



Gender Equality and Economic Diversification

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Received: 27 February 2019; Accepted: 3 April 2019; Published: 11 April 2019



Abstract: We show that gender inequality decreases the variety of goods countries produce and export, in particular in low-income and developing countries. We argue that this happens through at least two channels: first, gender gaps in opportunity, such as lower educational enrollment rates for girls than for boys, harm diversification by constraining the potential pool of human capital available in an economy. Second, gender gaps in the labor market impede the development of new ideas by decreasing the efficiency of the labor force. Our empirical estimates support these hypotheses, providing evidence that gender-friendly policies could help countries diversify their economies.

Keywords: gender inequality; female labor force participation; economic diversification; economic growth; human capital; low-income and developing countries

1. Introduction

Although commodity prices have recently somewhat recovered, the preceding episode of decline and the associated volatility of commodity prices are a powerful reminder for countries—especially those rich in resources—to diversify their output and export bases. The drop in oil and other commodity prices in recent years has put substantial pressure on many resource-intensive countries, with growth declining in many of them, and significant macroeconomic adjustment needs arising, since export and fiscal revenues have declined markedly (IMF 2016). While oil prices have been increasing from their low of less than 30 USD/barrel in early 2016, they are still significantly lower than their peak 2013 levels. Commodity prices more generally are expected to remain at only a fraction of their high levels in the medium term. As a result, reforms to stimulate product and export diversification have gained renewed importance on policy makers' agendas, in particular in resource-intensive economies.

Indeed, a substantial body of the literature has highlighted economic diversification as a driver of sustainable growth at the early stages of development (see Figures 1 and 2). A long-held tenet of international trade, Ricardo's theory of comparative advantage, promotes the idea that countries should specialize in the production of goods and services they can produce at lower relative opportunity cost. Historically, many low-income countries have relied on relatively few trading partners and specialized in commodity and primary products, mainly due to their resource endowments, as might be predicted by the Heckscher-Ohlin model. Yet, as many countries have experienced, even in the current episode of lower commodity prices, the lack of diversification is associated with both lower economic growth and higher volatility. The literature has now well established that diversification and structural transformation—the continued, dynamic reallocation of resources to more productive sectors and activities—are associated with economic growth, particularly at the early stages of development

(IMF 2014a; Papageorgiou and Spatafora 2012).¹ Given this particular importance of diversification at earlier stages of development, this article focuses in particular on low-income and developing countries, while also examining a larger set of countries (106 countries for export diversification and 108 countries for output diversification in the baseline regressions, as listed in Appendix B).

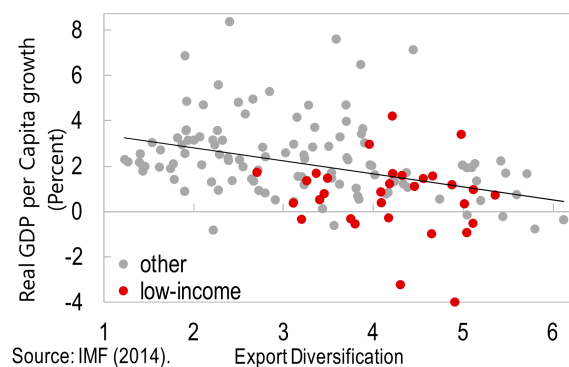


Figure 1. Export product diversification and output growth, 1962–2010. (*higher diversification values = less diversification*).

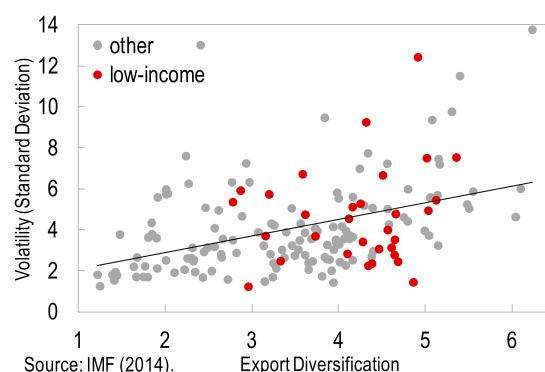


Figure 2. Export diversification and output volatility, 1962–2010. (*higher diversification values = less diversification, volatility = standard deviation over 1962–2010*).

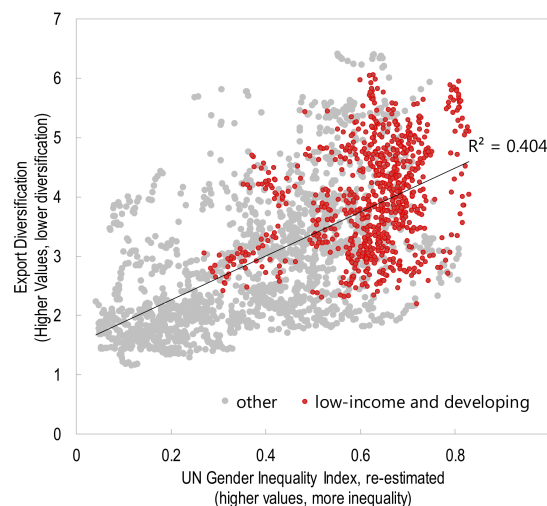
Several structural country characteristics and policies have been shown to be associated with economic diversification, with education taking on a prominent role. IMF (2014b) highlights a range of country characteristics and policies, such as the level of development, institutional quality, stronger infrastructure, and a higher degree of globalization being strongly associated with diversification. In addition, it confirms the results of other studies that have shown a well-educated workforce matters not only for diversification, but also is strongly associated with export quality upgrading (Dabla-Norris et al. 2013). Increasing human capital accumulation fosters economic diversification by promoting the development of skill-intensive industries and new technologies and by facilitating technological diffusion between firms (Bal-Gunduz et al. 2015). Whereas primary and secondary education can enable a country to imitate frontier technology, tertiary education can increase its possibility of innovating (Aghion and Howitt 2006).

Building on this literature, we introduce gender equality as an additional determinant of economic diversification with two main hypotheses:

¹ The process of structural transformation is characterized by two dimensions: horizontal (across sectors) and vertical (within a sector). Diversification into new higher value-added sectors is the horizontal dimension. Quality upgrading is the vertical dimension and focuses on producing higher quality (and generally higher priced) products within existing sectors (IMF 2014a).

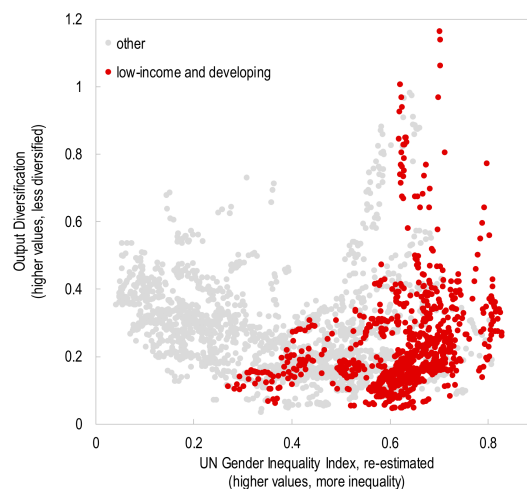
- First, gender gaps in opportunity, such as in education, harm diversification directly by constraining the potential pool of human capital. In particular, in countries where girls' education lags that of boys, female human capital cannot accumulate optimally, therefore slowing down technology adoption and innovation ("human capital channel").
- Second, gender gaps impede the development of new ideas indirectly by decreasing the efficiency of the labor force. Gender gaps in labor force participation shrink the pool of talent from which employers can hire and limit the number of female entrepreneurs (Cuberes and Teignier 2016; Esteve-Volart 2004; Christiansen et al. 2016a, 2016b). This limitation, in turn, impedes a country's ability to create and execute ideas, i.e., to diversify ("resource allocation channel").

In fact, a look at the data shows that gender inequality and economic diversification indeed appear to be inter-linked phenomena (see Figures 3 and 4). High levels of gender inequality, as measured by an extended version of the United Nations' Gender Inequality Index, are associated with lower levels of export diversification (a combined measure of export product variety and equality in export shares), while they are negatively related with output diversification (a measure of equality in the contribution of sectors to real output, including services) mainly in low-income and developing countries. Box 1 in Section 3 describes these indices in more detail.



Sources: World Bank World Development Indicators, United Nations, IMF (2014), and IMF staff calculations.

Figure 3. Export diversification and gender inequality, 1990–2010.



Sources: World Bank World Development Indicators, United Nations, IMF (2014), and IMF staff calculations.

Figure 4. Output diversification and gender inequality, 1990–2010.

Our empirical analysis shows that gender inequality in opportunities and in the labor market is associated with lower diversification. In particular, the contribution of this paper is three-fold:

- *First, we present empirical evidence that gender inequality is negatively associated with both output and export diversification in low-income and developing economies.* The effect of gender inequality on economic diversification comes on top of the effect of the standard drivers of diversification identified in the literature. While the negative effects of gender inequality and the positive effects of diversification on economic growth have found support in these two separate literatures, to our knowledge, the connection between them has not yet been established.
- *Second, our results suggest that both inequality of opportunities and lower female labor force participation are associated with lower economic diversification.* These findings support our two main hypotheses. The negative relationship between inequality of opportunity and diversification supports the hypothesis of the *human capital channel*, while the association between female labor force participation and diversification supports the premise of the *resource allocation channel*, which reduces the creation of ideas and development of sectors.
- *Third, we provide evidence on causality.* Gender inequality and diversification are interlinked phenomena and, as described in more detail in Section 2, the literature so far has mainly focused on how structural transformation coincides with episodes of improvements in gender equality (Akbulut 2011; Olivetti and Petrongolo 2014; Ngai and Petrongolo 2017; Rendall 2013). The novel aspect of our study is to examine whether gender inequality affects diversification and to address endogeneity concerns in our regressions.

The paper proceeds as follows. Section 2 provides a brief overview of the literature. Section 3 presents the more technical details on the empirical strategy and data. Box 1 highlights the development in some of the key indicators of gender inequality and relates them to measures of economic diversification. Section 4 highlights the paper's main results, and Section 5 concludes.

2. Literature Review

To our knowledge, this article is the first empirical study to establish the negative effects of gender inequality on economic diversification. However, as outlined in this section, the positive effects of diversification and the negative effects of gender inequality on economic growth have been documented at length in these two separate literatures.

According to the former literature, diversification, development, and growth are closely interlinked, in particular in low-income countries.

- *Despite significant cross-country heterogeneity, greater diversification has been associated with improved macroeconomic performance:* higher growth, reduced volatility, and increased resilience to external shocks (Koren and Tenreyro 2007; Cadot et al. 2011). Singer (1950) demonstrated that a country's initial level of diversification is positively correlated with economic growth. Using an Instrumental Variable Bayesian Model Averaging approach to move beyond correlations, IMF (2014a) finds that for low-income countries, extensive diversification (introducing new product lines), intensive diversification (creating a more balanced mix of existing products), and the broader process of output diversification are indeed drivers of economic growth. Diversification also involves shifting resources from sectors with high volatility, such as mining and agriculture, to sectors with less volatility, such as manufacturing, resulting in greater stability. Countries with more diversified production structures tend to have lower volatility of output, consumption, and investment (Moore and Walkes 2010; Mobarak 2005).
- *There is a non-linear relationship between diversification and development (Imbs and Wacziarg 2003).* As countries develop, they diversify until they reach a critical point. Beyond this point, they start specializing in low-volatility sectors (Imbs and Wacziarg 2003; Koren and Tenreyro 2007; Cadot et al. 2011).

The latter literature documents a negative link between real GDP per capita growth and gender inequality (Elborgh-Woytek et al. 2013). On a macro level, the relationship between gender inequality and economic growth has been a topic of increasing interest in the academic and policy literature in recent decades. Dating back to the early 1990s, a special issue of World Development was dedicated to introducing a gender lens to macroeconomics (Çağatay et al. 1995). Since then, a broad literature has developed on the topic of gender inequality and its connection to economic development and growth (see, e.g., the World Bank (2012) World Development Report: Gender Equality and Development).

Economic development has been shown to decrease gender inequality, while persistent discrimination against women can also adversely affect development (Goldin 1994; Hill and King 1995; Dollar and Gatti 1999; Tzannatos 1999; Stotsky 2006; Cuberes and Teignier 2014). Whereas our article focuses on the latter direction of causality, many others have explored the former (e.g., Galor and Weil 1996; Fernandez 2007; Alesina et al. 2013; Dufo 2012). The following results demonstrate some of the channels through which gender inequality can negatively impact macroeconomic performance:

- **Education.** Studies have confirmed the negative effect of gender inequality in education on growth (Hill and King 1995; Engelbrecht 1997; Forbes 2000; Dollar and Gatti 1999; Klasen 1999; Knowles et al. 2002; Klasen and Lamanna 2009; Seguino 2010). Dollar and Gatti (1999) find that gender inequality in education negatively impacts growth in countries where female educational attainment is high. Klasen (1999) demonstrates that the negative effect is present in all economies.² Berge and Wood (1994) provide support for the hypothesis that an educated female labor force is a determinant of manufacturing exports growth. Using broader measures of gender inequality going beyond education gaps, a recent study by Amin et al. (2015) confirms their strong negative impact on economic growth but only in poor countries.

We hypothesize that these negative effects of gender inequality in educational opportunities affect growth at least in part by obstructing the economic diversification process.

- **Occupation.** Occupational choice models are based on the assumption that men and women have the same distribution of talent (Cuberes and Teignier 2012; Esteve-Volart 2004). Gender gaps in entrepreneurship distort the efficient allocation of talent (Cuberes and Teignier 2012). As a certain percentage of women are prevented from becoming entrepreneurs, they are forced to work as employees, thus increasing the supply of labor. As a result, equilibrium wages and aggregate productivity fall. Gender gaps in labor force participation are modeled as preventing a fraction of women from supplying labor to the market, hence decreasing income per capita. Cuberes and Teignier (2016) present an updated version of the model in which women also have the choice to become self-employed, in addition to being entrepreneurs and workers. In this version of the model, women face two additional exogenous restrictions: only a fraction can become self-employed, and those who become workers receive lower wages than men do. The main results are not qualitatively different. Esteve-Volart (2004) makes explicit the negative endogenous effect of gender gaps in education on growth: the suboptimal allocation of managerial talent explicitly leads to lower female human capital accumulation and thus, slower technology adoption and innovation, which reduces aggregate output and obstructs economic growth. The negative effects of gender discrimination in managerial talent allocation are more serious for sectors where high-level skills are needed, such as the non-agricultural sector, whereas restricted female labor force participation in general impacts all sectors, including agriculture. Finally, using a model of endogenous savings, fertility, and labor market participation, Cavalcanti and Tavares (2016)

² Earlier studies have shown somewhat different results: Barro and Lee (1994) and Barro and Sala-i-Martin (1995) find that female secondary education has a negative impact on growth, as low female educational attainment signifies “backwardness” and hence higher growth potential. Klasen (1999) and Lorgelly and Owen (1999), however, suggest that the finding may reflect multicollinearity problems resulting from the inclusion of both female and male education variables in the regression analysis and the disproportionate influence of a few outlier countries.

show that an increase of 50 percent in the gender wage gap could lead to a decrease in income per capita by 35 percent.

We explore whether the channels posited in these models affect growth via their effects on the dynamic process of diversification and structural transformation of the economy.

- **Aggregate measures of gender inequality and growth.** Recent empirical evidence, using an extended version of the UN's Gender Inequality Index (GII), shows that several dimensions of gender inequality (health, empowerment, education attainment, and labor force participation) are strongly associated with lower growth, in particular in low-income countries (Gonzales et al. 2015b; Hakura et al. 2016). Box 1 in Section 3 describes the GII in more detail.

In this study, we test whether measures of gender inequality are also related to lower export and output diversification.

- **Gender wage inequality** has had a positive effect on export-led growth in semi-industrialized export-oriented economies, while it has had a negative effect in low-income agricultural countries (Seguino 2000, 2010). On the other hand, accounting for the different productivity of male and female workers, Schober and Winter-Ebmer (2011) do not find support for the hypothesis that increased gender inequality contributes to growth, but argue that it may indeed hamper it.

Due to the lack of extensive and reliable data on wage inequality in low-income and developing countries, in this article, we focus instead on proxies, such as gender inequality in reproductive health, empowerment, and labor market participation, the sub-components of the multi-dimensional GII.

Some studies have explored the connection between the two literatures of gender inequality and economic diversification but they focus on the reverse direction of causation from the one explored in this article. Structural transformation has been shown to coincide with episodes of decreases in gender inequality, in particular in the service sector. Several studies examine the relationship between women's economic participation and structural transformation, and have focused predominantly on the influence of the service sector (Akbulut 2011; Olivetti and Petrongolo 2014; Ngai and Petrongolo 2017; Rendall 2013). Rendall (2013) finds that structural transformation has been important in reducing gender inequality and argues that this has happened by decreasing the labor demand for physical ("brawn") attributes. Economies with lower "brawn" requirements offer better labor market opportunities because they allow women to take advantage of their comparative advantage in less physical ("brain") attributes. For example, in Mauritius, the development of the textile industry coincided with an increase in female labor force participation of nearly 60 percent between 1983 and 1999 (Svirydzenka and Petri 2014). Cavalcanti and Tavares (2008) link increases in female labor force participation to increases in government expenditures, leading to higher demand for services provided by the government. This in turn further encourages female labor force participation, especially when the public sector typically employs more women.

These studies emphasize the direction of causation from structural transformation of the economy to women's economic participation. The novelty of our research is that we explore the reverse relationship, namely whether greater gender equality can enhance and support the process of structural transformation.

Gender-based legal restrictions exert a significant impact on women's economic participation, which is why we argue they are valid instruments to address endogeneity concerns in our regressions and provide evidence for the hypothesized direction of causality from gender inequality to economic diversification. Restrictions to participation have been shown to negatively affect women's access to finance (Demirgüç-Kunt et al. 2013), employment (Amin and Islam 2014), labor force participation (Gonzales et al. 2015b), asset ownership and wealth (Deere et al. 2013), property rights (Razavi 2003), and adoption of new technologies (Quisumbing and Pandolfelli 2010). A recent IMF study uses the comprehensive database compiled in the World Bank's Women, Business and the Law Report (World Bank 2013; World Bank Group 2015) to demonstrate that restrictions on women's rights to

inheritance and property, as well as legal impediments to economic activity, such as the right to open a bank account or to freely pursue a profession, significantly exacerbate gender gaps in labor force participation (Gonzales et al. 2015a).

Our study uses the results from this stream of the literature to argue that gender-based legal restrictions are valid instruments to tackle endogeneity concerns in the analysis of the impact of gender inequality on diversification: legal restrictions exacerbate gender inequality, which, in turn, impedes output and export diversification.

To our knowledge, there are no theoretical studies on the impact of gender inequality in opportunities and outcomes on output and export diversification. Most theoretical studies of the impact of gender inequality on growth have examined the causal channels of fertility and the education of children (Galor and Weil 1996; Lagerlöf 2003; Cavalcanti and Tavares 2016; Doepke and Tertilt 2009; Agénor et al. 2014). Hence, the empirical investigation in this study is broadly based on the theoretical occupational choice models of Cuberes and Teignier (2012) and Esteve-Volart (2004), which examine the effects of gender discrimination on aggregate output and economic growth.

We explore whether the channels posited in these models are similarly at play concerning the process of diversification of the economy.

3. Empirical Strategy

We analyze the effect of gender inequality on diversification together with determinants previously highlighted in the literature. To obtain unbiased estimates, we control for unobservable variables that differ across countries, as well as common effects over time in the following relationship for the period 1990–2010 in our baseline estimations:

$$\begin{aligned} \text{Diversification}_{it} = & \beta_1 \text{Gender Inequality}_{it} + \beta_2 \text{GenderInequality}_{it} \cdot \text{LIDC} + \\ & \gamma' \text{Structural Characteristics}_{it} + \delta' \text{Policies}_{it} + \varphi' \text{Institutions}_{it} + \\ & + \tau' \text{Cyclical Factors}_{it} + \mu_i + \theta_t + \varepsilon_{it}, \end{aligned}$$

in which

- *Diversification_{it}* represents the measure of either export or output diversification as defined in Box 1 for country *i* at time *t*.
- *The main contribution of our paper is to test whether gender inequality exerts a significant effect on diversification. Gender Inequality_{it}* tests for this effect at two levels: first, to account for the combined effect of several dimensions of gender inequality, we use the extended version of the United Nations Gender Inequality Index, i.e., a combination of gaps in labor force participation, education, and reproductive health, as well as female seats in parliaments as described in Box 1. In a second step, to test for the effect of individual measures of gender inequality, the index is replaced by the female-to-male gross enrollment ratio in secondary school, the female labor force participation rate, the share of female seats in parliament, the adolescent fertility rate, and the risk of maternal death. As the relationship between diversification and gender inequality may vary across levels of development, we include a low-income and developing country interaction term (*LIDC*) in our main regressions.
- *Structural Characteristics_{it}* may significantly impact a country's ability to diversify. We therefore include real GDP per capita and its square in the regression to account for the overall level of development, as well as the turning point after which countries re-concentrate their export or output structure (IMF 2014b; Dabla-Norris et al. 2013). The baseline regressions also include population size to capture the pool of workers potentially able to produce different products in a country, along with an index of human capital to account for a country's ability to generate and implement new ideas. In addition, we test whether being resource-rich exhibits a negative effect on diversification by introducing the share of mining in GDP or the share of fuel exports into the regressions.
- *Institutions_{it}* shape the environment in which businesses operate and the ease of entering a market to implement an idea or to produce a new product. To account for this impact, our regressions use both

general institutional quality (e.g., Frasier Institute Summary Index), as well as specific dimensions of the regulatory environment (e.g., legal systems and property rights).

- *Cyclical factors_{it} may boost or compress a certain sector in the short term, therefore impacting diversification over time.* We therefore introduce macroeconomic variables, such as terms-of-trade, real effective exchange rates, and real GDP growth into our regressions.
- *Policies_{it} may foster economic diversification.* Here, we test for several policy dimensions, such as more openness to trade (through an index of globalization, the degree of freedom to trade internationally, and average tariff rates), financial development (an index of financial reform, and interest rate controls and private sector credit-to-GDP as robustness checks), the scale of investment in the economy (investment in percent of GDP and per worker), and infrastructure development (density of landlines and length of road network).
- *To capture other factors over time and by country, we include μ_i and θ_t , that is country fixed effects and time fixed effects into our baseline regressions.* ε_{it} represents the error term.

In our baseline regressions, the data sample includes 106 countries (35 LIDC) for export diversification and 108 countries (36 LIDC) for output diversification, as listed in Appendix B. Please see Appendix A for detailed summary statistics and sources of the variables.

Box 1. Defining export and output diversification and gender inequality.

Export product diversification. We use the Theil index of export diversification from IMF (2014b), which follows Cadot et al. (2011). The index can be decomposed into a “between” and a “within” sub-index:

$$\begin{aligned} \text{Theil Index} &= \frac{1}{N} \sum_i^N \frac{\text{Export Value}_i}{\text{Average Exp. Value}} \cdot \ln \frac{\text{Export Value}_i}{\text{Average Exp. Value}} \\ &= \text{Theil}_{\text{between}} + \text{Theil}_{\text{within}}, \end{aligned}$$

in which i is the product index and N the total number of products. The “between” Theil index captures the *extensive margin* of diversification, i.e., the number of products, while the “within” Theil index captures the *intensive margin* (product shares). Lower values of the Theil index indicate higher levels of export product diversification. The index is available for 188 countries from 1962 to 2010 from the IMF Diversification Toolkit.

Output diversification. As services are not included in the calculation of export product diversification, we additionally use the output diversification Theil index in our regressions to account for the impact of changes in the service sector. Following the methodology used for the export Theil index described above, the output diversification index was constructed for the real subsectors from the UN’s sectoral database in IMF (2014b). The index covers 188 countries from 1970 to 2010 from the IMF Diversification Toolkit.

Gender Inequality Index (GII). The gender inequality index is the extended version of the United Nations Gender Inequality Index (Gonzales et al. 2015b; Stotsky et al. 2016), which captures gender inequality across areas of health (maternal mortality ratios and adolescent fertility rates), empowerment (share of parliamentary seats and education attainment at the secondary level for both males and females), and labor force participation (rates by sex). While the GII has drawbacks (such as a complicated functional form and a combination of indicators that compare men and women with indicators that pertain only to women), it is preferable to alternatives such as the GDI (in which one of the main components is not observed and is imputed). The index spans values between 0 and 1, with higher values indicating higher gender inequality. The index is available for 141 countries from 1990 to 2013 from the IMF GDI GII Database.

In addition to the fixed effects specifications, we address the endogenous relationship between economic diversification and gender inequality by using the instrumental variable generalized method of moments (IV-GMM) technique.³ Gender inequality in outcomes and opportunities may cause lower levels of export and output diversification, but lower levels of diversification may lead to larger gender inequalities in outcomes and opportunities. Therefore, to determine the direction of causality, we use IV-GMM in addition to the fixed effects specifications as highlighted above.⁴ In particular, the

³ See Bandiera and Natraj (2013) for a discussion of panel regressions and the endogenous relationship between gender inequality and growth.

⁴ All regressions are estimated using heteroskedasticity-robust Huber-White standard errors.

instrumental variables (IV) approach isolates the causal effect of the country-specific degree of gender inequality, as measured by the GII, on export and output diversification.

We introduce legal rights for women as instruments into our specifications. To be valid, an instrument needs to fulfill two criteria: (i) not have a direct impact on export and output diversification (be uncorrelated with the error term of the regression) and (ii) be highly correlated with gender inequality, the endogenous regressor of interest. Similar to the institutions and growth literature, we draw from a large dataset of legal restrictions on women's economic activity. We argue that gender-based legal restrictions—the mere existence of laws on the books of a country—do not exert a direct impact on export and output diversification, thus fulfilling the first condition of exogeneity, which we confirm with the Hansen statistical test. As argued in the previous section, legal rights have been shown in various strands of the literature to have a direct and strong impact on gender inequality. This makes them good candidates to fulfill the second condition of relevance of the instrument in theory. We also confirm this in the next section, which discusses our results.

4. Results: Gender Inequality Impedes Diversification

In line with our hypothesis, we find that gender inequality is strongly and negatively associated with export diversification in low-income and developing countries, even after accounting for the other drivers of diversification discussed earlier. Table 1 presents our baseline regressions, which test for gender being a driver of export diversification along with explanatory factors previously highlighted in the literature, including a large set of structural country characteristics, policies, and cyclical factors. The regressions also include country and time fixed effects to capture common fluctuations across time and inherent country characteristics. In particular, we find the following:

- *Gender inequality, as measured by the extended version of the UN's Gender Inequality Index, is strongly associated with export diversification, in line with our hypothesis.* In particular, moving from a situation of absolute gender inequality to perfect gender equality measured by the index could decrease the Theil index of export diversification, i.e., increase export diversification in low-income and developing countries, by 0.6 to 2 units. The magnitude of this effect is equivalent to up to about two standard deviations of the index across low-income and developing countries. Looking beyond low-income and developing countries, the results show that higher levels of gender inequality are significantly associated with lower levels of export diversification across all levels of development.

Controlling for other factors and policies highlights the robustness of the association of gender equality with economic diversification, with the gender equality remaining positively associated with diversification even as other variables are added in Table 1, columns (2) to (9). In particular:

- *The effect of gender inequality comes on top of structural characteristics previously highlighted in the literature.* Our results confirm the U-shaped relationship between export diversification and development (Dabla-Norris et al. 2013) in which countries diversify until they reach a certain level of development but re-concentrate afterwards. As expected, a higher share of mining in output is associated with a less diversified export base. In line with a larger pool of talent, population size (in most of our specifications) and human capital (in some specifications) are associated with higher export diversification.
- *The impact of gender inequality remains when controlling for policies associated with export diversification.* In particular, we show that institutions—creating a better business environment, e.g., as measured by the Frasier Summary Index of Institutions or legal systems and property rights—are significantly and positively associated with higher levels of diversification. A higher degree of openness in international trade expands the possible pool of trading partners and demand for exports, and our results confirm a positive and significant relationship with export diversification. Better infrastructure is also strongly associated with higher degrees of export diversification.

- Finally, macroeconomic factors also appear to play a role. Real exchange rate appreciation and terms-of-trade improvement are associated with lower degrees of export diversification, possibly reflecting the effect of lower price competitiveness in the short term and higher quantities of exports of main sectors when their prices are high.⁵

Gender inequality is negatively associated with output diversification in low-income and developing countries. To capture the role that the service sector may play in the economy, we examine output diversification in a similar empirical setup. In Table 2, the results for structural characteristics and policies are broadly comparable to the ones on export diversification described above. We find that gender inequality in low-income and developing countries is negatively associated with output diversification in all our specifications. However, we find mixed results on gender inequality for the remainder of countries. There is a significant and positive association of gender inequality and output diversification in some of the regressions for these countries, likely reflecting the fact that low gender inequality may result in greater participation of women in the service sector, in which countries tend to re-concentrate production as they develop.

In addition, our results provide evidence on two main channels through which gender inequality inhibits economic diversification. To test for the contribution of different dimensions of gender inequality, we include female labor force participation, gender gaps in education, female representation in parliament, and indicators of female health (maternal mortality and adolescent fertility) simultaneously into our regressions. The results in Table 3 highlight that there is some evidence for the *human capital channel*—a higher female-to-male enrollment ratio is significantly and positively related to export diversification, particularly in low-income and developing countries. In addition, there is evidence for the *resource allocation channel*, as higher female labor force participation rates are associated with higher export diversification levels in low-income and developing economies. The results also provide some evidence that better health outcomes, in terms of lower maternal mortality ratios and adolescent fertility rates are positively associated with export diversification. Table 4 highlights that the results are broadly similar for output diversification, where higher female labor force participation and higher educational enrollment ratios for girls relative to boys in low-income and developing countries are associated with higher output diversification when controlling for policies and institutions.

Finally, we also find evidence for causality in the specifications by instrumenting gender inequality with legal rights, which helps us to strengthen confirmation of our hypothesis by providing evidence of a causal effect of gender equality on diversification. Table 5 highlights gender inequality as a significant determinant of export and output diversification, even after including legal rights for women, such as the right to be the head of a household or full community marital property rights, as instruments for gender inequality in GMM regressions. The instruments we use pass standard econometric and rule-of-thumb tests. Each of the instruments is individually significant in the first-stage regressions and the F-statistics of the IV regressions are well above the rule-of-thumb threshold value of 10. In addition, in specifications with two or more instruments, the p-values of the Hansen J-statistic do not allow us to reject the joint null hypothesis that the instruments are uncorrelated with the error term, supporting our hypothesis that the excluded instruments are indeed correctly excluded from the estimated equation. These results suggest that gender inequality may indeed be a cause of lower economic diversification.

⁵ The results hold when real GDP per capita growth is used as an alternative to capture cyclical components. Several measures of income inequality were included in the regressions but did not yield significant results.

Table 1. Fixed effects regressions: Drivers of export diversification. Dependent variable: Theil index of export diversification.

Explanatory Variables:	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Gender Inequality									
Gender Inequality Index	0.703 ** (0.273)	0.752 *** (0.278)	0.776 *** (0.277)	1.156 *** (0.319)	1.141 *** (0.284)	1.381 *** (0.282)	1.078 *** (0.294)	0.983 *** (0.298)	0.665 ** (0.264)
– in LIDC	1.014 ** (0.431)	0.983 ** (0.438)	1.113 ** (0.435)	0.338 (0.457)	0.880 ** (0.432)	0.120 (0.440)	0.274 (0.405)	0.538 (0.417)	0.630 (0.426)
Structural Factors									
Log(Population)	−0.707 *** (0.133)	−0.560 *** (0.135)	−0.568 *** (0.136)	−1.059 *** (0.156)	−0.434 *** (0.146)	−0.222 (0.145)	−0.682 *** (0.138)	−0.450 *** (0.148)	−0.101 (0.147)
Lag Human capital index	0.0460 (0.109)	0.0406 (0.110)	0.0743 (0.110)	−0.112 (0.127)	−0.0729 (0.111)	0.0309 (0.111)	−0.286 ** (0.116)	−0.285 ** (0.118)	0.0887 (0.103)
Log(Real GDP per capita)	−1.838 *** (0.294)	−2.371 *** (0.289)	−1.712 *** (0.308)	−0.215 (0.310)	−1.736 *** (0.297)	−0.970 *** (0.311)	−1.166 *** (0.296)	−1.750 *** (0.301)	−0.971 *** (0.328)
– squared	0.114 *** (0.0174)	0.140 *** (0.0172)	0.103 *** (0.0182)	0.0245 (0.0190)	0.108 *** (0.0179)	0.0605 *** (0.0188)	0.0704 *** (0.0178)	0.112 *** (0.0182)	0.0516 *** (0.0191)
Mining as share of GDP	0.00937 ** (0.00396)	0.00694 * (0.00398)	0.0119 *** (0.00416)	0.0253 *** (0.00377)	0.00694 * (0.00407)	0.0119 *** (0.00407)	0.0221 *** (0.00392)	0.0266 *** (0.00407)	0.0236 *** (0.00472)
Policies									
1. Institutions									
Fraser Institute Sum. Index	−0.116 *** (0.0137)								−0.0700 *** (0.0178)
Legal Syst.& Property Rights		−0.0358 *** (0.0102)							
2. Openness									
Freedom to trade			−0.0646 *** (0.00858)						−0.0219 * (0.0114)
Globalization Index				−0.0123 *** (0.00268)					
3. Infrastructure									
Length of road network					−0.0300 ** (0.0144)				
Log(landlines/1000 workers)						−0.129 *** (0.0177)			−0.110 *** (0.0180)

Table 1. Cont.

Explanatory Variables:	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Macro/Cyclical Factors									
Terms of Trade							0.00313 *** (0.000347)		0.00427 *** (0.000440)
Log(REER)								0.186 *** (0.0519)	0.305 *** (0.0490)
Constant	11.90 *** (1.201)	13.69 *** (1.209)	10.78 *** (1.273)	5.434 *** (1.209)	10.21 *** (1.232)	6.928 *** (1.284)	8.737 *** (1.223)	9.483 *** (1.263)	5.712 *** (1.436)
Observations	1841	1835	1836	1798	1726	1726	1903	1909	1583
Countries	100	100	100	105	89	89	100	102	84
R-squared	0.181	0.141	0.174	0.108	0.110	0.136	0.127	0.118	0.271

Standard errors in parentheses: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Note: All specifications include country and time fixed effects.

Table 2. Fixed effects regressions: Drivers of output diversification. Dependent variable: Theil index of output diversification.

Explanatory Variables:	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Gender Inequality									
Gender Inequality Index	−0.0552 * (0.0310)	−0.0344 (0.0315)	−0.0867 ** (0.0425)	−0.1000 *** (0.0369)	0.0283 (0.0305)	0.0397 (0.0307)	−0.0103 (0.0398)	−0.0932 ** (0.0369)	0.0404 (0.0393)
– in LIDC	0.188 *** (0.0488)	0.203 *** (0.0495)	0.212 *** (0.0709)	0.310 *** (0.0527)	0.194 *** (0.0462)	0.158 *** (0.0476)	0.302 *** (0.0629)	0.190 ** (0.0794)	0.268 ** (0.119)
Structural Factors									
Log(Population)	−0.0376 ** (0.0150)	−0.0318 ** (0.0154)	−0.0352 (0.0224)	−0.0524 *** (0.0179)	−0.0309 ** (0.0157)	−0.0269 * (0.0158)	−0.0568 *** (0.0196)	−0.0466 ** (0.0221)	−0.0424 * (0.0240)
Lag Human capital index	0.0350 *** (0.0121)	0.0346 *** (0.0122)	0.0271 (0.0171)	0.0376 *** (0.0145)	0.0219 * (0.0120)	0.0281 ** (0.0121)	0.0254 * (0.0154)	0.0509 *** (0.0153)	0.0320 ** (0.0152)
Log(Real GDP per capita)	−0.215 *** (0.0336)	−0.238 *** (0.0353)	−0.225 *** (0.0431)	−0.233 *** (0.0360)	−0.221 *** (0.0319)	−0.192 *** (0.0339)	−0.340 *** (0.0407)	−0.132 ** (0.0543)	−0.208 *** (0.0707)
– squared	0.0103 *** (0.00199)	0.0112 *** (0.00209)	0.0108 *** (0.00262)	0.0107 *** (0.00220)	0.00985 *** (0.00192)	0.00809 *** (0.00205)	0.0179 *** (0.00254)	0.00578 * (0.00302)	0.0121 *** (0.00398)
Mining as share of GDP	0.000214 (0.000449)	0.000283 (0.000474)	−0.00216 *** (0.000583)	−0.000943 ** (0.000430)	0.00355 *** (0.000427)	0.00381 *** (0.000434)	−0.000152 (0.000461)	−0.00367 *** (0.000638)	−0.00519 *** (0.000741)

Table 2. Cont.

Explanatory Variables:	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Policies									
1. Institutions									
Fraser Institute Sum. Index	−0.00961 *** (0.00155)								−0.00816 *** (0.00191)
2. Openness									
Freedom to trade		−0.00224 ** (0.000976)							
Average Tariff Rates			0.0290 *** (0.0108)						0.0647 *** (0.0111)
Globalization Index				−0.00105 *** (0.000307)					
3. Infrastructure/Investment									
Length of road network					−0.00464 *** (0.00153)				
Log(Landlines/1000 workers)						−0.00716 *** (0.00193)			−0.00452 * (0.00234)
Investment per worker							-3.79×10^{-6} *** (7.98×10^{-7})		-5.94×10^{-6} *** (7.84×10^{-7})
4. Financial Development									
Financial reform index								−0.0760 *** (0.0126)	−0.0293 ** (0.0128)
Constant	1.386 *** (0.137)	1.440 *** (0.146)	1.435 *** (0.174)	1.550 *** (0.141)	1.325 *** (0.132)	1.221 *** (0.140)	1.895 *** (0.170)	1.101 *** (0.232)	1.322 *** (0.307)
Observations	1880	1875	1410	1839	1752	1752	1783	1128	1027
Countries	102	102	100	107	90	90	108	75	67
R-squared	0.165	0.146	0.108	0.209	0.221	0.223	0.190	0.167	0.220

Standard errors in parentheses: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Note: All specifications include country and time fixed effects.

Table 3. Fixed effects regressions: Drivers of export diversification—focus on dimensions of gender inequality. Dependent variable: Theil index of export diversification.

Explanatory Variables:	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Gender Inequality									
Female labor force participation rate	0.473 (0.472)	0.970 ** (0.466)	0.758 (0.468)	1.762 *** (0.532)	0.995 ** (0.457)	0.859 * (0.462)	1.562 *** (0.478)	1.478 *** (0.467)	−0.0324 (0.423)
- in LIDC	−2.748 *** (0.844)	−3.458 *** (0.867)	−2.935 *** (0.851)	−3.146 *** (0.888)	−3.400 *** (0.980)	−3.111 *** (1.004)	−2.609 *** (0.833)	−2.185 *** (0.811)	−2.092 ** (1.066)
Secondary enrollment ratio	−0.00603 (0.281)	0.0555 (0.284)	0.0444 (0.283)	−0.580* (0.315)	−0.374 (0.270)	−0.328 (0.270)	−0.333 (0.291)	−0.279 (0.282)	0.316 (0.247)
- in LIDC	−0.986 ** (0.480)	−0.987 ** (0.490)	−1.034 ** (0.484)	0.119 (0.446)	−0.195 (0.456)	−0.167 (0.456)	−0.0318 (0.461)	−1.012 ** (0.424)	−1.590 *** (0.590)
Women in parliament	−0.00265 (0.00278)	−0.00212 (0.00282)	−0.00271 (0.00292)	−0.00525 * (0.00315)	−0.00250 (0.00283)	−0.00292 (0.00283)	−0.00337 (0.00293)	−0.000692 (0.00315)	0.00444 (0.00277)
- in LIDC	0.00691 (0.00482)	0.00482 (0.00487)	0.00606 (0.00493)	0.00452 (0.00517)	0.00752 (0.00461)	0.00800 * (0.00459)	0.00418 (0.00487)	0.00650 (0.00484)	0.00578 (0.00471)
Maternal mortality ratio	0.00142 ** (0.000695)	0.00156 ** (0.000700)	0.00151 ** (0.000700)	0.00154 * (0.000800)	0.00130 * (0.000676)	0.00104 (0.000692)	0.00145 ** (0.000719)	0.00152 ** (0.000699)	0.00169 *** (0.000629)
- in LIDC	−0.000415 (0.000735)	−0.000884 (0.000755)	−0.000411 (0.000741)	−0.00141 * (0.000830)	−0.000186 (0.000727)	−1.73 × 10 ^{−5} (0.000733)	−0.00129 * (0.000750)	−0.000668 (0.000733)	−0.00111 (0.000672)
Adolescent fertility rate	0.000586 (0.00271)	0.000761 (0.00274)	−0.000966 (0.00277)	0.00377 (0.00309)	0.00172 (0.00266)	0.00231 (0.00265)	0.00318 (0.00284)	0.00267 (0.00288)	0.00341 (0.00254)
- in LIDC	−0.00143 (0.00409)	0.00138 (0.00419)	−0.000821 (0.00411)	0.00640 (0.00403)	0.00476 (0.00457)	0.00393 (0.00461)	0.00702 * (0.00396)	0.00436 (0.00375)	0.0122 ** (0.00535)
Structural Factors									
Log(Population)	−0.0711 (0.234)	0.171 (0.237)	0.181 (0.236)	−0.742 *** (0.271)	0.329 (0.239)	0.340 (0.238)	−0.305 (0.240)	0.239 (0.247)	0.667 *** (0.238)
Lag Human capital index	−0.358 ** (0.155)	−0.310 ** (0.158)	−0.392 ** (0.157)	−0.244 (0.185)	−0.313 ** (0.152)	−0.288 * (0.152)	−0.483 *** (0.162)	−0.419 *** (0.158)	−0.387 *** (0.139)
Log(Real GDP per capita)	−2.059 *** (0.609)	−2.261 *** (0.624)	−2.051 *** (0.622)	1.137 * (0.595)	−1.698 *** (0.608)	−1.626 *** (0.610)	0.248 (0.586)	−0.766 (0.563)	−0.848 (0.617)
- squared	0.125 *** (0.0356)	0.131 *** (0.0365)	0.120 *** (0.0363)	−0.0578 (0.0354)	0.106 *** (0.0357)	0.101 *** (0.0358)	−0.0136 (0.0348)	0.0550 (0.0335)	0.0495 (0.0362)
Mining as share of GDP	0.0114 ** (0.00566)	0.00874 (0.00573)	0.0151 ** (0.00607)	0.0122 ** (0.00557)	0.0142 ** (0.00562)	0.0151 *** (0.00565)	0.0143 *** (0.00549)	0.0191 *** (0.00587)	0.0390 *** (0.00629)

Table 3. Cont.

Explanatory Variables:	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Policies									
1. Institutions									
Fraser Institute Sum. Index	−0.115 *** (0.0221)								−0.124 *** (0.0245)
Legal Syst. & Property Rights		−0.0437 *** (0.0169)							
2. Openness									
Freedom to trade			−0.0516 *** (0.0149)						−0.00345 (0.0168)
Globalization Index				−0.0114 *** (0.00368)					
3. Infrastructure									
Length of road network					−0.0276 (0.0188)				
Log(landlines) per 1000 workers						−0.0499 * (0.0271)			−0.0532 ** (0.0261)
4. Macro/Cyclical factors									
Terms of Trade							0.00287 *** (0.000536)		0.00485 *** (0.000607)
Log(REER)								−0.00341 (0.0798)	0.236 *** (0.0759)
Constant	12.64 *** (2.540)	12.50 *** (2.640)	11.78 *** (2.583)	−0.198 (2.488)	8.799 *** (2.590)	8.703 *** (2.588)	2.426 (2.484)	4.838 * (2.498)	3.450 (2.704)
Observations	1033	1034	1032	954	989	989	1083	1084	927
Countries	96	97	96	101	86	86	96	98	81
R-squared	0.203	0.162	0.194	0.133	0.174	0.175	0.149	0.168	0.354

Standard errors in parentheses *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Note: All specifications include country and time fixed effects.

Table 4. Fixed Effects Regressions: Drivers of output diversification—focus on dimensions of gender inequality. Dependent variable: Theil index of output diversification.

Explanatory Variables:	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Gender Inequality									
Female labor force participation rate	0.0160 (0.0449)	0.0562 (0.0446)	0.0665 (0.0642)	−0.0181 (0.0505)	0.0776 * (0.0469)	0.0624 (0.0474)	0.0613 (0.0543)	0.0580 (0.0681)	0.0127 (0.0704)
- in LIDC	0.0695 (0.0810)	0.0358 (0.0821)	−0.292 ** (0.123)	−0.00734 (0.0843)	0.0568 (0.0924)	0.0749 (0.0939)	0.0384 (0.0978)	−0.421 *** (0.156)	−0.342 ** (0.174)
Secondary enrollment ratio	0.124 *** (0.0263)	0.133 *** (0.0265)	0.0947 *** (0.0359)	0.119 *** (0.0293)	0.110 *** (0.0271)	0.117 *** (0.0271)	0.0922 *** (0.0315)	0.0751 ** (0.0371)	0.0313 (0.0384)
- in LIDC	−0.0407 (0.0453)	−0.0502 (0.0458)	−0.107 * (0.0595)	−0.0965 ** (0.0417)	−0.0696 (0.0439)	−0.0695 (0.0439)	−0.0760 * (0.0452)	−0.0776 (0.0740)	−0.282 *** (0.105)
Women in parliament	−0.000203 (0.000260)	−0.000462 * (0.000277)	−0.000449 (0.000352)	−0.000178 (0.000295)	−0.000573 ** (0.000283)	−0.000619 ** (0.000284)	−0.000456 (0.000324)	−0.000386 (0.000360)	−0.000523 (0.000379)
- in LIDC	−0.000136 (0.000458)	3.47×10^{-5} (0.000470)	0.000539 (0.000696)	7.31×10^{-5} (0.000473)	6.11×10^{-5} (0.000439)	0.000167 (0.000440)	0.000372 (0.000531)	0.000137 (0.00102)	−0.000579 (0.00128)
Maternal mortality ratio	0.000162 ** (6.51×10^{-5})	0.000171 *** (6.58×10^{-5})	6.44×10^{-5} (9.43×10^{-5})	6.27×10^{-5} (7.45×10^{-5})	0.000203 *** (6.89×10^{-5})	0.000174 ** (7.02×10^{-5})	0.000117 (7.95×10^{-5})	5.40×10^{-5} (9.70×10^{-5})	1.97×10^{-7} (9.69×10^{-5})
- in LIDC	-8.27×10^{-5} (6.94×10^{-5})	-7.99×10^{-5} (7.02×10^{-5})	4.73×10^{-5} (1.00×10^{-4})	7.38×10^{-5} (7.79×10^{-5})	−0.000136 * (7.44×10^{-5})	−0.000115 (7.52×10^{-5})	3.04×10^{-5} (8.34×10^{-5})	-1.43×10^{-6} (0.000105)	7.79×10^{-6} (0.000122)
Adolescent fertility rate	0.000925 *** (0.000258)	0.00101 *** (0.000264)	0.000931 ** (0.000373)	0.000769 *** (0.000290)	0.000327 (0.000270)	0.000428 (0.000270)	0.000535 * (0.000305)	0.000401 (0.000474)	0.000758 (0.000488)
- in LIDC	0.000974 ** (0.000387)	0.000993 ** (0.000391)	0.00119 ** (0.000523)	0.00110 *** (0.000378)	0.00163 *** (0.000443)	0.00153 *** (0.000450)	0.00132 *** (0.000421)	0.00181 *** (0.000655)	−0.000633 (0.00104)
Structural Factors									
Log(Population)	−0.0116 (0.0224)	−0.00569 (0.0227)	0.0546 (0.0334)	0.0336 (0.0259)	−0.00484 (0.0244)	−0.00776 (0.0243)	0.0438 (0.0286)	−0.0281 (0.0396)	−0.00629 (0.0422)
Lag Human capital index	0.0183 (0.0148)	0.0191 (0.0150)	0.0146 (0.0221)	0.0234 (0.0176)	0.0199 (0.0157)	0.0230 (0.0157)	0.0158 (0.0188)	0.0479 ** (0.0239)	0.0459 * (0.0239)
Log(Real GDP per capita)	−0.0755 (0.0585)	−0.0619 (0.0602)	−0.196 ** (0.0801)	−0.217 *** (0.0566)	−0.124 ** (0.0607)	−0.115 * (0.0613)	−0.240 *** (0.0691)	0.0433 (0.129)	−0.243 (0.152)
- squared	0.00239 (0.00342)	0.000941 (0.00352)	0.00921 * (0.00476)	0.00980 *** (0.00337)	0.00422 (0.00356)	0.00368 (0.00360)	0.0112 *** (0.00416)	−0.00306 (0.00718)	0.0141 (0.00859)
Mining as share of GDP	0.000103 (0.000544)	−0.000482 (0.000587)	−0.00373 *** (0.000827)	-9.53×10^{-5} (0.000510)	0.00241 *** (0.000554)	0.00252 *** (0.000562)	0.000476 (0.000556)	−0.00552 *** (0.00123)	−0.00855 *** (0.00131)

Table 4. Cont.

Explanatory Variables:	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Policies									
1. Institutions									
Fraser Institute Sum. Index	−0.0107 *** (0.00212)								−0.00724 ** (0.00341)
2. Openness									
Freedom to trade		−0.00251 * (0.00144)							
Average Tariff Rates			0.0636 *** (0.0183)						0.143 *** (0.0251)
Globalization Index				−0.00121 *** (0.000351)					
3. Infrastructure/Investment									
Length of road network					−0.00474 ** (0.00190)				
Log(landlines) per 1000 workers						−0.00573 ** (0.00267)			−0.00403 (0.00390)
Investment per worker							$-1.67 \times 10^{-6} *$ (9.59×10^{-7})		$-5.39 \times 10^{-6} ***$ (1.41×10^{-6})
4. Financial Development									
Financial reform index								−0.115 *** (0.0201)	−0.0630 *** (0.0226)
Constant	0.541 ** (0.245)	0.451 * (0.251)	0.933 *** (0.330)	1.109 *** (0.236)	0.719 *** (0.257)	0.717 *** (0.258)	1.124 *** (0.291)	0.138 (0.554)	1.389 ** (0.656)
Observations	1063	1062	681	987	1014	1014	942	552	485
Countries	98	98	95	103	87	87	104	73	65
R-squared	0.245	0.229	0.231	0.330	0.259	0.258	0.294	0.276	0.341

Standard errors in parentheses *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Note: All specifications include country and time fixed effects.

Table 5. Instrumental variable GMM: Drivers of export and output diversification.

Panel A: Dependent Variable: Export Diversification			Panel B: Dependent Variable: Output Diversification		
Explanatory Variables:	(1)	(2)	Explanatory Variables:	(1)	(2)
GII Index	5.785 *** (1.942)	3.534** (1.739)	GII Index	1.778 *** (0.361)	0.153 *** (0.0387)
Log(Population)	−0.976 *** (0.214)	−0.252 (0.271)	Log(Population)	−0.0830 ** (0.0396)	−0.134 *** (0.0230)
Lag Human capital index	0.0251 (0.196)	0.420 *** (0.162)	Lag Human capital index	0.131 *** (0.0321)	−0.00844 (0.0116)
Log(GDP per capita)	−1.307 *** (0.337)	−0.666 * (0.343)	Log(GDP per capita)	−0.390 *** (0.0726)	−0.222 *** (0.0809)
- squared	0.0931 *** (0.0201)	0.0360 * (0.0196)	- squared	0.0230 *** (0.00446)	0.0141 *** (0.00473)
Mining as share of GDP	0.0318 *** (0.00710)	0.0105 (0.00659)	Mining as share of GDP	0.00129 (0.00126)	−8.56 × 10 ^{−5} (0.000944)
Fraser Institute Sum. Index		−0.0498 (0.0363)	Fraser Institute Sum. Index		−0.0114 *** (0.00169)
Freedom to trade		−0.0405 *** (0.0141)	Average Tariff Rates		0.0361 *** (0.0105)
Log(landlines) per 1000 workers		−0.0919 *** (0.0281)	Log(landlines) per 1000 workers		−0.00201 (0.00190)
Terms of Trade		0.00427 *** (0.000609)	Investment per worker		−5.89 × 10 ^{−6} *** (1.02 × 10 ^{−6})
Log(REER)		0.301 *** (0.0588)	Financial reform index		−0.00467 (0.0124)
Constant	5.515 *** (2.046)	3.438 (2.466)	Constant	0.923 *** (0.329)	1.578 *** (0.354)
Observations	1552	1204	Observations	1554	833
p-value of Hansen J statistic	0.296	0.248	p-value of Hansen J statistic	0.548	0.276
Instrument F-test	13.27	12.85	Instrument F-test	16.28	33.44

Standard errors in parentheses *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Note: (Panel A) All specifications include country and time fixed effects. (Panel B) Note: All specifications include country and time fixed effects. Lesotho and Mauritania are dropped from the estimation due to insufficient observations.

5. Conclusions

This paper presents, to the best of our knowledge, the first empirical evidence that gender inequality impacts both export and output diversification. Using a multi-dimensional index to capture gender inequality, as well as individual gender inequality indicators, we demonstrate that gender inequality, both in outcomes and in opportunities, negatively impacts export and output diversification, especially in low-income and developing countries. In particular, moving from a situation of absolute gender *inequality* to perfect gender *equality* measured by the index could *increase* export diversification in low-income and developing countries, by 0.6 to 2 units—equivalent to up to about two standard deviations of the index across low-income and developing countries.

In the rest of the panel, when data on the service sector is included, as in the regressions of output diversification, we get mixed results. This could be due to the fact that low gender inequality may result in greater participation of women in the service sector, in which countries tend to re-concentrate production as they develop and could be further explored in future research. Our research provides evidence that both gender equity in opportunities, as well as outcomes, matter for economic diversification. In particular, we show that both gender inequalities in opportunities, such as education, and lower female labor force participation, are negatively associated with diversification. The former supports the hypothesis of inequality constraining the level of human capital, which limits diversification—and could be tested along generalized inequality of opportunity in future research. The latter supports the theory of inefficient allocation of resources leading to suboptimal creation of ideas and development of sectors.

Our empirical work provides support for causality between gender inequality and diversification. We separate the effect of gender inequality on diversification from the reverse effect of diversification on gender inequality, due to our empirical estimation strategy, which uses the country-specific *de jure* laws and regulations as instruments for gender inequality. These legal restrictions, such as restrictions the right to be the head of a household, skew the efficient allocation of resources by impeding women's economic participation and by preventing households from giving the same opportunities to daughters as they do to sons.

By linking gender inequality to lower economic diversification—which is widely acknowledged as a source of sustainable growth—we highlight a new channel through which gender equality boosts growth.

Author Contributions: Conceptualization, M.N., L.K., K.K. and R.K.; methodology, M.N., L.K. and R.K.; software, R.K., L.K. and M.N.; validation, R.K., L.K. and M.N.; formal analysis, M.N., L.K. and R.K.; investigation, M.N., L.K. and R.K.; resources, L.K., M.N., K.K. and R.K.; data curation, R.K. and L.K.; writing—original draft preparation, M.N., L.K. and R.K.; writing—review and editing, K.K., M.N., L.K. and R.K.; visualization, M.N., L.K. and R.K.; supervision, K.K., M.N. and L.K.; project administration, M.N. and L.K.; funding acquisition, L.K. and K.K.

Funding: This paper is part of a research project on macroeconomic policy in low-income countries supported by the U.K.'s Department for International Development (DFID), grant number (IATI Identifier) GB-1-202960, and it should not be reported as representing the views of the International Monetary Fund or of DFID.

Acknowledgments: The authors wish to thank Sonali Jain-Chandra, seminar participants at the Strategy, Policy, and Review department, as well as reviewers from the African, European, Middle East and Central Asia, Communications, and Research Departments of the International Monetary Fund (IMF) for helpful comments.

Conflicts of Interest: The authors declare no conflict of interest. The funders had no role in the design of the study; in the collection, analyses, or interpretation of data; in the writing of the manuscript, or in the decision to publish the results.

Appendix A. Summary Statistics

Table A1. Key Variables and Summary Statistics.

Variable	Source	Full Sample					LIDC				
		Obs	Mean	Std. Dev.	Min	Max	Obs	Mean	Std. Dev.	Min	Max
Export Diversification Theil	IMF Diversification Toolkit	6378	3.5	1.2	1.0	6.4	2159	4.2	0.9	1.8	6.4
Output Diversification Theil	IMF Diversification Toolkit	7065	0.3	0.2	0.0	1.7	2259	0.3	0.2	0.0	1.6
Log(GDP per capita)	World Economic Outlook	6141	8.5	1.2	5.2	11.7	1910	7.2	0.5	5.2	8.8
Log(Population)	PWT 8.1	6141	1.7	1.9	−3.2	7.2	1910	1.8	1.4	−2.6	5.1
Human capital index (5-year lag)	PWT 8.1/ Barro Lee	4385	2.1	0.6	1.0	3.6	1289	1.6	0.4	1.0	2.9
Mining as share of GDP	IMF Jobs and Income Surveillance toolkit	4831	21.0	11.6	0.8	85.6	1865	17.7	11.6	0.8	75.9
GII Index	IMF GDI GII database	2580	0.5	0.2	0.0	0.8	774	0.6	0.1	0.3	0.8
Ratio of female tertiary teachers	WDI	2105	0.3	0.1	0.0	0.8	521	0.2	0.1	0.0	0.8
Unmarried women; equal property rights	Women, Business, and the Law	3707	0.9	0.3	0.0	1.0	1470	0.9	0.3	0.0	1.0
Married women; equal property rights	Women, Business, and the Law	3688	0.8	0.4	0.0	1.0	1431	0.7	0.5	0.0	1.0
Married women; head household	Women, Business, and the Law	3723	0.6	0.5	0.0	1.0	1466	0.5	0.5	0.0	1.0
Married women; legal proceedings	Women, Business, and the Law	3763	0.9	0.3	0.0	1.0	1506	0.8	0.4	0.0	1.0
Married women; bank account	Women, Business, and the Law	3742	0.9	0.3	0.0	1.0	1490	0.8	0.4	0.0	1.0
Equal inheritance, sons and daughters	Women, Business, and the Law	3688	0.7	0.5	0.0	1.0	1431	0.6	0.5	0.0	1.0
Joint titling of property	Women, Business, and the Law	3582	0.4	0.5	0.0	1.0	1354	0.4	0.5	0.0	1.0
Full community marital property regime	Women, Business, and the Law	3589	0.1	0.2	0.0	1.0	1351	0.0	0.2	0.0	1.0
Partial community marital property regime	Women, Business, and the Law	3589	0.4	0.5	0.0	1.0	1351	0.3	0.5	0.0	1.0
Separate property marital property regime	Women, Business, and the Law	3589	0.4	0.5	0.0	1.0	1351	0.5	0.5	0.0	1.0
Guaranteed equity	Women, Business, and the Law	3734	0.9	0.3	0.0	1.0	1501	0.9	0.3	0.0	1.0
Nondiscrimination clause	Women, Business, and the Law	3734	0.4	0.5	0.0	1.0	1501	0.4	0.5	0.0	1.0
Valid customary law	Women, Business, and the Law	3734	0.3	0.5	0.0	1.0	1501	0.5	0.5	0.0	1.0
Female labor force participation rate	WDI	3591	0.5	0.2	0.1	0.9	1197	0.6	0.2	0.1	0.9
Secondary enrollment ratio	WDI	4371	0.9	0.3	0.0	3.1	1230	0.7	0.3	0.0	2.1
Women in parliament	WDI	2425	14.1	9.9	0.0	56.3	753	12.0	9.1	0.0	56.3
Maternal mortality ratio	WDI	3591	272.0	374.7	3.0	2900.0	1218	623.4	422.8	29.0	2900.0
Adolescent fertility rate	WDI	3696	65.0	49.3	3.1	228.6	1218	106.6	48.3	18.0	222.4
Fraser Institute Summary Index	Fraser Institute	3655	5.9	1.4	2.0	9.2	1100	5.1	1.1	2.0	7.5
Legal system and property rights	Fraser Institute	3509	5.3	1.9	1.1	9.6	989	4.0	1.1	1.6	6.8
Freedom to trade	Fraser Institute	3820	5.8	2.4	0.0	10.0	1215	4.3	2.1	0.0	8.8
Globalization index	KOF Index of Globalization	4451	46.3	19.2	9.6	92.9	1728	31.4	10.2	9.6	63.1
Length of road network	Calderon-Serven database	3755	−1.2	1.4	−5.2	1.6	1043	−2.0	1.4	−5.2	0.0
Log(Landlines per 1000 workers)	Calderon-Serven database	3765	3.7	2.0	−0.6	7.2	1043	1.8	1.1	−0.6	5.2
Terms of Trade	World Economic Outlook	4334	109.7	48.7	5.5	602.9	1477	124.5	69.5	5.5	602.9
Log(REER)	IFS	3350	4.7	0.7	0.7	15.3	1171	4.9	1.0	0.7	15.3
Average Tariff Rates	Trade Index	3194	0.7	0.2	0.0	1.0	999	0.7	0.2	0.0	1.0
Investment per worker	PWT	4012	4589.1	5361.8	−832.1	46,086.0	1500	537.9	555.2	−832.1	5207.6
Financial reform index	IMF Index of Financial reform	2527	0.5	0.3	0.0	1.0	558	0.3	0.2	0.0	0.9
Gini index	WDI	1035	40.7	10.3	16.2	99.9	233	43.0	9.1	25.9	69.5
Income ratio (top 20%/bottom 20%)	WDI	1034	10.5	11.6	2.2	278.2	233	12.6	20.0	3.7	278.2
Agriculture, value added (% of GDP)	WDI	5099	19.0	15.5	0.0	74.3	1806	33.5	13.6	3.1	74.3
Rural population	WDI	7044	50.9	24.8	0.0	97.2	2259	71.3	14.8	23.0	97.2
Fuel exports	WDI	4516	16.2	29.1	0.0	359.3	1025	13.2	30.1	0.0	359.3
Domestic credit to private sector	WDI	5731	37.8	35.8	0.1	312.2	1781	15.3	11.5	0.2	114.7
Real GDP per capita growth rate	World Economic Outlook	5981	0.0	0.1	−1.1	1.0	1861	0.0	0.1	−0.7	0.7

Appendix B. Country Sample

Appendix B.1. Non-LIDC Countries

Albania, Argentina, Armenia, Australia, Austria, Belgium, Brazil, Bulgaria, Canada, Chile, China, Colombia, Costa Rica, Croatia. Denmark, Dominican Republic, Ecuador, Arab, Republic of Egypt, El Salvador, Estonia, Finland, France, Germany, Greece, Guatemala, Hungary, India, Indonesia, Islamic Republic of Iran, Iraq, Ireland, Israel, Italy, Jamaica, Japan, Jordan, Kazakhstan, Latvia, Lithuania, Malaysia, Mexico, Morocco, Namibia*

Netherlands, New Zealand, Norway, Pakistan, Panama, Paraguay, Peru, Philippines, Poland, Portugal, Romania, Saudi Arabia, Singapore, Slovak Republic, Slovenia, South Africa, Spain, Sri Lanka, Sweden, Switzerland, Syrian Arab Republic, Thailand, Tunisia, Turkey, Ukraine, United Kingdom, United States, Uruguay, Venezuela

Appendix B.2. LIDC Countries

Bangladesh, Benin, Bolivia, Burundi, Cambodia, Cameroon, Central African Republic, Democratic Republic of Congo, Republic of Congo, Côte d'Ivoire, Ghana, Honduras, Kenya, Kyrgyz Republic, People's Democratic Republic of Lao, Lesotho*, Liberia, Malawi, Mali, Mauritania, Moldova, Mongolia, Mozambique, Nepal, Niger, Rwanda, Senegal, Sierra Leone, Sudan, Tajikistan, Tanzania, Togo, Uganda, Republic of Yemen, Zambia, Zimbabwe

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