.15786	0.15793
$a_2 = -0.3500$	-0.28358	-0.40123	-0.36730	-0.34446	-0.21271	-0.36634	-0.35419	-0.34545
b_1 = 0.8000	0.33046	0.72731	0.76196	0.77122	0.28415	0.70397	0.74542	0.75947
b_2 = 0.6000	0.36387	0.50014	0.56053	0.58345	0.40609	0.51357	0.55673	0.57349
δ (%)=0.0000	30.68012	6.56738	3.93705	2.92556	33.17133	6.90524	4.02630	3.13498

7. Conclusions

The solution to overcoming the identification difficulties of the dual-rate Hammerstein system is to derive a dual-rate identification model that must be suitable for the dual-rate sampled data, and it is easy to use a standard least squares method. By combining the polynomial transformation technique with the key variable separation technique, this paper converts the discrete-time Hammerstein system into a dual-rate linear regression model about all parameters, which is suitable for the dual-rate sampled data and presents a recursive least squares algorithm for the dual-rate linear-in-parameter model to get the parameters of the dual-rate model. Finally, the parameters of the single-rate model are computed by a least squares algorithm from the obtained parameters of the dual-rate model. The proposed method can be used in the soft measurement of dual-rate nonlinear systems.

Author Contributions

Ying-Ying Wang prepared the manuscript. Xiang-Dong Wang assisted in the work and the simulation. Dong-Qing Wang was in charge of the overall research and critical revision of the paper.

Conflicts of Interest

The authors declare no conflict of interest.

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