

Article

A Regional Approach for Investigation of Temporal Precipitation Changes

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Abstract: Climatic variability is one of the fundamental aspects of the climate. Our scope of knowledge of this variability is limited by unavailable long-term high-resolution spatial data. Climatic simulations indicate that warmer climate increases extreme precipitations but decreases high-frequency temperature variability. As an important climatologic variable, the precipitation is reported by the IPCC to increase in mid and high altitudes and decrease in subtropical areas. On a regional scale, such a change needs spatio-parametric justification. In this regard, a regionalization approach relying on frequency characteristics and parameters of heavy precipitation may provide better insight into temporal precipitation changes, and thus help us to understand climatic variability and extremes. This study introduces the “index precipitation method”, which aims to define hydrologic homogeneous regions throughout which the frequency distribution of monthly maximum hourly precipitations remains the same and, therefore, investigate whether there are significant temporal precipitation changes in these regions. Homogenous regions are defined based on L-moment ratios of frequency distributions via cluster analysis and considering the spatial contiguity of gauging sites via GIS. Regarding the main hydrologic characteristics of heavy precipitation, 12 indices are defined in order to investigate the existence of regional trends by means of *t*- and Mann–Kendall tests for determined homogenous regions with similar frequency behaviors. The case study of Japan, using hourly precipitation data on 150 gauges for 1991–2010, shows that trends that statistically exist for single-site observations should be regionally proved. Trends of heavy precipitation have region-specific properties across Japan. Homogenous regions beneficially define statistically significant trends for heavy precipitation.



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Keywords: regional analysis; homogenous region; L-moment; index precipitation; climatic variability

1. Introduction

Climatic variability has been recognized as an important problem causing fundamental changes to human life, as well as the natural environment [1,2]. In order to address how this variability has developed and will affect available water resources, one of the main climatologic quantities needing to be investigated is precipitation [3–6]. Considering the possible effects of climate change on precipitation, many researchers have analyzed and tested the existence of trends in precipitation using basic statistics and updated data for different parts of the world [7–9]. Although these studies support, on a global scale, the IPCC report [10–12], claiming that precipitation is set to increase in the Northern Hemisphere, mid and high latitudes, but to decrease in subtropical land areas [13] on a local/regional scale, their results are generally of high spatial variation and lack areal support from vicinity sites. In fact, the consistency of calculated hydrologic trends cannot be realistically determined solely by using single-site observations, due to the biases resulting from data-related problems. This problem is more pronounced for the analyses relying on average statistics [14], since average statistics are not highly sensitive to climatic variability [15,16]. Considering that societies and ecosystems are indeed much more vulnerable to extremes compared to averages, a regional approach based on the regional

frequency behavior of heavy precipitations and a regional trend indicator based on the multitude of trend indices relying on the basic characteristics of heavy precipitation will be more helpful to verify the existence and consistency of a regional trend of precipitation. This should provide better insight into temporal precipitation changes, as well as climatic variability and extremes.

There are many approaches in the literature for the investigation of the spatial pattern of rainfall, relying on mathematical transformations, statistical regressions, or clustering techniques, such as multivariate regression, spatial interpolation, spatial correlation, harmonic analysis, L-moments, principal component analysis, hierarchical and non-hierarchical clustering techniques [17–19]. These approaches use basic basin physiographic and climatic properties for defining the rainfall pattern in a considered area and have been studied by many researchers to obtain a robust methodology that can be applied within a defined homogenous region [20–23]. One of the eminent methodologies in the literature is the index flood method [24,25]. The index flood method derives homogeneous regions based on regional frequency behaviors and uses local statistics to obtain scaled estimations. Therefore, it preserves both the regional frequency behavior and local statistics of observations within the boundary of the homogenous region. It assumes a common dimensionless regional frequency curve along the homogenous region. The product of this regional frequency curve and the index flood, which is generally taken as the mean of the station data, yields an at-site frequency curve for a considered flow section. Since this method is first developed for flood estimations, it is known as the index flood method [18,24]. On the other hand, ordinary clustering approaches directly use site-specific basin or climatic parameters to define homogenous regions; therefore, such classification approaches are generally quite sensitive to the parameters selected and data used and may not be appropriate for the estimation of spatio-temporal trends in such classified regions.

A homogeneous region is defined in such a way that frequency distributions of monthly maximum hourly precipitations statistically identical throughout the region can more realistically reveal temporal changes in heavy precipitations. This will eliminate difficulties with single-site observations, such as the non-existence of long-term data, differences in observation periods, and inhomogeneous data. Heavy precipitation trend indices on such homogenous regions and thereby a regional trend indicator based on these indices will expose a more realistic picture of a study area for the existence of a significant, consistent trend. This study presents the index precipitation method to define hydrologic homogeneous regions and, therefore, to investigate temporal changes on heavy precipitation. In the example of Japan, homogenous regions were delineated so as to get similar frequency behaviors throughout each homogenous region. L-moment estimations of frequency distributions were used for this purpose, since L-moments produce unbiased frequency distributions and are less sensitive to outliers, making analyses of multi-site observations easier. Thereafter, regions with a similar frequency behavior—homogenous regions—were determined using L-moment ratios diagrams. This was achieved by cluster analysis, classifying gauging sites with the same theoretical distribution and then with similar L-moment ratios. Finally, a geographical information system analysis was implemented to differentiate noncontiguous homogenous classes, or more specifically, to get homogenous regions. Regional trends in these regions were then investigated using different precipitation indices, as well as the regional trend index.

2. Study Area and Data

Hourly precipitations from the Japan Meteorological Agency for the period 1991–2010 are used in the analyses. Records of 150 stations, for which consistent and continuous observations are available, are organized in a visual basic environment in order to obtain continuous hourly time series of precipitation records. These time series of single-site records are used in regional trend analyses in order to determine the spatio-temporal pattern of heavy precipitations throughout Japan. Figure 1 shows the locations and spatial distribution of the stations.

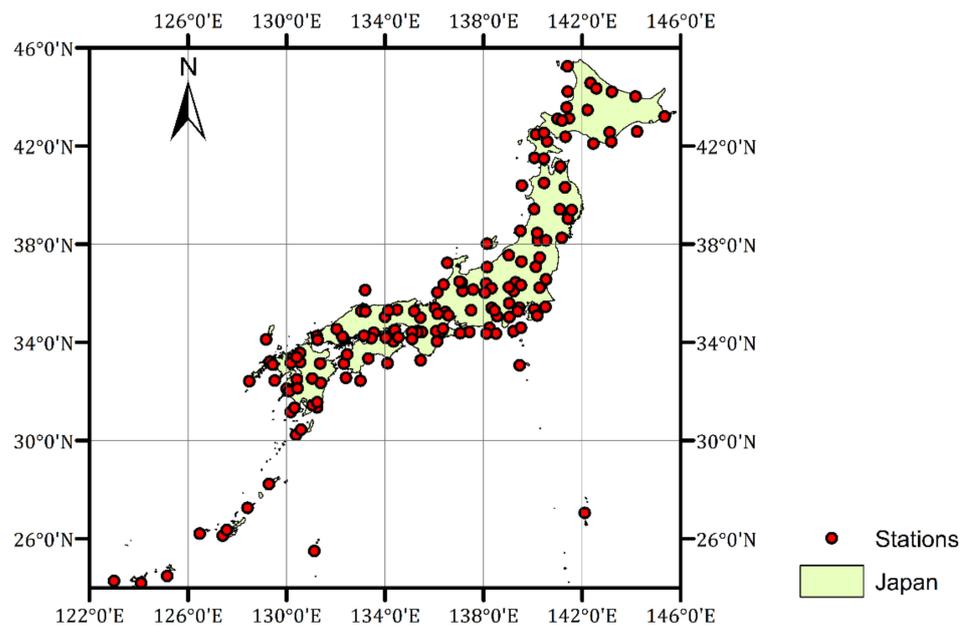


Figure 1. The locations of 150 rainfall gauges having consistent records throughout the period of 1991–2010 for the study area.

The Japan Meteorological Agency digitized precipitation data in 2000, for the full period of operation, which go back to 1900 for some sites [3]. Using annual data, the Japan Meteorological Agency [26] found a slightly decreasing trend in precipitation for the last century. There are also some researches testing long-term trends in hourly, daily, and monthly data, focusing only on the sites with a significant length of records that consist of the digital data after 1961 and the microfilm data before 1960 [3,27,28]. For example, Fujibe et al., (2005) analyzed daily, four-hourly and hourly precipitations for the period 1898–2003. They found an increase in heavy precipitation at a rate of 20–30%, throughout the century. Nonetheless, Fujibe and Yamazaki (2006) emphasize that the data of those analyses include many doubtful records and thus, are insufficient to verify a long-term trend in heavy precipitation. After checking the quality of the daily precipitation data (1901–2004) of the Japan Meteorological Agency, Fujibe and Yamazaki (2006) investigated long-term changes in the heavy precipitation in Japan, for 10 categories of the intensity, frequency and variety indices of heavy precipitation. They found that heavy precipitations have increased mainly in Eastern Japan and observed a pronounced trend generally in autumn.

3. Methodology

3.1. Regional Homogeneity Analysis

3.1.1. The Index Precipitation Method

Accurate estimation of the temporal changes in heavy precipitation may only be possible by developing a regional perspective, which reduces difficulties related to single-site data. In this context, the concept of the homogenous region has long been considered in hydrology, as an area enclosing geographically contiguous stations [29]. Such a homogenous region approach is not successful for hydrologic extremes since the geographical proximity does not warrant a common regional frequency. Although some advanced methods that combine catchment characteristics and proximity-based weights have been suggested in the literature to provide an improved definition of proximity-based homogenous regions [22], their results are generally far from obtaining a regional frequency behavior. A prominent regional approach in the literature is the index flood method. It uses frequency statistics of different sites for the delineation of the homogenous region with an identical frequency distribution [23,30]. Sveinsson et al. (2002) applied the same principle to annual

maximum precipitations of different gauging stations. Their study involves three main steps [18,31,32]. These are the identification of the sites belonging to the homogeneous region, the determination of the regional frequency distribution and the calculation of the site-specific scaling factor, which is usually taken as the mean of the observations in the considered site.

This study recommends a regionalization approach in order to define homogeneous regions for the temporal analysis of the heavy precipitation. It uses the main logic of the index flood method, defining a homogenous region with identical flood frequency distributions, but is developed for the analysis of heavy precipitation. The regionalization approach developed hereby is called The Index Precipitation Method. The method requires that a homogeneous region should satisfy the following conditions. (1) The frequency distributions of the precipitations on different sites are identical and can be represented by a regional frequency function. (2) The frequency parameters of these sites with identical frequency distributions are also identical for different sites and can be represented by their regional estimates. (3) The sites in a homogeneous region need to be geographically contiguous.

The most crucial part of the index precipitation method is the identification of appropriate frequency distributions and distribution parameters for site observations. Reviewing 12 regional frequency analyses, Cunnane (1988) suggested that probability-weighted moments provide a good regional approximation. Hosking (1990) showed that L-moments are less sensitive to the sampling variation of observations, thus, provide more realistic estimates for frequency distributions of hydrologic variables. Similarly, Vogel and Fennessey (1993) also found that product moments are not suitable for discriminating between the frequency distributions of daily stream flows. From the regional precipitation perspective, Gutmann (1993) indicated that a regional analysis based on L-moments could provide an improved probability assessment by using a denser spatial network [33]. L-moments are linear combinations of order statistics. They characterize a wide range of frequency distributions. They are more robust for outliers and give less biased estimations. They are more effective for small samples and approximates of asymptotic normal distribution in finite samples [34]. To use these advantages, The Index Precipitation Method uses L-moments in order to define both site and regional frequency distributions. For an observed series in an ascending order $x_1 \leq x_2 \leq \dots \leq x_n$, the r th L-moment of x is defined by Equation (1)

$$\lambda_{r+1} = \sum_{k=0}^r (-1)^{r-k} \binom{r}{k} \binom{r+k}{k} \beta_k, \quad r = 0, 1, 2, \dots \quad (1)$$

In particular,

$$\begin{aligned} \lambda_1 &= \beta_0 \\ \lambda_2 &= 2\beta_1 - \beta_0 \\ \lambda_3 &= 6\beta_2 - 6\beta_1 + \beta_0 \\ \lambda_4 &= 20\beta_3 - 30\beta_2 + 12\beta_1 - \beta_0 \end{aligned} \quad (2)$$

where β_r are the probability weighted moments and can be estimated by

$$\hat{\beta}_r = \frac{1}{n} \binom{n-1}{r}^{-1} \sum_{j=r+1}^n \binom{j-1}{r} x_j = \frac{1}{n} \sum_{j=1}^n \frac{(j-1)(j-2)\dots(j-r)}{(n-1)(n-2)\dots(n-r)} x_j \quad (3)$$

and the L-moment ratios can be shown in dimensionless L-moment forms as follows.

$$\begin{aligned} \tau_2 &= \lambda_2 / \lambda_1 \quad (\text{L - variation}) \\ \tau_3 &= \lambda_3 / \lambda_2 \quad (\text{L - skewness}) \\ \tau_4 &= \lambda_4 / \lambda_2 \quad (\text{L - kurtosis}) \end{aligned} \quad (4)$$

where τ_2 , τ_3 , τ_4 are alternate measures for the coefficient of variation, the coefficient of skewness and the coefficient of kurtosis, respectively.

The first step of a homogeneity analysis is to remove discordant sites and eliminate them from the regional analysis. The sites whose L-moments are significantly different than

others are identified as discordant sites. A common discordancy measure for detection of multivariate outliers is Wilk's test given by Equation (5). The sites with $D_i \geq 3$ are identified as discordant [31,34].

$$D_i = \frac{1}{3}(u_i - \bar{u})^T S (u_i - \bar{u})$$

$$S = \frac{1}{N} \sum_{i=1}^N (u_i - \bar{u})(u_i - \bar{u})^T \quad (5)$$

where;

D_i : The discordancy measure for the site i

S : the covariance matrix of L – moment ratios

$$u_i = \begin{bmatrix} \tau_{2,i} \\ \tau_{3,i} \end{bmatrix} \text{ for 2 parameter distributions}$$

$$u_i = \begin{bmatrix} \tau_{3,i} \\ \tau_{4,i} \end{bmatrix} \text{ for 3 parameter distributions}$$

\bar{u} : The mean vector of u_i

$\tau_{r,i}$: The r th dimensionless L-moment of the site i

3.1.2. Frequency Distributions and Their L-Moment Estimations

The index precipitation method suggested in this study requires first to define the region throughout which the frequency distributions of observed hydrologic variables are identical. Hosking (1990) showed that L-moments provided the best estimations for frequency parameters. L-moments characterize a wide range of frequency distributions, are more robust for outliers, are more accurate especially for small samples, give less biased estimations, approximate to an asymptotic normal distribution in finite samples [34–36]. L-moment estimations of commonly-used frequency distributions are considered here.

The frequency distributions of Normal (N), Gumbel (Gum) and Exponential (Exp); Lognormal 2 (LN2), Gamma (Gam) and Weibull (W2); Lognormal 3 (LN3), Generalized Parteo (GP), Generalized Extreme Value (GE), Pearson Type 3 (PEIII), Generalized Logistic (GL) are examined in this study. The ratio of the L-skewness coefficient to the L-kurtosis coefficient yields a single value for Normal, Gumbel and Exponential distributions. Lognormal 2 (LN2), Gamma (Gam) and Weibull 2 (W2) distributions are 2 parameter distributions; hence, they are less flexible for defining the sites with similar frequency distributions via L-moment ratios. Due to these deficiencies of the one and two parameter distributions, this study focuses only on the 3 parameter distributions such as Generalized Parteo (GP), Generalized Logistic (GL), Pearson Type 3 (PEIII), Generalized Extreme Value (GE), Lognormal 3 (LN3) distributions. Among these distributions, Generalized Parteo (GP), Generalized Logistic (GL), Pearson Type 3 (PEIII) distributions are considered for the analyses in order to discard similar distributions and keep homogeneous regions larger.

The frequency distributions and L-moments of Generalized Parteo (GP), Generalized Logistic (GL), Pearson Type 3 (PEIII) distributions are given below [34].

Generalized Parteo Distribution

$$F(x) = 1 - \left[1 - \frac{\kappa}{\beta}(x - \alpha) \right]^{\frac{1}{\kappa}}$$

$$\alpha = 0 \text{ or known} \quad (6)$$

$$\beta = \lambda_1(1 + \kappa)$$

$$\kappa = \frac{\lambda_1}{\lambda_2} - 2$$

Generalized Logistic Distribution

$$\begin{aligned}
 F(x) &= \left\{ 1 + \left(1 - \frac{\kappa}{\beta} (x - \alpha) \right)^{1/\kappa} \right\}^{-1}; \kappa \neq 0 \\
 \alpha &= \lambda_1 - \beta \left(\frac{1}{\kappa} - \frac{\pi}{\sin \kappa \pi} \right) \\
 \beta &= \frac{\lambda_2}{\kappa \pi} \sin(\kappa \pi) \\
 \kappa &= -\tau_3
 \end{aligned} \tag{7}$$

Pearson Type 3 Distribution

$$\begin{aligned}
 F(x) &= \frac{1}{\Gamma(\kappa)\beta} \int_{\alpha}^x \left(\frac{x-\alpha}{\beta} \right)^{\kappa-1} \exp\left(-\frac{x-\alpha}{\beta}\right) dx \\
 \alpha &= \lambda_1 - \beta\kappa \\
 \beta &= \sqrt{\pi}\lambda_2 \frac{\Gamma(\kappa)}{\Gamma(\kappa+0.5)} \\
 \kappa &= \begin{cases} \frac{1+0.2906z}{z+0.1882z^2+0.0442z^3}; z = 3\pi\tau_3^2; \text{ for } |\tau_3| < \frac{1}{3} \\ \frac{0.36067z-0.59567z^2+0.25361z^3}{1-2.78861z+2.56096z^2-0.77045z^3}; z = 1 - |\tau_3|; \text{ for } 1/3 \leq |\tau_3| < 1 \end{cases}
 \end{aligned} \tag{8}$$

The variables in Equations (7)–(9) are defined below.

$F(x)$: Cumulative frequency distribution.

α , β and κ : The parameters of the frequency distributions.

λ_1 and λ_2 : The first and second order L-moments

$\Gamma(\kappa)$: Gamma function

3.1.3. Delineation of Homogenous Regions

A widely accepted approach for the regional homogeneity is to delineate the stations having similar hydrologic frequencies [31]. Such an analysis needs first to know the frequency behaviors of site observations. The method of L-moments provides, at this point, a novel approach for deciding the type of the frequency distribution of a hydrologic variable for a considered site. Since there is a distinct relationship between the L-moment ratios of observations for various theoretical frequency distributions, L-moment ratio diagrams can be beneficially used to decide the type of the theoretical frequency distribution of site observations. In L-moment ratio diagrams, for 2 parameter distributions, L-variation and L-skewness coefficients and for 3 parameter distributions, L-skewness and L-kurtosis coefficients are plotted reciprocally. Based on these L-moment ratio diagrams, the theoretical frequency distribution of a site is determined by finding the best fitting frequency distribution [23].

This study recommends a 3-step identification procedure, the index precipitation method, for determination of homogeneous regions. The first step of the index precipitation method is the identification of the sites that can be statistically assumed as having an identical frequency distribution. Here, L-moment ratio diagrams are suggested to use for the determination of the theoretical frequency distribution representing site observations. A visual inspection of the distances on the L-moment ratio diagram is subjective for finding the frequency distributions best fitting to the observations, therefore Euclidean distances can be used practically and effectively. For a considered site i , the frequency distribution that gives the minimum average weighted distance (AWD in Equation (9)) can be considered as the best fitting frequency distribution for precipitation observations on that site [37,38]. This analysis is repeated until all sites are assigned into their best fitting frequency distribution classes, once discordant sites are refined by the discordancy test in Equation (5) or any outlier removing techniques in literature.

$$\begin{aligned}
 d_i &= \sqrt{(\tau_3 - \hat{\tau}_3)^2 + (\tau_4 - \hat{\tau}_4)^2} \\
 AWD &= \frac{\sum_{i=1}^N n_i d_i}{\sum_{i=1}^N n_i}
 \end{aligned} \tag{9}$$

where,

AWD: The average weighted distances

i: Index for sites

N: The number of sites

n_i : The number of the observations at the site *i*,

d_i : The distance between the L-moment ratios of the theoretical distributions and observations for the site *i*,

$\tau_{3,i}$, $\tau_{4,i}$: the 3th and 4th dimensionless L-moments of site observations

$\hat{\tau}_{3,i}$, $\hat{\tau}_{4,i}$: the 3th and 4th dimensionless L-moments of a considered theoretical distribution.

The second step of the index precipitation method is the determination of the sites that can be represented by regional parameters. For the sites that have an identical precipitation frequency distribution, determined at the first step, the Cluster Analysis provides a simple tool that can further classify the sites with an identical frequency distribution into the classes with a closely similar behavior. Accordingly, for the sites with an identical frequency distribution, the sites that have similar L-moment parameter estimations α_i , β_i , κ_i are classified into the same homogenous class. The similarity here is defined by Euclidean Distances between L-moment parameter estimations for different sites (Equation (10)). For Euclidean Distance calculations, the standardized values of the parameter estimations α_i , β_i , κ_i make possible to give equal importance to each parameter. A hierarchical classification classifies each site into a closest class, at each hierarchical step. Calculations are conducted until all sites with the identical distribution are classified into one single class. However, final homogenous classes are determined by considering only the significant number of hierarchical steps after which the distances between classes or sites increase significantly.

$$d_{i,j} = \sqrt{(\alpha_i - \alpha_j)^2 + (\beta_i - \beta_j)^2 + (\kappa_i - \kappa_j)^2} \quad (10)$$

where;

$d_{i,j}$: The Euclidean Distance between L-moment parameters of the sites *i* and *j*

α_i , β_i , κ_i : L-moment parameters of the frequency distribution of the site *i*

α_j , β_j , κ_j : L-moment parameters of the frequency distribution of the site *j*

The third and last step of the index precipitation method is the delineation of geographically contiguous sites, regarding homogenous classes determined in the steps 1 and 2. If the sites with both the identical frequency distribution and similar L-moment parameters are spatially contiguous, the area delineating these sites are defined as a homogenous region. Thiessen polygon analysis or any other interpolation techniques that consider topography or basin physiographic parameters can be used for the delineation process, conveniently via Geographical Information System –GIS– analyses. Once homogenous regions are delineated, they can be tested by any methods in the literature. Here, the regional homogeneity test of Hosking and Wallis (1993, 1997) is suggested relying on the advantages of L-moment methods (see Section 3.1.4). For tested homogenous regions, the regional parameters may be calculated by taking the weighted means (by the length of the observations, n_i) of the distribution parameters α_j , β_j , κ_j in each homogenous region.

3.1.4. Test of Homogenous Regions

Contrary to traditional single-site hydrologic analyses, a regional analysis needs multi-site observations of hydrologic quantities. For a delineated homogenous region, a regional model overcomes the difficulties came up with the data and irregular sampling interval. A fundamental assumption for the most of the regional analyses is “regional homogeneity”, which defines a region throughout which considered hydrologic properties

remain unchanged. Hosking and Wallis (1993, 1997) suggests the following test of regional homogeneity, relying on L-moment ratios.

$$\begin{aligned}
 H_r &= \frac{V_r - \bar{V}_r}{\sigma_{V_r}} ; r = 1, 2, 3 \\
 V_1 &= \frac{\sum_{i=1}^N n_i (\tau_{2,i} - \bar{\tau}_2)^2}{\sum_{i=1}^N n_i}, \\
 V_2 &= \frac{\sum_{i=1}^N n_i [(\tau_{2,i} - \bar{\tau}_2)^2 + (\tau_{3,i} - \bar{\tau}_3)^2]^{\frac{1}{2}}}{\sum_{i=1}^N n_i}, \\
 V_3 &= \frac{\sum_{i=1}^N n_i [(\tau_{3,i} - \bar{\tau}_3)^2 + (\tau_{4,i} - \bar{\tau}_4)^2]^{\frac{1}{2}}}{\sum_{i=1}^N n_i}
 \end{aligned} \tag{11}$$

where;

i : index for sites

N : The number of sites

n_i : Sample size for the site i

$\tau_{2,i}, \tau_{3,i}, \tau_{4,i}$: The 1th, 2nd, and 3rd L-moment ratios for the site i (see Equation (4))

$\bar{\tau}_2, \bar{\tau}_3, \bar{\tau}_4$: The means of $\tau_{2,i}, \tau_{3,i}, \tau_{4,i}$ in a considered homogenous region.

\bar{V}_r : The regional mean of the parameter V_r

σ_{V_r} : The regional standard deviation of the parameter V_r

H_r : Heterogeneity Measure

According to the test, if $H_r \leq 1$ then the region is acceptably homogenous; if $1 \leq H_r \leq 2$ then the region is possibly heterogeneous; if $H_r \geq 2$ then the region is definitely heterogeneous.

3.2. Trend Analyses on Homogenous Regions

3.2.1. Trend Indices

The Intergovernmental Panel on Climate Change [13] reports that climate change should be primarily addressed by the changes in the precipitation, the changes in the climate extremes, the internal consistencies of observed trends and the speed of the changes, as well as the other temperature and atmospheric/oceanographic circulation related changes [39]. Investigation of these changes suffer from the following facts; A rise in the mean is not necessarily related to a rise in the extremes; A change in the mean may cause a shift in the distribution and thus may have a major impact on the society and ecosystem [40]; Hydrologic temporal changes cannot be realistically determined solely by using single-site observations, and should be regionally consistent. These facts demand regional studies that focus on climatic extremes, and further investigations of daily/hourly data [41,42].

The IPCC Third Assessment Report also points out that the difficulties related to global data hamper the studies on climatic changes and urge to develop climate indicators [13]. Traditionally, the studies related to the changes on climatic extremes focus on single-site monthly, seasonal or annual data. However, a single-site data rarely gives a significant proof on the existence of climatic changes unless there is a long-term observation and it is supported by surrounding hydrologic observations. In other words, the classical trend analysis of single-site observations cannot accurately detect hydrologic changes on heavy precipitations. Instead, a regional analysis relying on the precipitation indices that characterize heavy precipitations provides more robust tool for investigation of temporal precipitation changes and thus climate variabilities.

This study recommends 12 indices to potentially cover all aspects of the temporal change in hourly precipitation. Once I_1, I_2, \dots, I_{12} indices are estimated for the same months of subsequent years and organized in time series for a homogenous region considered,

the existence of a significant trend on heavy precipitations can be verified by t and Mann-Kendall Tests in Section 3.2.2.

$$\begin{aligned}
 I_1 &= \frac{1}{m \cdot n} \sum_{i=1}^m \sum_{t=1}^n P_i(t) \\
 I_2 &= \max \left(\sum_{t=1}^n \frac{P_1(t)}{n}, \sum_{t=1}^n \frac{P_2(t)}{n}, \dots, \sum_{t=1}^n \frac{P_m(t)}{n} \right) \\
 I_3 &= \max(P_i(t); i = 1 \text{ to } m; t = 1 \text{ to } n) \\
 I_4 &= \frac{1}{m} \sum_{i=1}^m \max(P_i(t); t = 1 \text{ to } n) \\
 I_5 &= \frac{1}{m} \sum_{i=1}^m \frac{p_i^+}{n} \\
 I_6 &= \max \left(\frac{p_1^+}{n}, \frac{p_2^+}{n}, \dots, \frac{p_m^+}{n} \right) \\
 I_7 &= \frac{1}{m} \sum_{i=1}^m \frac{p_i^*}{n} \\
 I_8 &= \max \left(\frac{p_1^*}{n}, \frac{p_2^*}{n}, \dots, \frac{p_m^*}{n} \right) \\
 I_9 &= \frac{1}{m} \sum_{i=1}^m \frac{pv_i}{n} \\
 I_{10} &= \max \left(\frac{pv_1}{n}, \frac{pv_2}{n}, \dots, \frac{pv_m}{n} \right) \\
 I_{11} &= \frac{1}{m} \sum_{i=1}^m \frac{p_i^\#}{n} \\
 I_{12} &= \max \left(\frac{p_1^\#}{n}, \frac{p_2^\#}{n}, \dots, \frac{p_m^\#}{n} \right)
 \end{aligned} \tag{12}$$

where;

m is the number of the stations in the considered homogenous region

n is the number of the hours in the considered month

i is the index of station

t is the index of time (hours)

$P_i(t)$ is the precipitation for time t and site i (mm/hour).

p_i^+ is the number of the rainy hours

p_i^* is the number of the rainy hours for which $p_i \geq \tilde{p}_i$. Where \tilde{p}_i is the precipitation corresponding to α significance level for the frequency distribution fitted to hourly precipitations for the considered month τ .

pv_i is the precipitation of the rainy hours for $p_i \geq \tilde{p}_i$ (i.e., extreme precipitations)

$p_i^\#$ is the length of the maximum wet period in the considered month τ (hours)

I_1 : the mean hourly precipitation in the homogenous region considered

I_2 : The regional maximum of the mean hourly precipitations.

I_3 : The regional maximum of the observed hourly precipitations.

I_4 : The regional mean of the observed maximum hourly precipitation in the sites.

I_5 : The regional mean of the rainy hours in a month

I_6 : The regional maximum of rainy hours in a month

I_7 : The regional mean of the rainy hours (in a month) with extreme precipitation

I_8 : The regional maximum of the rainy hours (in a month) with extreme precipitation

I_9 : The regional mean of the extreme precipitations in a month

I_{10} : The regional maximum of the extreme precipitations in a month

I_{11} : The regional mean of the length of maximum wet period for a considered month τ

I_{12} : The regional maximum of the length of maximum wet period for a considered month τ

The regional trend index (RTI) defined by Equation (13) is used to verify regional trends on the heavy precipitation. The regional trend index indicates the percentage of I_1, I_2, \dots, I_{12} indices for which there are statistically significant trends.

$$RTI = \frac{1}{12} \sum_{j=1}^{12} TI \quad (13)$$

RTI: The regional trend index

TI: The number of the I_1, I_2, \dots, I_{12} indices with a significant trend.

3.2.2. Trend Tests

Trend tests determine if a random variable changes over the time. They are valuable especially for estimation of possible effects of the climate change [43]. The statistical significance of a calculated trend can be decided by parametric or non-parametric tests. While parametric tests assume that the random variable is normally distributed, non-parametric tests rely on order statistics and do not take the frequency distribution into account. t -test (parametric) and Mann Kendall test (non-parametric) are powerful tests and can be interchangeably used in the practice. t -test provides better results than Mann Kendall tests when the frequency distribution is normal, however Mann Kendall test is powerful when the frequency distribution is skewed [43]. Both t test and Mann-Kendall test are used in this study in order to test the existence of significant trends in precipitation indices and thus on heavy precipitations throughout homogenous regions determined.

The t -test with the $(n - 2)$ degree of freedom tests the hypothesis $H_0 : \hat{r} = 0$ vs. $H_1 : \hat{r} \neq 0$ to verify if a statistically significant trend exists for a confidence level α [44]. The H_0 hypothesis is rejected when $t < t_{cr,\alpha/2}$.

$$t = r \frac{\sqrt{n-2}}{\sqrt{1-r^2}} \quad (14)$$

where;

n is the sample size

r is the coefficient of correlation

$t_{cr,\alpha/2}$ is the critical t value for the confidence level $\alpha/2$.

Mann-Kendall Test uses order statistics of observations. For a precipitation series $p(1), p(2), \dots, p(n)$, Mann-Kendall statistic is given by Equation (15).

$$S = \sum_{k=1}^{n-1} \sum_{j=k+1}^n \text{sign}[p(j) - p(k)] \quad (15)$$

$$\text{sign}[p(j) - p(k)] = \begin{cases} 1 & \text{if } p(j) - p(k) > 0 \\ 0 & \text{if } p(j) - p(k) = 0 \\ -1 & \text{if } p(j) - p(k) < 0 \end{cases}$$

It is known that the z statistic given by Equation (16) follows a standard normal distribution. Therefore, the z statistic can be used for testing the null hypothesis of "there is no trend". The null hypothesis $H_0 : S = 0$ vs. $H_1 : S \neq 0$ is rejected, if $z > z_{\alpha/2}$ for the confidence level $\alpha/2$.

$$z = \begin{cases} \frac{(S-1)}{\sigma_s} & \text{for } S > 0 \\ 0 & \text{for } S = 0 \\ \frac{(S+1)}{\sigma_s} & \text{for } S < 0 \end{cases} \quad (16)$$

$$\sigma_s = \sqrt{\frac{n(n-1)(2n+5) - \sum_{k=1}^n t_k(t_k-1)(2t_k+5)}{18}}$$

where;

t_k is the number of the tied observations (a set of sample data with same value) in the k th group

g is the number of the tied groups

σ_s is the standard deviation of S statistic

4. Application and Results

The hourly precipitations of 150 gauging sites from Japan Meteorological Agency are organized in time series for the period 1991–2010. To have a base for the regional trend analyses and to see how calculated trends on the single sites are regionally consistent, an initial trend analysis is implemented for monthly maximum precipitations—i.e., maximum hourly precipitations in subsequent months.

Both the t -test and Mann Kendall tests are used to see if there are significant trends on monthly maximum precipitations. Test results are given in Table 1 for the sites with statistically significant trends. The results show that monthly maximum precipitations may have trends for the sites 575, 585, 616, 622, 836, 940. Out of 150 sites, only the sites 585, 616, 940 have passed both of the trend tests.

Table 1. Resultant statistics of the trend analyses of the maximum hourly precipitations on the observation sites for 1991–2010 (the t -test and Mann-Kendall test are used to test the existences of significant trends for 0.95 confidence level).

Trends in Sites						
Sites	575	585	616	622	836	940
b (mm/month)	0.011	0.023	0.020	0.020	0.040	0.035
t values	*	2.266	2.090	2.370	1.961	2.464
z values	1.989	2.093	1.537	*	*	1.963

* there is no statistically significant trend for this site, based on the test used; b : regression coefficient; t : t statistics (see Equation (14)); z : z statistic (see Equation (16)).

Another prior analysis on single-site observations is implemented directly for hourly precipitation data. Contrary to the monthly maximum precipitations, the hourly precipitations of many sites appear to have significant trends, though the t -test and Mann Kendall tests give slightly different results. Figure 2 shows statistically significant positive and negative trends in hourly precipitations of single-site observations through Japan. The sites in high and low latitude regions have mainly negative trends, while the sites of middle latitude inner regions have positive trends.

For the sites 402 and 575, which give respectively the highest negative and positive t scores, the considered lengths of hourly precipitation series are changed systematically (by 3 months) to see if t -scores change by the length of the series considered (Figure 3a). Moreover, in order to see the effect of different observation periods on t -scores, the time spans of the observations are changed systematically so as to have subseries on different time periods of whole length of the hourly series. For 26,280 h, 52,560 h and 78,840 h (i.e., 3, 6 and 9 years) series lengths, the subseries are so created that the time inceptions are systematically increased by 8640 h (approximately 1 year). Figure 3b–d shows the change of t -scores by these subseries of different observation periods, respectively for 26,280 h, 52,560 h and 78,840 h length series. This analysis indicated that the t -score therefore the acceptance of a trend hypothesis highly depends on the length of series. The lengthier series, more accurate the estimations on trend analyses. The considered periods of the observations are also vital to decide to the existence of a statistically significant trend, especially for small size series. For larger series, the difference between the t -statistic and critical t -values is getting larger, this helps us to clarify if there is a statistically significant trend.

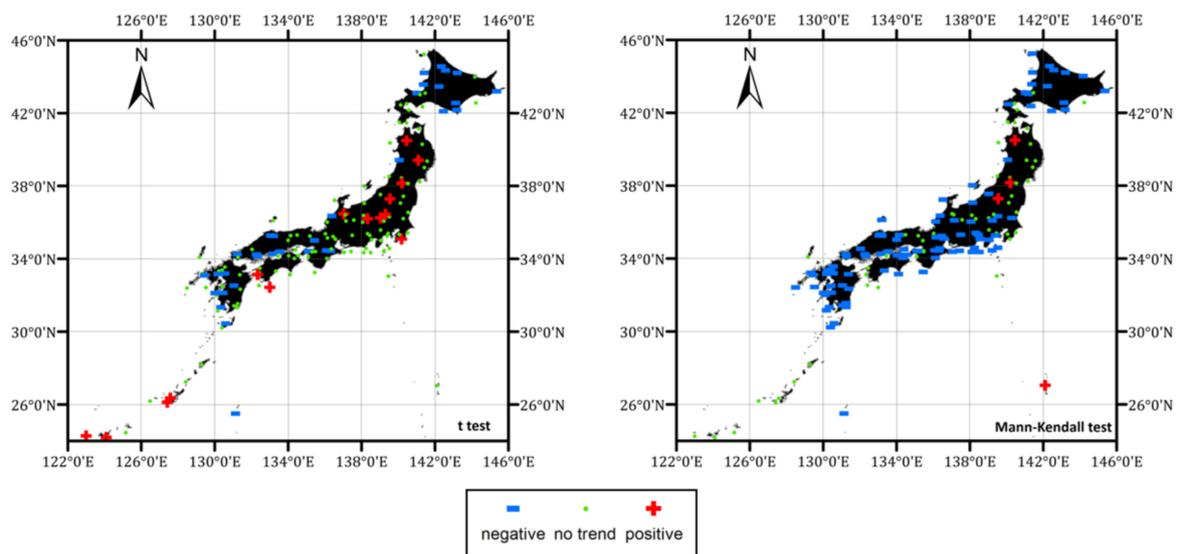


Figure 2. Trends in hourly precipitations for 1991–2010 (trends approved by both the t -test and Mann-Kendall test for 0.95 confidence level).

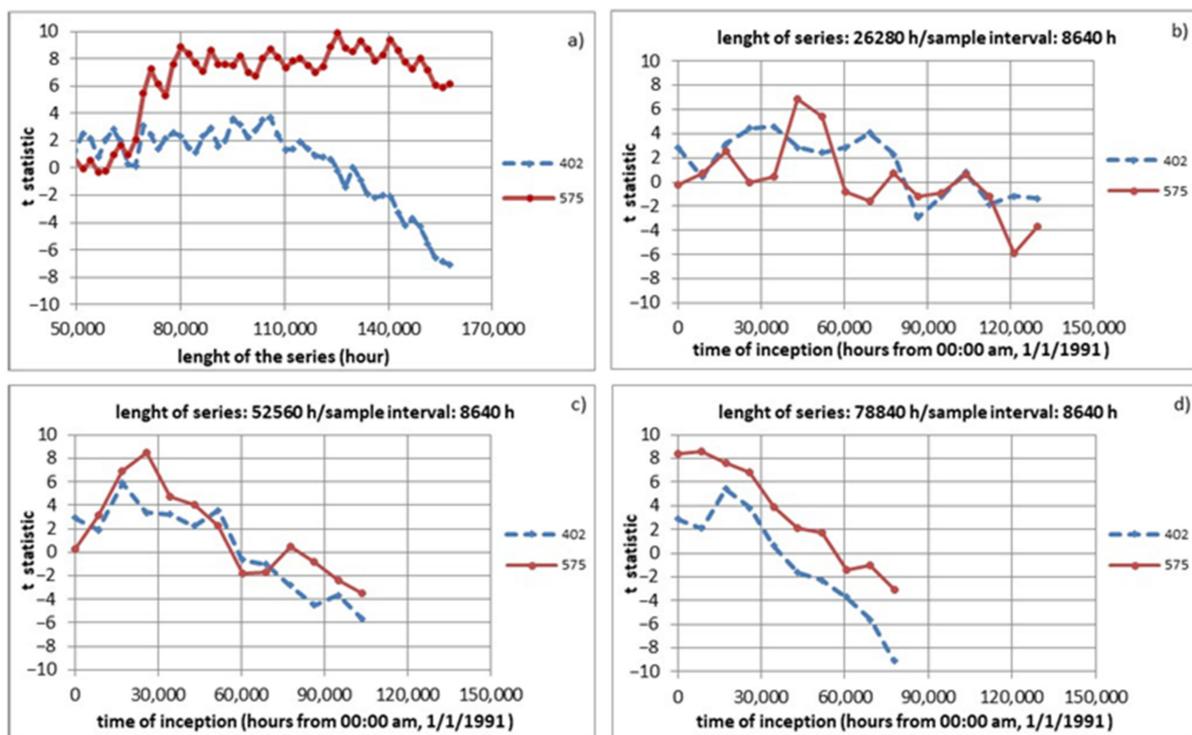


Figure 3. Sampling behaviors of the t -statistic for hourly precipitations with different length of series. (a) shows the change of the t -scores by the length of series. (b–d) shows the effect of the different observation periods on the t -scores, corresponding to the different time spans of the observations, shifted systematically by 8640 h, for 26,280 h, 52,560 h and 78,840 h length series, respectively.

For single-site observations, the existence of trend is also investigated for monthly maximum precipitations of the same months of subsequent years, regarding that climatic extremes are more sensitive to climatic changes. The monthly maximum precipitation series are organized into 12 new time series, representing months. Both the t -test and Mann Kendall tests are used to test these data series for possible trends. The sites with significant trends are given in Figure 4. According to Figure 4, some sites have significant trends for

some months, nonetheless no general pattern can be seen and proves a consistent trend supported by multi stations, geographically contiguous.

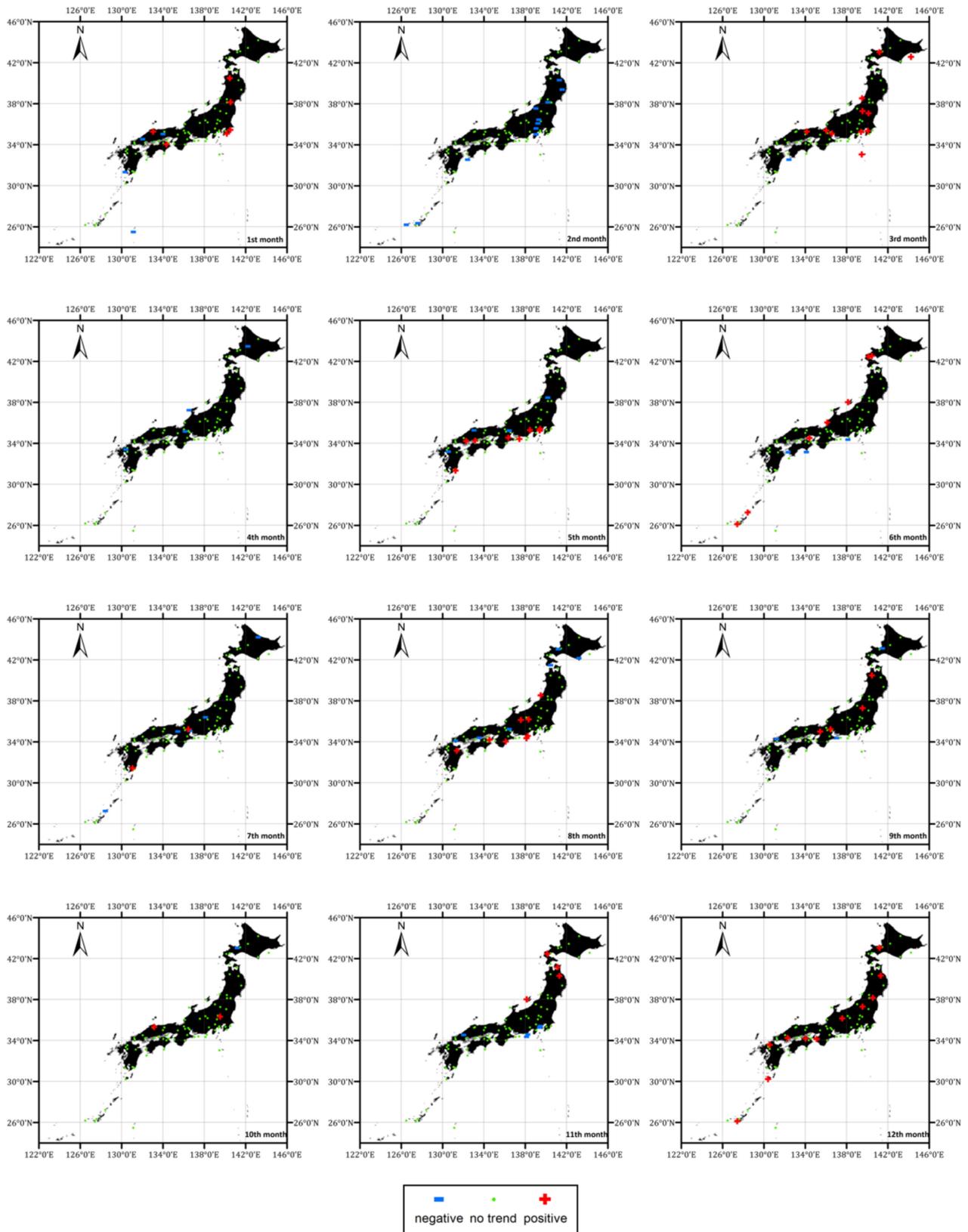


Figure 4. Trends in monthly maximum precipitations for the same months along 1991–2010 (Trends approved by both the *t*-test and Mann Kendall test for 0.95 confidence level).

These prior analyses confirmed that trend analyses are highly sensitive to the length of series, observation period and the type of the test applied. The calculated trends lack geographical consistency and generally consider only a single statistic (e.g., maximum, mean etc.) of heavy precipitation. These facts demand regional studies that focus on climatic extremes, and further investigations of hourly data [3,41]. The need for further investigation on multi-station trend analysis has created the main motivation for this study. Aiming to develop a regional approach for investigation of a regional trend on heavy precipitations, this study introduces “the index precipitation method”, which defines a hydrologic homogeneous region with an identical frequency distribution of multi-site observations in order to verify the existence of a regional trend on heavy precipitations via various precipitation indices.

The index precipitation method needs a preliminary effort to refine discordant gauging sites. Two methods are applied to decide on discordant stations. The first one is Wilk’s test in Equation (5), for which the sites with $D_i \geq 3$ are identified as discordant. The second one is Euclidean Distances between L-moment ratios of observations and theoretical distributions. This method considers the sites with comparably high distances as outlier sites. For a specific month, the breaking point on a graph of the sites vs. relevant distances (in increasing order) give the critical distance for the outliers. Calculated discordant (outlier) sites are given in Table 2.

Table 2. Discordancy Analyses for the period of 1991–2010 (The discordant stations are the sites for which Discordancy Measure D_i of Wilks’s test are ≥ 3 or for which Euclidean Distances d_{min} between the L-moment ratios of the observations and theoretical distributions are—considerably high compared to the rest of the data set—larger than the critical distance for that month).

1th Month				2th Month				3th Month				4th Month			
Station	d_{min}	Station	D												
622	0.167	784	5.525	666	0.182	413	7.802	401	0.184	401	9.586	612	0.149	684	4.342
784	0.166	512	4.671	653	0.171	636	6.138	640	0.143	740	7.453	769	0.127	675	4.272
672	0.154	585	4.425	742	0.150	776	5.695	656	0.119	430	6.164	670	0.123	651	4.216
817	0.153	898	4.293	424	0.118	587	4.872	942	0.118	626	5.252	622	0.116	605	3.635
780	0.134	762	4.212	819	0.116	653	4.190	750	0.104	741	4.343	651	0.104	893	3.605
766	0.131	440	4.175	570	0.111	895	3.961	754	0.103	899	4.334	893	0.095	662	3.498
649	0.105	672	3.767	631	0.110	666	3.942	831	0.096	784	4.043	768	0.091		
626	0.104	815	3.566	822	0.110	837	3.840	610	0.094	656	3.902	755	0.085		
616	0.101			421	0.097	742	3.146	597	0.077			662	0.068		
674	0.095			636	0.077			612	0.061			657	0.062		
618	0.082							585	0.060			767	0.060		
654	0.077							592	0.052			830	0.055		
												421	0.053		
												606	0.052		
												618	0.040		
												674	0.039		
												570	0.037		

5th Month				6th Month				7th Month				8th Month			
Station	d_{min}	Station	D												
837	0.126	777	6.064	754	0.133	824	8.727	807	0.138	575	6.005	632	0.208	632	9.369
629	0.111	780	5.651	824	0.114	624	6.802	674	0.120	812	5.655	822	0.122	684	8.802
624	0.094	626	4.335	822	0.112	649	5.819	819	0.116	776	4.675	587	0.105	440	6.681
835	0.092	837	3.826	684	0.104	684	4.473	605	0.106	945	4.471	655	0.103	746	4.936
607	0.092	663	3.741	682	0.098	602	3.691	405	0.091	637	3.511	674	0.090	598	4.677
821	0.091	649	3.701	648	0.096	835	3.487	597	0.090	426	3.447	768	0.079	761	3.491
407	0.090	618	3.141	590	0.090	770	3.298	892	0.089	407	3.243	684	0.062	822	3.125
772	0.078			636	0.084	655	3.287	762	0.088	819	3.232			762	3.064
756	0.076			778	0.083			893	0.079	807	3.051			411	3.047
893	0.076			428	0.082			581	0.077						
570	0.074			823	0.081			616	0.074						
777	0.073			805	0.080			668	0.063						

Table 2. Cont.

5th Month				6th Month				7th Month				8th Month			
Station	d_{\min}	Station	D	Station	d_{\min}	Station	D	Station	d_{\min}	Station	D	Station	d_{\min}	Station	D
411	0.073			831	0.080			595	0.062						
672	0.071			423	0.076			404	0.054						
512	0.070			654	0.070			656	0.053						
831	0.070			581	0.070										
598	0.068			582	0.066										
648	0.063														
663	0.060														
940	0.058														
636	0.040														
9th Month				10th Month				11th Month				12th Month			
Station	d_{\min}	Station	D	Station	d_{\min}	Station	D	Station	d_{\min}	Station	D	Station	d_{\min}	Station	D
574	0.257	574	6.647	632	0.187	620	6.130	592	0.159	675	5.012	570	0.326	404	9.117
830	0.211	584	4.482	620	0.165	604	5.897	624	0.154	626	4.679	588	0.212	750	7.069
822	0.167	830	4.453	784	0.162	435	4.899	807	0.153	807	4.344	584	0.185	570	6.661
584	0.148	822	4.163	759	0.148	800	4.584	836	0.149	575	4.031	587	0.167	770	5.135
610	0.115	632	3.937	815	0.132	767	3.817	424	0.145	592	4.009	428	0.163	421	4.678
423	0.114	654	3.583	420	0.116	617	3.649	616	0.130	424	3.982	843	0.134	892	4.468
766	0.112	440	3.363	624	0.111	815	3.418	626	0.126	651	3.837	402	0.132	590	4.295
626	0.110	653	3.150	607	0.106	632	3.342	822	0.100	836	3.825	756	0.130	585	4.172
817	0.105			598	0.094	612	3.293	823	0.088	588	3.208	404	0.128	581	4.048
426	0.105			433	0.085	759	3.277	769	0.080			838	0.109	741	3.914
648	0.105			829	0.081	420	3.201	675	0.075			835	0.083	588	3.096
512	0.093			604	0.068							898	0.056	417	3.049
741	0.093			893	0.057							590	0.045		
827	0.076														
682	0.070														

The index precipitation method requires a three-step verification for delineation of a homogeneous region. As a first step, in order to define the sites with identical frequency distributions of maximum hourly precipitations, L-moment ratios of the site observations and theoretical distributions were calculated. Based on the L-moment ratio diagram, distances between the L-moment ratios of the observations and theoretical distributions were calculated. The frequency distributions best fitting to hourly precipitations were then investigated by means of the AWD measure in Equation (9). The L-moment ratios—L skewness/L kurtosis—of the 150 sites and the L-moment ratios of Generalized Pareto (GP), Generalized Logistic (GL) and Pearson Type 3 (PEIII) distributions are plotted in Figure 5. Both the discordant sites and the sites with identical frequency distributions are shown in the Figure 5.

The second step of the index precipitation method requires to determine the sites with identical regional parameters, for the same frequency distribution sites. Hierarchical Clustering was applied to further classify the sites with an identical frequency distribution into the sites with similar frequency distribution parameters. Euclidean Distances between the frequency parameters α_i , β_i , κ_i were used to define the similarities for each month via Equation (10).

The significant numbers of hierarchical levels were decided based on the graphs of Euclidean distances vs. hierarchical steps, by finding the step after which the graph begins to level off. The obtained homogenous classes through which α_i , β_i , κ_i are statistically identical are given in Table 3. Where, n is the number of the sites in homogenous classes.

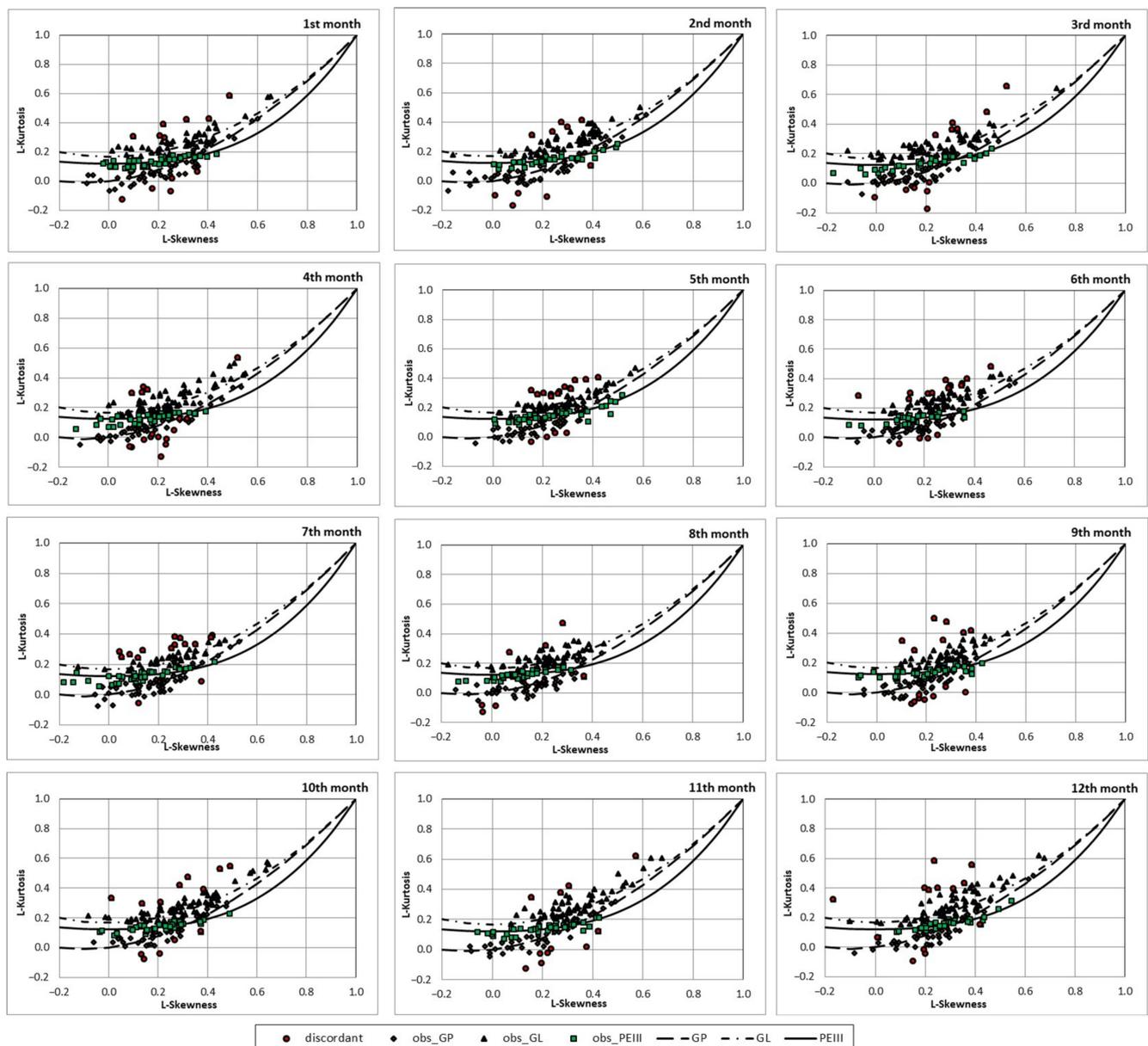


Figure 5. L-moment ratio diagrams (Graphs show fitness of the L-moment ratios of the 150 sites and different theoretical distributions such as Generalized Parteo GP, Generalized Logistic GL and Pearson Type 3 PEIII distributions).

The third step of the index precipitation method was implemented by delineating homogenous classes into geographically contiguous homogeneous regions. Thiessen polygon method was used for this purpose. Each of the calculated homogenous regions therefore provided the assumptions of (1) identical frequency distribution, (2) identical frequency distribution parameters, (3) geographically continuity. The homogenous regions are shown in Figure 6.

The regional parameters obtained for the homogenous regions are given in Table 4.

Defining hydrologic characteristics of heavy precipitations, the 12 indices given in Equation (11) were used to verify the existence and consistency of the regional trends on the homogenous regions. For each homogenous region, the estimates of 12 indices were organized in time series of same months. Regional trends were tested for the homogenous regions by the *t*-test and Mann Kendall tests. The *t* and *z* values are given in Table 5.

Table 3. Statistics of the probabilistic distribution parameters in homogeneous classes (Calculated by Equations (6)–(8)).

Month	Class	Generalized Pareto							Generalized Logistic							Pearson Type III						
		n	Mean			Std Dev			n	Mean			Std Dev			n	Mean			Std Dev		
			α	k	β	α	k	β		α	k	β	α	k	β		α	k	β			
1	1	12	0	1.565	2.565	0	0.242	0.242	23	0.932	−0.170	0.221	0.040	0.075	0.073	8	0.241	2.156	0.365	0.155	0.412	0.105
	2	14	0	2.387	3.387	0	0.198	0.198	19	0.835	−0.362	0.240	0.041	0.041	0.066	8	0.406	1.005	0.628	0.189	0.293	0.246
	3	11	0	0.939	1.939	0	0.122	0.122	3	0.595	−0.615	0.233	0.081	0.056	0.045	3	−0.324	10.861	0.122	0.347	0.338	0.030
	4	13	0	0.363	1.363	0	0.218	0.218								4	0.161	4.090	0.206	0.262	0.660	0.062
	5	2	0	5.412	6.412	0	0.055	0.055								2	0.122	8.054	0.109	0.103	0.484	0.006
	6															2	0.055	5.668	0.169	0.476	0.287	0.093
2	1	14	0	1.443	2.443	0	0.176	0.176	22	0.909	−0.230	0.226	0.029	0.049	0.051	23	0.172	1.949	0.620	0.197	1.128	0.428
	2	19	0	0.713	1.713	0	0.259	0.259	24	0.816	−0.383	0.243	0.042	0.024	0.050	3	−0.036	7.843	0.131	0.299	1.430	0.029
	3	7	0	2.182	3.182	0	0.221	0.221	3	0.976	−0.124	0.116	0.008	0.031	0.021							
	4	2	0	4.689	5.689	0	0.026	0.026	3	0.979	−0.052	0.234	0.014	0.021	0.062							
	5	4	0	2.830	3.830	0	0.104	0.104	2	0.812	−0.526	0.150	0.028	0.069	0.014							
	6	4	0	3.774	4.774	0	0.249	0.249														
	7	2	0	5.688	6.688	0	0.354	0.354														
3	1	21	0	2.442	3.442	0	0.340	0.340	20	0.837	−0.370	0.225	0.030	0.058	0.027	16	0.435	1.309	0.490	0.156	0.626	0.177
	2	17	0	1.314	2.314	0	0.320	0.320	31	0.941	−0.180	0.199	0.020	0.064	0.047	4	0.121	3.650	0.240	0.135	0.200	0.031
	3	11	0	3.849	4.849	0	0.469	0.469	7	1.003	0.009	0.171	0.015	0.049	0.027	2	−0.759	26.178	0.067	0.639	0.412	0.023
	4															2	−0.593	14.802	0.108	0.000	1.632	0.012
4	1	17	0	1.751	2.751	0	0.138	0.138	27	0.879	−0.263	0.251	0.039	0.053	0.035	5	−0.124	7.009	0.161	0.351	0.297	0.052
	2	13	0	1.037	2.037	0	0.159	0.159	3	0.828	−0.467	0.169	0.020	0.039	0.003	25	0.251	2.368	0.382	0.141	1.005	0.188
	3	5	0	2.130	3.130	0	0.066	0.066	16	0.961	−0.099	0.232	0.022	0.049	0.033	2	−0.798	14.358	0.127	0.541	0.845	0.045
	4	4	0	3.083	4.083	0	0.037	0.037	6	0.713	−0.486	0.262	0.045	0.055	0.023							
	5	2	0	0.633	1.633	0	0.097	0.097														
	6	2	0	2.713	3.713	0	0.098	0.098														
5	1	7	0	2.863	3.863	0	0.031	0.031	6	0.911	−0.290	0.168	0.011	0.018	0.010	10	0.290	1.755	0.412	0.162	0.404	0.081
	2	6	0	2.473	3.473	0	0.076	0.076	22	0.898	−0.234	0.246	0.025	0.040	0.034	5	0.307	2.831	0.245	0.102	0.230	0.029
	3	15	0	1.216	2.216	0	0.145	0.145	7	0.958	−0.126	0.203	0.010	0.032	0.029	8	0.440	0.578	1.037	0.118	0.172	0.345
	4	10	0	1.823	2.823	0	0.205	0.205	13	0.820	−0.414	0.214	0.025	0.039	0.030	4	0.215	3.724	0.212	0.079	0.199	0.029
	5															2	−0.057	11.899	0.089	0.109	0.108	0.010
	6															2	−0.147	9.931	0.115	0.116	0.124	0.010
	7															2	0.026	5.204	0.188	0.261	0.120	0.055
	8															2	−0.124	6.974	0.162	0.275	0.361	0.048

Table 3. Cont.

Month	Class	Generalized Pareto							Generalized Logistic							Pearson Type III						
		n	Mean			Std Dev			n	Mean			Std Dev			n	Mean			Std Dev		
			α	k	β	α	k	β		α	k	β	α	k	β		α	k	β			
6	1	16	0	1.576	2.576	0	0.125	0.125	17	0.908	−0.205	0.258	0.019	0.033	0.028	7	0.269	1.596	0.496	0.112	0.526	0.145
	2	5	0	1.996	2.996	0	0.061	0.061	19	0.843	−0.323	0.258	0.032	0.039	0.030	6	−0.269	6.235	0.206	0.138	0.760	0.030
	3	17	0	1.083	2.083	0	0.164	0.164	5	0.917	−0.262	0.178	0.007	0.018	0.010	4	−0.302	9.193	0.143	0.153	0.762	0.024
	4	6	0	2.918	3.918	0	0.112	0.112	5	0.971	−0.080	0.229	0.009	0.035	0.044	7	0.173	3.678	0.224	0.209	0.651	0.044
	5	5	0	2.407	3.407	0	0.138	0.138	4	0.762	−0.460	0.242	0.015	0.037	0.037	3	−0.605	12.560	0.130	0.269	1.032	0.033
	6	2	0	3.331	4.331	0	0.044	0.044	2	0.846	−0.501	0.133	0.030	0.053	0.001							
	7	2	0	0.687	1.687	0	0.053	0.053														
7	1	19	0	1.644	2.644	0	0.157	0.157	15	0.931	−0.155	0.264	0.024	0.054	0.039	15	0.188	2.114	0.442	0.296	1.053	0.184
	2	10	0	0.615	1.615	0	0.128	0.128	15	0.854	−0.288	0.282	0.020	0.031	0.043	5	−0.722	9.731	0.178	0.132	0.794	0.024
	3	15	0	1.107	2.107	0	0.147	0.147	4	0.748	−0.427	0.283	0.039	0.037	0.027	9	−0.107	5.722	0.194	0.206	0.562	0.038
	4	4	0	2.130	3.130	0	0.053	0.053	3	0.868	−0.395	0.165	0.024	0.051	0.001	2	−0.337	14.275	0.093	0.391	0.694	0.023
	5	2	0	3.410	4.410	0	0.097	0.097														
	6	3	0	2.736	3.736	0	0.207	0.207														
8	1	18	0	1.794	2.794	0	0.170	0.170	15	0.829	−0.325	0.278	0.039	0.033	0.042	18	0.015	3.390	0.381	0.180	2.027	0.178
	2	10	0	1.267	2.267	0	0.089	0.089	25	0.897	−0.215	0.278	0.021	0.026	0.048	7	−0.912	13.642	0.142	0.442	2.953	0.028
	3	10	0	0.837	1.837	0	0.130	0.130	8	0.943	−0.121	0.280	0.013	0.017	0.034	2	−2.109	51.334	0.061	1.151	1.310	0.024
	4	7	0	0.391	1.391	0	0.079	0.079	5	0.998	−0.006	0.255	0.016	0.041	0.049							
	5	5	0	2.373	3.373	0	0.165	0.165														
9	1	16	0	1.366	2.366	0	0.074	0.074	33	0.855	−0.299	0.265	0.028	0.049	0.032	13	0.349	1.132	0.620	0.117	0.334	0.214
	2	2	0	2.618	3.618	0	0.001	0.001	4	0.772	−0.462	0.226	0.040	0.044	0.010	7	−0.101	4.600	0.239	0.199	0.335	0.039
	3	12	0	1.707	2.707	0	0.124	0.124	8	0.939	−0.144	0.257	0.014	0.041	0.041	4	−0.861	16.101	0.116	0.154	0.378	0.008
	4	5	0	0.522	1.522	0	0.132	0.132								6	0.194	2.355	0.346	0.147	0.461	0.049
	5	8	0	1.003	2.003	0	0.103	0.103														
	6	5	0	2.192	3.192	0	0.116	0.116														
	7	4	0	3.049	4.049	0	0.072	0.072														
10	1	18	0	1.180	2.180	0	0.258	0.258	24	0.862	−0.261	0.297	0.020	0.041	0.031	11	0.243	1.597	0.480	0.146	0.258	0.095
	2	13	0	0.430	1.430	0	0.214	0.214	20	0.782	−0.395	0.274	0.039	0.034	0.032	5	−0.160	4.913	0.237	0.088	0.304	0.027
	3	5	0	2.197	3.197	0	0.302	0.302	3	1.009	0.033	0.208	0.011	0.045	0.046	3	−0.110	3.442	0.324	0.039	0.363	0.027
	4								12	0.943	−0.125	0.270	0.015	0.029	0.035	4	0.310	0.706	1.039	0.114	0.174	0.354
	5								4	0.628	−0.628	0.206	0.047	0.029	0.048	2	−0.001	2.376	0.431	0.122	0.374	0.119
	6								2	0.710	−0.543	0.218	0.024	0.043	0.016							

Table 4. Cont.

Month	Class	Generalized Pareto									Generalized Logistic						Pearson Type III					
		n	Mean			Std Dev			n	Mean			Std Dev			n	Mean			Std Dev		
			α	k	β	α	k	β		α	k	β	α	k	β		α	k	β			
3	1	4	0	2.302	3.302	0	0.203	0.203	3	0.806	-0.424	0.221	0.020	0.045	0.016	3	0.384	1.301	0.614	0.091	0.729	0.379
	2	5	0	2.508	3.508	0	0.398	0.398	3	0.827	-0.371	0.237	0.026	0.029	0.013							
	3	5	0	1.214	2.214	0	0.134	0.134	3	0.940	-0.141	0.253	0.032	0.079	0.029							
	4	4	0	4.115	5.115	0	0.301	0.301	6	0.946	-0.194	0.164	0.020	0.079	0.027							
	5								6	0.933	-0.196	0.198	0.021	0.049	0.048							
	6								5	0.946	-0.154	0.212	0.014	0.055	0.033							
4	1	5	0	1.701	2.701	0	0.190	0.190	15	0.871	-0.274	0.256	0.037	0.047	0.042	4	0.205	1.703	0.519	0.205	0.792	0.154
	2	3	0	1.808	2.808	0	0.064	0.064	3	0.828	-0.467	0.169	0.020	0.039	0.003	7	0.306	1.831	0.500	0.090	1.243	0.251
	3	5	0	1.033	2.033	0	0.168	0.168	4	0.956	-0.091	0.258	0.040	0.077	0.049							
	4								3	0.956	-0.109	0.238	0.017	0.028	0.037							
5	1	4	0	1.254	2.254	0	0.197	0.197	3	0.914	-0.287	0.164	0.015	0.027	0.010							
	2	3	0	1.212	2.212	0	0.082	0.082	3	0.965	-0.100	0.211	0.010	0.028	0.009							
	3								3	0.835	-0.402	0.204	0.021	0.017	0.028							
6	1	3	0.084	0.802	1.834	0.146	0.872	0.817	4	1.564	1.447	0.149	1.646	1.964	0.101							
	2	3	0.086	0.527	1.575	0.105	0.760	1.272	3	1.252	0.223	0.178	0.493	0.518	0.156							
	3								3	0.916	-0.184	0.249	0.063	0.129	0.018							
7	1	8	0	1.682	2.682	0	0.149	0.149								4	0.151	2.318	0.406	0.339	1.275	0.138
	2	3	0	1.170	2.170	0	0.097	0.097								3	0.289	1.308	0.653	0.140	0.793	0.284
8	1	4	0	1.912	2.912	0	0.032	0.032	4	0.830	-0.341	0.263	0.021	0.011	0.036	3	-0.039	3.589	0.409	0.236	2.192	0.293
	2								3	0.898	-0.200	0.297	0.017	0.025	0.032	3	0.080	3.055	0.413	0.231	2.580	0.223
	3								3	0.885	-0.234	0.280	0.015	0.016	0.037							
9	1	5	0	1.382	2.382	0	0.080	0.080	4	0.848	-0.305	0.272	0.016	0.038	0.030							
	2								14	0.848	-0.312	0.263	0.030	0.044	0.037							
	3								3	0.835	-0.345	0.249	0.029	0.044	0.004							

Table 4. Cont.

Month	Class	Generalized Pareto						Generalized Logistic						Pearson Type III								
		n	Mean			Std Dev			n	Mean			Std Dev			n	Mean			Std Dev		
			α	k	β	α	k	β		α	k	β	α	k	β		α	k	β			
10	1	4	0	0.973	1.973	0	0.166	0.166	6	0.865	-0.254	0.298	0.025	0.044	0.028	5	0.186	1.020	0.848	0.081	0.290	0.231
	2	3	0	1.582	2.582	0	0.146	0.146	4	0.853	-0.280	0.290	0.008	0.012	0.019							
	3	5	0	0.452	1.452	0	0.167	0.167	3	0.870	-0.268	0.271	0.030	0.060	0.027							
	4								4	0.862	-0.241	0.327	0.021	0.051	0.035							
	5								4	0.797	-0.405	0.247	0.043	0.019	0.044							
	6								3	0.947	-0.114	0.273	0.017	0.026	0.022							
11	1	3	0	0.486	1.486	0	0.190	0.190	6	0.861	-0.262	0.295	0.029	0.042	0.029	4	0.172	1.148	0.806	0.082	0.518	0.252
	2							7	0.866	-0.262	0.286	0.023	0.033	0.036								
12	1	3	0	0.301	1.301	0	0.214	0.214	11	0.802	-0.340	0.302	0.053	0.052	0.050	4	0.172	1.148	0.806	0.082	0.518	0.252
	2	8	0	0.315	1.315	0	0.142	0.142	6	0.835	-0.290	0.314	0.014	0.023	0.021							
	3								4	0.770	-0.357	0.332	0.054	0.064	0.023							
	4								4	0.761	-0.314	0.413	0.013	0.034	0.035							
	5								3	0.921	-0.153	0.305	0.023	0.032	0.039							
	6								4	0.929	-0.228	0.182	0.007	0.045	0.032							

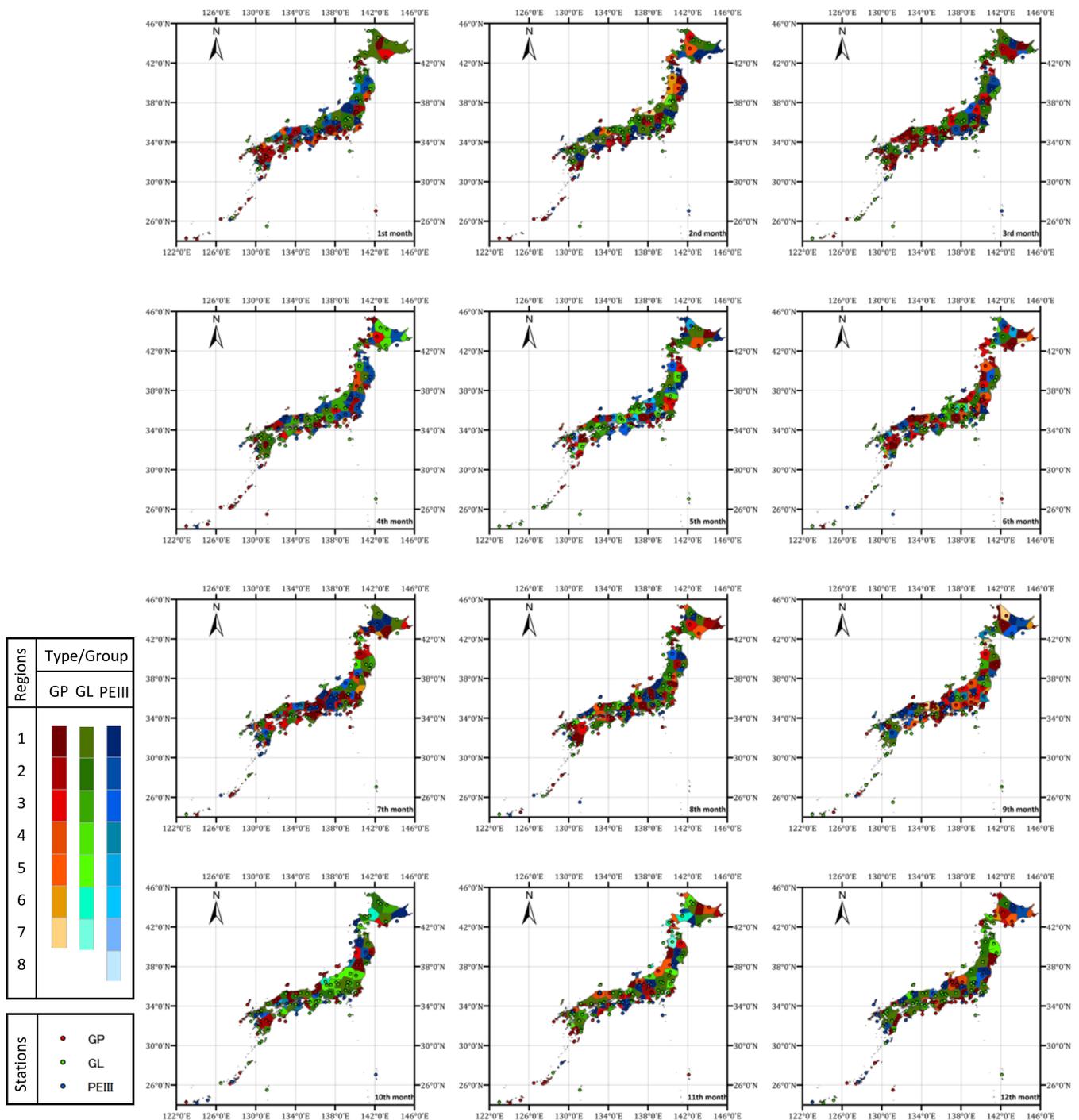


Figure 6. Homogenous regions (They are defined in three steps; the first step is determination of the sites with identical frequency distributions such as Generalized Parteo (GP), Generalized Logistic (GL) and Pearson Type 3 (PEIII) distributions; the second step is determination of the sites with similar frequency parameters such as α_i , β_i , κ_i , among the sites with an identical frequency distribution; the last step is delineation of these sites with an identical distribution and similarly equal frequency parameters into geographically contiguous regions).

Figure 7 shows homogenous regions with significant trends, regarding the regional trend index.

Table 5. The *t* and *z* statistics for different indices (Bold values show statistically significant trends in the relevant indices for the considered homogenous regions and month. The regional trend index RTI indicates the percentage of the indices with significant trends, for a homogenous region and month).

Months	Polygon	I ₁		I ₂		I ₃		I ₄		I ₅		I ₆		I ₇		I ₈		I ₉		I ₁₀		I ₁₁		I ₁₂		Regional Trend Index	
		<i>t</i>	<i>z</i>	<i>t</i>	<i>z</i>	<i>t</i>	<i>z</i>	<i>t</i>	<i>z</i>																		
1	41	-0.357	-0.455	-0.604	-0.455	-0.607	-0.190	-0.523	-0.342	-0.127	0.000	-0.732	-0.720	-2.446	-1.756	-2.446	-1.756	-2.383	-1.682	-2.383	-1.682	0.702	0.608	1.024	0.954	-0.333	
	33	-0.918	-0.833	-1.111	-1.023	-1.768	-1.988	-1.949	-1.823	-0.673	-0.455	-0.771	-0.341	-0.903	-0.708	-1.182	-0.829	-0.829	-0.765	-1.003	-0.765	0.513	1.142	0.325	1.067	-0.083	
	12	0.057	0.076	-0.086	0.038	-1.857	-0.726	-1.258	-0.303	0.036	0.038	-0.227	0.342	-1.473	-0.861	-0.991	-0.418	-1.487	-0.773	-1.190	-0.773	1.104	1.742	2.055	0.083		
	45	-0.437	0.000	-0.593	0.114	-1.434	-0.190	-1.821	-0.912	0.844	1.136	0.889	1.327	-1.684	-1.000	-1.181	-0.796	-1.840	-1.102	-1.588	-0.997	1.869	2.064	1.713	1.982	0.167	
	67	1.045	0.872	1.048	0.530	1.668	1.602	1.535	1.258	0.314	0.493	0.582	0.720	1.922	1.769	1.922	1.769	1.884	1.707	1.884	1.707	1.696	2.048	1.832	2.167	0.167	
	52	1.743	1.894	2.068	1.970	2.003	1.155	1.902	1.240	0.997	0.833	1.544	1.061	1.472	1.121	1.328	1.121	1.384	1.121	1.384	1.121	2.785	2.771	2.773	2.929	0.333	
	29	0.756	1.061	1.089	0.985	2.768	2.352	2.622	2.312	-0.983	0.000	-0.753	0.265	2.383	2.131	2.443	2.138	2.440	2.126	2.371	2.223	0.231	0.379	0.457	0.152	0.500	
2	80	0.497	0.000	0.852	0.152	-0.593	-0.381	-0.823	-1.027	1.204	1.212	1.743	1.709	-1.530	-2.102	-1.530	-2.037	-1.530	-1.962	-1.530	-1.962	0.483	-0.266	0.511	0.000	-0.333	
	85	-0.557	-0.455	-1.322	-0.682	-0.950	-1.645	-0.383	-0.684	-0.032	-0.227	-0.458	-0.379	-1.287	-1.962	-1.287	-1.822	-1.287	-1.962	-1.287	-1.962	0.557	-0.190	-0.376	-0.950	-0.250	
	12	-0.935	-0.909	-1.001	-0.833	-2.158	-1.818	-2.139	-1.970	-0.922	-0.152	-0.864	0.000	-0.807	-1.121	-1.150	-1.121	-0.945	-1.121	-1.388	-1.121	0.525	0.645	0.594	0.836	-0.167	
	43	0.641	-0.379	0.557	-0.227	0.986	0.343	0.959	0.038	0.399	-0.303	-1.229	-1.972	1.134	0.462	1.358	0.548	1.185	0.544	1.455	0.670	0.656	0.341	0.217	-0.228	-0.083	
	82	0.304	0.341	0.302	0.606	0.756	-0.343	0.769	-0.190	1.681	1.554	2.031	1.439	0.834	0.771	0.834	0.771	0.834	0.771	0.834	0.771	0.185	-0.494	0.697	-0.153	0.083	
	3	14	-2.443	-2.085	-2.458	-2.121	-1.992	-1.593	-2.536	-1.937	-2.321	-1.742	-2.452	-1.972	-0.748	-0.781	-0.748	-0.781	-0.660	-0.773	-0.660	-0.773	0.047	-0.607	-0.427	-0.959	-0.500
12	-0.763	-0.758	-0.719	-0.682	-0.936	-1.061	-0.740	-0.909	-2.089	-1.894	-2.777	-2.088	-1.795	-2.071	-1.795	-2.071	-1.603	-1.731	-1.603	-1.731	-1.445	-1.296	-1.442	-1.449	-0.333		
1	-2.024	-2.500	-1.868	-2.121	-0.602	-1.140	-1.148	-1.253	-2.107	-2.204	-1.753	-1.554	-0.423	-0.879	-0.423	-0.879	-0.445	-0.870	-0.445	-0.870	-1.341	-1.070	-0.805	-0.535	-0.250		
22	-0.833	-0.948	-0.519	-0.455	-0.152	-0.228	-0.314	-0.229	-1.165	-0.720	-1.056	-0.759	-2.259	-1.096	-1.953	-1.009	-2.355	-1.091	-2.115	-1.091	-1.064	-1.366	-0.056	-0.800	-0.250		
28	-1.097	-0.379	-0.560	0.000	0.924	0.873	0.380	1.138	-2.076	-2.045	-2.124	-1.972	0.915	1.356	1.448	1.458	0.844	1.353	1.361	1.450	-0.132	-0.266	-0.121	-0.343	-0.167		
33	-1.298	-1.175	-1.093	-0.606	1.623	1.480	2.547	2.428	-2.445	-2.312	-2.457	-1.669	1.476	1.376	1.221	1.376	1.695	1.454	1.599	1.363	-1.327	-1.596	-0.893	-1.485	-0.083		
5	0.018	-0.303	0.444	-0.076	1.260	1.062	1.427	1.101	-0.453	-0.872	0.832	0.000	1.603	1.944	1.432	1.654	1.486	2.049	1.543	1.882	0.388	-0.569	0.826	-0.383	0.083		
15	0.564	0.530	0.885	0.455	0.623	0.228	0.707	0.114	-1.005	-0.456	-1.251	-1.023	0.630	0.000	0.596	0.000	0.682	0.097	0.599	0.000	2.284	1.026	1.289	0.195	0.083		
50	-0.264	-0.379	-0.097	-0.341	1.856	1.833	1.123	1.329	-0.673	-0.606	0.136	0.530	1.187	1.691	2.125	2.103	1.325	1.674	2.113	1.717	-0.114	-0.910	0.498	-0.153	0.167		
27	1.169	0.758	1.946	1.515	2.004	1.889	2.003	2.128	0.588	0.303	0.913	0.303	0.862	1.202	0.227	0.789	1.329	2.007	0.794	1.893	1.088	0.341	0.420	0.000	0.250		
4	10	-2.012	-1.894	-2.018	-1.667	-1.839	-1.061	-0.653	-0.948	-2.178	-1.515	-1.816	-1.214	-0.963	-1.372	-1.432	-1.567	-1.062	-1.380	-1.663	-1.463	-1.542	-0.418	-1.230	-0.769	-0.250	
	50	-0.233	-0.379	-0.043	-0.190	-2.088	-2.015	-1.199	-0.991	-0.521	-0.379	-0.312	0.000	0.227	0.372	0.227	0.372	-0.611	0.157	-0.611	0.157	0.123	-0.038	0.129	-0.114	-0.083	
	69	-0.439	-0.455	-0.373	0.000	-1.275	-1.719	-1.243	-1.979	-0.197	-0.152	-0.149	-0.341	0.000	0.118	0.000	0.118	-0.042	0.059	-0.055	0.000	0.165	0.152	0.229	0.304	-0.083	
5	38	-0.487	0.152	-0.384	-0.379	-2.027	-1.290	-1.596	-0.645	0.918	1.138	1.587	1.556	-1.961	-1.297	-1.633	-1.185	-2.000	-1.354	-1.849	-1.354	-1.374	-0.533	-0.560	-0.153	-0.250	
	27	-1.454	-1.515	-1.076	-1.591	0.510	1.714	-0.787	0.797	-1.622	-1.364	-1.637	-1.251	0.253	1.160	1.311	1.327	0.093	1.102	0.900	1.156	-2.717	-1.671	-2.659	-1.982	-0.167	
	10	-1.084	0.038	-1.176	0.000	-0.712	-0.229	-0.924	-0.190	-0.442	0.493	-0.478	0.342	-0.808	-0.590	-0.784	-0.590	-0.806	-0.530	-0.776	-0.647	-2.453	-1.138	-1.704	-0.879	-0.083	
	30	-1.162	-0.530	-1.302	-0.455	-0.807	-0.190	-1.118	-0.303	-0.417	0.190	-0.571	0.000	-0.227	0.000	-0.227	0.000	-0.419	-0.262	-0.419	-0.262	-1.928	-1.101	-1.972	-1.609	-0.083	
	49	-0.490	0.152	-1.103	-0.227	-2.125	-1.371	-1.707	-0.569	0.200	0.758	-0.299	0.152	-1.004	-0.735	-0.873	-0.735	-0.956	-0.682	-0.843	-0.787	-0.396	0.000	0.680	0.191	-0.083	
	15	1.071	1.742	0.663	1.439	0.560	1.366	0.717	1.327	2.166	1.894	1.875	1.897	0.494	0.586	0.494	0.586	0.292	0.483	0.292	0.483	1.502	1.441	1.314	1.446	0.083	
	19	2.077	0.872	1.771	1.061	-1.664	-0.493	-1.070	-0.227	1.973	1.136	1.946	1.175	-1.003	-0.875	-1.003	-0.875	-1.164	-0.870	-1.164	-0.870	0.961	0.267	0.882	0.114	0.167	
	12	2.009	0.909	1.153	0.833	-0.575	-0.721	0.157	0.076	1.357	0.531	1.139	0.455	-0.817	-0.740	-1.503	-0.983	-0.940	-0.727	-1.560	-1.000	3.292	2.244	2.971	1.646	0.250	
	6	81	-1.147	-1.591	-1.079	-1.439	-0.481	-0.797	-0.438	-0.228	-0.967	-1.212	-0.725	-0.833	0.502	0.633	0.643	0.789	0.572	0.682	0.644	0.682	-2.393	-2.352	-2.647	-2.436	-0.167
		2	-0.362	-0.493	-0.389	-0.985	0.693	-0.266	0.884	-0.038	-0.692	-0.909	-0.342	-0.796	0.903	0.089	0.610	0.000	1.052	0.000	0.854	0.000	-1.686	-0.949	-2.443	-1.485	-0.083
10		-1.176	-0.455	-1.065	-0.227	0.242	0.190	0.149	0.265	-1.309	-0.227	-1.075	-0.228	0.169	0.422	0.456	0.584	0.198	0.367	0.438	0.472	-2.107	-1.253	-1.415	-0.570	-0.083	
12		-0.667	-1.212	-0.691	-1.212	0.845	0.076	0.491	-0.152	-1.342	-1.818	-1.400	-1.970	0.349	0.000	0.349	0.000	0.548	0.091	0.548	0.091	-1.246	-1.329	-1.094	-1.520	-0.083	
16		-1.338	-1.439	-1.345	-1.591	-0.388	-0.531	-0.621	-0.835	-1.825	-1.742	-2.331	-2.165	0.285	0.315	0.000	0.211	0.249	0.367	0.002	0.262	0.278	0.000	-0.119	0.000	-0.083	
19		-1.360	-1.061	-1.222	-1.591	-0.787	-0.948	-1.127	-1.212	-0.915	-1.253	-0.958	-1.402														

Table 5. Cont.

Months	Polygon	I ₁		I ₂		I ₃		I ₄		I ₅		I ₆		I ₇		I ₈		I ₉		I ₁₀		I ₁₁		I ₁₂		Regional Trend Index	
		t	z	t	z	t	z	t	z	t	z	t	z	t	z	t	z	t	z	t	z	t	z	t	z		
7	40	-0.169	-0.379	-0.272	-0.682	-0.218	-0.417	0.327	-0.266	-0.333	-0.114	-0.378	0.000	-0.683	-0.422	-0.809	-0.478	-0.758	-0.577	-0.897	-0.577	-1.640	-2.236	-1.592	-2.012	-0.167	
	5	-0.017	-0.530	-0.239	-0.303	0.258	-0.190	0.353	0.114	0.008	-0.190	0.015	-0.379	0.810	-0.158	0.810	-0.158	0.732	-0.157	0.732	-0.157	-1.903	-1.711	-1.986	-1.907	-0.083	
	76	-0.855	-0.985	-1.003	-0.833	-0.071	0.341	0.004	0.379	-1.839	-1.742	-1.730	-1.517	0.643	0.368	0.643	0.368	0.693	0.367	0.693	0.367	-1.773	-1.857	-1.853	-2.088	-0.083	
8	56	-1.019	-0.455	-1.016	-0.303	-0.674	-0.910	-0.557	-0.797	-0.928	-0.493	-0.870	-0.456	-1.606	-2.014	-1.606	-2.014	-1.661	-2.178	-1.661	-2.178	0.092	-0.265	0.211	-0.115	-0.333	
	4	0.982	0.833	0.935	1.515	0.910	1.405	1.558	1.972	0.239	0.000	0.086	0.000	1.193	1.172	0.977	1.084	1.018	1.160	0.818	1.063	0.412	0.076	0.675	0.495	0.083	
	86	0.634	0.455	0.483	0.530	2.736	2.197	2.171	1.630	-0.308	-0.114	-0.368	-0.493	-0.149	0.000	-0.149	0.000	-0.181	0.000	-0.181	0.000	-0.353	0.000	-0.704	-0.304	0.167	
9	48	-1.270	-1.212	-1.109	-1.439	-0.327	-1.175	-0.838	-1.517	-1.671	-1.970	-1.911	-2.312	-1.898	-2.065	-1.898	-1.850	-1.430	-1.999	-1.430	-1.908	-1.053	-1.251	-0.855	-1.181	-0.333	
	49	-2.222	-1.706	-2.511	-1.782	-1.836	-0.948	-1.932	-1.439	-1.595	-1.288	-1.526	-1.023	-1.399	-1.179	-1.399	-1.179	-1.398	-1.118	-1.398	-1.118	-1.391	-0.455	-1.612	-0.533	-0.167	
	9	-1.003	-1.136	-1.099	-1.023	-0.085	-0.493	0.158	-0.493	-1.763	-1.818	-1.825	-2.048	0.963	0.118	1.424	0.118	0.986	0.177	1.273	0.177	-0.394	-0.266	-0.214	-0.685	-0.083	
	47	-0.648	0.379	-0.658	0.114	-1.257	-1.101	-1.515	-1.099	-2.063	-0.909	-1.881	-0.720	-0.965	-0.590	-0.716	-0.472	-0.849	-0.530	-0.613	-0.412	0.916	1.368	0.901	1.066	-0.083	
	66	-0.956	-0.833	-1.121	-0.909	0.943	0.723	0.926	0.948	-0.982	-0.682	-1.049	-0.910	1.961	1.651	1.961	1.651	1.948	1.707	1.948	1.707	-2.217	-1.366	-2.067	-1.720	0.000	
	16	-0.228	0.227	-0.175	0.000	3.124	2.540	2.424	2.352	-0.502	-0.114	-0.382	-0.076	1.508	1.315	1.508	1.315	1.660	1.312	1.660	1.312	0.338	0.531	0.674	0.267	0.167	
10	25	0.259	0.000	-0.228	-0.758	-1.701	-1.061	-0.382	-0.606	-0.139	-0.682	-0.108	-0.455	-1.547	-0.958	-1.639	-1.056	-1.743	-0.813	-1.843	-0.894	-2.167	-2.276	-1.699	-1.714	-0.083	
	60	0.111	-0.076	0.336	0.000	2.211	1.824	2.231	2.168	-0.030	-0.417	0.212	-0.303	1.445	1.434	1.445	1.434	1.311	1.102	1.311	1.102	0.684	0.304	1.031	0.306	0.167	
	63	1.123	-0.455	0.624	-0.417	1.162	0.038	1.217	0.038	0.139	-1.212	-0.607	-1.288	0.603	0.000	0.603	0.000	1.031	0.052	1.031	0.052	2.012	0.986	2.007	1.072	0.167	
	8	2.603	1.288	2.797	1.515	1.681	0.988	1.692	1.253	1.427	0.379	1.345	0.152	0.282	0.280	0.330	0.280	0.294	0.280	0.317	0.280	2.532	1.515	2.416	1.214	0.333	
11	40	-0.688	-1.364	-0.695	-1.212	-0.848	-0.988	-0.808	-0.797	-0.412	-0.530	-0.295	-0.458	-1.731	-2.005	-1.691	-2.060	-1.720	-2.060	-1.672	-2.060	0.082	-0.379	0.148	-0.534	-0.333	
	90	0.712	0.758	0.723	0.909	-1.788	-0.722	-1.977	-1.364	1.629	1.630	2.346	2.236	0.322	0.000	0.322	0.000	0.164	0.000	0.164	0.000	0.932	1.062	0.359	0.381	0.000	
	7	0.754	1.364	-0.058	0.644	1.381	0.910	1.540	0.569	1.438	1.591	1.091	1.214	1.824	1.769	1.824	1.769	1.864	1.825	1.864	1.825	1.732	2.164	1.401	1.714	0.083	
	87	1.248	1.439	1.295	1.366	0.842	1.677	1.682	2.051	0.978	0.985	0.607	0.228	1.011	1.074	1.011	1.074	1.011	0.767	1.016	0.767	1.016	0.795	0.644	0.098	0.304	0.083
	86	0.360	0.114	0.361	-0.038	-1.307	-1.075	-1.179	-1.298	0.667	1.099	0.453	0.494	-0.903	-0.590	-0.904	-0.592	-0.996	-0.647	-1.088	-0.647	2.423	1.820	2.331	1.896	0.167	
	43	1.873	1.970	2.488	2.200	1.446	1.525	2.212	1.669	1.156	1.439	1.314	1.591	-0.265	0.000	-0.265	0.000	-0.103	0.000	-0.103	0.000	1.026	1.138	0.862	0.797	0.250	
12	43	1.277	0.530	1.348	0.227	2.094	1.758	1.655	1.327	0.669	-0.379	0.763	-0.455	1.056	1.682	1.183	1.682	1.024	1.682	1.144	1.682	1.718	1.821	1.581	1.603	0.083	
	60	1.391	1.364	1.163	1.667	1.933	1.517	2.003	1.824	0.661	0.758	0.436	0.796	0.649	1.017	0.929	1.077	1.040	1.181	1.368	1.272	0.848	1.177	0.726	1.677	0.083	
	36	2.345	1.818	2.299	1.818	-0.040	-0.531	0.809	0.683	1.560	1.364	1.856	1.402	1.375	1.366	1.511	1.471	1.262	1.416	1.277	1.416	1.911	1.672	1.119	0.880	0.167	
	45	0.977	0.606	0.913	0.758	2.286	1.575	1.970	0.913	0.693	0.682	0.675	0.190	1.606	0.354	1.606	0.354	1.631	0.412	1.631	0.412	1.629	1.027	1.498	0.496	0.167	
	47	2.358	1.667	2.196	1.970	-0.115	0.000	1.550	0.797	1.062	1.327	1.170	1.441	1.256	1.016	1.390	1.069	1.095	0.966	1.103	1.063	1.743	1.368	0.788	0.153	0.167	
	58	2.177	1.630	2.197	1.742	1.387	1.493	1.384	1.294	2.229	1.251	1.982	1.478	1.287	1.156	1.287	1.156	1.287	1.156	1.287	1.156	1.287	1.156	1.726	0.839	1.321	0.533
	77	2.127	1.439	2.269	1.894	0.511	1.103	0.615	1.142	2.396	1.708	2.561	1.745	1.249	0.825	1.095	0.825	1.003	0.765	0.750	0.765	1.726	0.721	1.420	0.457	0.333	
	10	2.248	1.251	2.095	1.515	1.694	1.410	1.673	1.292	2.091	1.517	2.018	1.441	1.634	1.542	1.616	1.542	1.582	1.542	1.556	1.542	1.980	0.836	2.045	0.651	0.500	

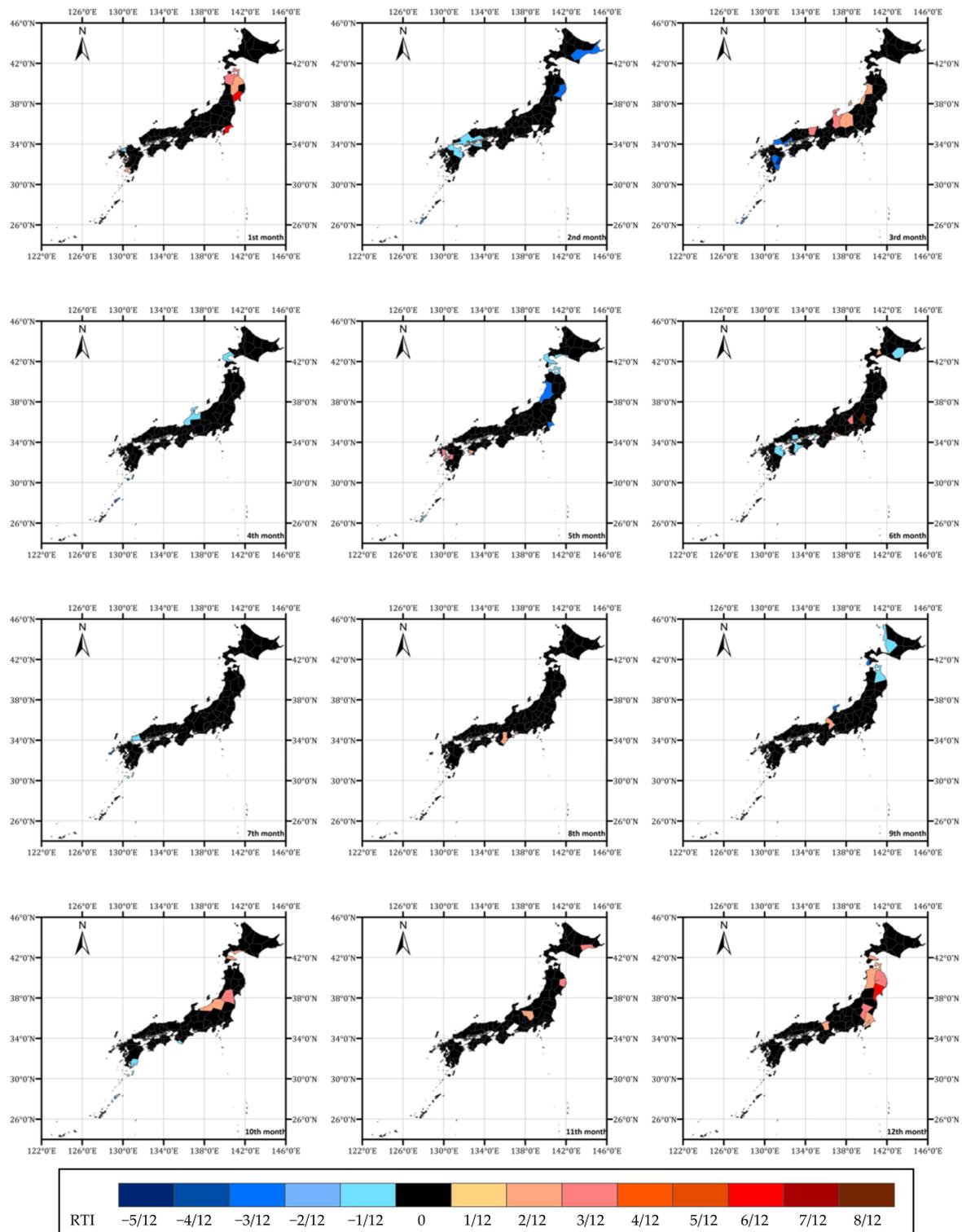


Figure 7. The regional trend indices for homogenous regions (The regional trend index RTI indicates the percentage of trend indices with significant trend, for a homogenous region and month).

The gauging sites in these homogenous regions with significant trends are listed in Table 6. The general picture of the regional analyses implemented along Japan has shown that temporal changes on heavy precipitations depend largely on season and geographical location. Statistically significant positive and negative trends have been detected, in

5. Conclusions

The Index Precipitation Method introduced herein provided a spatial consistency in the estimation of the temporal changes on heavy precipitations. Since homogenous regions are based on a statistically identical frequency distribution and parameters, the observed trends preserve statistical properties of precipitation frequencies on a considered study area. Homogenous regions and observed trends are less sensitive to sampling variation. Use of L-moment estimates in the delineation processes facilitated homogeneity analyses, while preserving high order statistical moments. Homogenous regions are geographically continuous and enclose geographically contiguous sites. In this respect, GIS provided an effective environment for the delineation processes and for the determination of homogenous regions.

Relying on the 12 indices that characterize heavy precipitations, the regional trend analyses made possible to take both the regional and temporal sampling variations into account. These indices define not solely the changes on the mean or extreme statistics of single-site observations, but also the regional changes on heavy precipitation. As a regional figure of merit and a trend severity measure, the regional trend index is a good indicator of the existence of a regionally verified consistent trend. Trend analyses along Japan show that temporal changes of heavy precipitations depend highly on geographical location. Upper, middle and lower latitudes have similar trend behaviour. The regional trends change monthly. Rainy and cold periods indicate more evident trends.

It is concluded in the study that the Index Precipitation Method usefully defines homogenous regions for which the frequency behaviors of multisite precipitation observations are preserved. It is valuable especially for removing negative effects of sampling variations, missing data, regionally unsupported trend tests, use of statistical moment-based parameters, on single-site trend analyses. The Index Precipitation Method would be also advantageous for various hydrologic studies that need regional homogeneity and spatial verification/support for temporal changes. It is expected to reduce ambiguities in hydrologic modelling related to single-site temporal processes.

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